

1 Did internal displacement from the 2010 earthquake in Haiti lead to long-term violence against
2 children? A matched pairs study design

3

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27

28 **Abstract**

29 **Background**

30 Empirical evidence is limited and contradictory on violence against children after internal
31 displacement from natural disasters. Understanding how internal displacement affects violence is
32 key in structuring effective prevention and response.

33

34 **Objective**

35 We examined the effect of internal displacement from the 2010 Haitian earthquake on long-term
36 physical, emotional, and sexual violence against children and outlined a methodological
37 framework to improve future evidence quality.

38

39 **Participants and setting**

40 We analyzed violence against adolescent girls and boys within the nationally representative, Haiti
41 Violence Against Children Survey.

42

43 **Methods**

44 We pre-processed data by matching on pre-earthquake characteristics for displaced and non-
45 displaced children and applied 95% confidence intervals from McNemar's exact test, with

46 sensitivity analyses, to evaluate differences in violence outcomes between matched pairs after the
47 earthquake.

48

49 **Results**

50 Internal displacement was not associated with past 12-month physical, emotional, and sexual
51 violence two years after the earthquake for girls and boys. Most violence outcomes were robust to
52 potential unmeasured confounding. Odds ratios for any form of violence against girls were 0.84
53 (95% CI: 0.52-1.33, $p = 0.500$) and against boys were 1.03 (95% CI: 0.61-1.73, $p = 1.000$).

54

55 **Conclusions**

56 Internal displacement was not a driver of long-term violence against children in Haiti. Current
57 global protocols in disaster settings may initiate services after the optimal window of time to
58 protect children from violence, and the post-displacement setting may be central in determining
59 violence outcomes. The combination of specific data structures and matching methodologies is
60 promising to increase evidence quality after rapid-onset natural disasters, especially in low-
61 resource settings.

62

63 **Keywords**

64 Natural disaster; internal displacement; humanitarian emergency; Haiti; violence

65

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68 or not-for-profit sectors.

69

70

Introduction

71 On January 12, 2010 Haiti experienced a 7.0 magnitude earthquake near the capital city of Port-
72 au-Prince (Doocy, Cherewick, & Kirsch, 2013). Port-au-Prince lost an estimated 23% of its
73 population from internal displacement to camps, informal settlements, and other regions of the
74 country (Lu, Bengtsson, & Holme, 2012). International humanitarian aid was substantial,
75 surpassing US\$9 billion, but the loss of infrastructure and high death toll among members of the
76 Haitian government and the United Nations fractured coordination (Kirsch, Sauer, & Guha Sapid,
77 2012; Ramachandran & Walz, 2015). Several evaluations highlighted that the provision of services
78 was insufficient to protect and respond to violence against internally displaced persons, or IDPs,
79 in both communities and camp settings (Center for Human Rights and Global Justice, 2012; The
80 Interuniversity Institute for Research and Development, 2010). Prior studies have reported
81 widespread criminality and sexual violence against women and girls committed by criminal gangs
82 and armed men in the immediate aftermath of the earthquake and months and years that followed
83 (Amnesty International, 2011; Kolbe et al., 2010). A study that isolated effect of the Haitian
84 earthquake on the probability of intimate partner violence among adult women nationally found
85 that physical violence increased in the most devastated areas of the country and decreased in
86 minimally affected areas. The probability of sexual violence likewise differed, decreasing by over
87 300 percent in moderately affected areas as opposed to devastated regions. Both physical and sexual
88 violence importantly were higher among women in IDP camps than the general population but not
89 significant different among women who were displaced by the earthquake (Weitzman & Behrman,
90 2016). It remains uncertain how earthquake exposure and internal displacement was associated
91 with violence against children within affected households and caregiving networks.

92

93 The pathways between natural disasters and violence against children are highly complex and
94 indirect. Population movement is often inherently part of natural disaster exposure. As Rashid and
95 Michaud (2000) highlight in their interviews with flood-affected communities in Bangladesh, girls
96 attributed new cases of sexual violence to the influx of unknown young men—some affiliated with
97 criminal groups—into their neighborhoods. Large-scale displacement into communities poses
98 security risks, especially when coupled with a breakdown of social systems of protection and
99 policing, common in disaster events. Predatory acts of sexual violence moreover can occur when
100 children are displaced to new environments, such as IDP camps or informal settlements. Camps
101 and informal settlements are often overcrowded transitional spaces that lack security in terms of
102 policing, secure housing and private living spaces, and lighting (Davis & Bookey, 2011; Standing,
103 Parker, & Bista, 2016).

104

105 Natural disasters and internal displacement lead to changes in the family system. Caregivers may
106 be physically separated from their children or die during or after a natural disaster. A lack of
107 supervision makes children vulnerable to violence but also threatens their overall wellbeing, which
108 elevates future violence risk (Cas, Frankenberg, Suriastini, & Thomas, 2014). Natural disasters
109 and internal displacement increase economic hardship, especially for low-income families
110 (Miljkovic & Miljkovic, 2014). New financial strains may cause caregivers to migrate to different
111 areas for work opportunities, leaving their children in the care of others or unaccompanied.
112 Particularly girls may travel independently over great distances to collect firewood or water, which
113 is a known risk for sexual violence (Spangaro et al., 2013). Children further may be forced to work
114 to support their families and face new threats of violence and exploitation from their employers.

115 Social support to families often is reduced, given the strain on social networks that similarly may
116 have experienced the natural disaster or increases in distance after displacement to separate regions
117 (Morris & Deterding, 2016).

118

119 Natural disasters and internal displacement cause extreme upheaval that psychologically can affect
120 individuals. Mental distress and psychopathology among caregivers and affected community
121 members is common and may produce increases in violence against children (Biswas, Rahman,
122 Mashreky, Rahman, & Dalal, 2010; Neria, Nandi, & Galea, 2008). Negative coping behaviors,
123 such as hazardous alcohol usage, often increase after natural disasters, which present an additional
124 risk factor for violence. After the 2004 Indian Ocean tsunami, for instance, alcohol usage by Sri
125 Lankan fathers was significantly associated with physical, emotional, and sexual violence ($\beta =$
126 $0.16, r_{bp} = 0.18, p < 0.01$) (Catani, Jacob, Schauer, Kohila, & Neuner, 2008). Economic hardship
127 further leads to caregiver stress, resulting in harsh acts of physical or emotional violence (Biswas
128 et al., 2010).

129

130 Children who are internally displaced are particularly vulnerable to experiencing violence, because
131 they remain in unstable settings within their countries of origin, and within families and
132 communities that experience disproportionate levels of distress during natural disasters and
133 displacement. A global meta-analysis of 56 mental health studies confirmed that IDPs tend to have
134 higher levels of psychopathology than refugees who left their country of origin ($Q = 65.47, R^2 =$
135 $0.05, p < 0.001$) (Porter & Haslam, 2005). Specific to Haiti, the prevalence estimates for
136 posttraumatic stress disorder (PTSD) and major depressive disorder two to four months after the

137 earthquake were 29.7% and 28.8% for IDPs, as compared to 19.1% and 21.9% for the general
138 population in Port-au-Prince ($p < 0.01$) (Cerdá et al., 2013).

139

140 Eliminating violence against children is a global commitment outlined in Sustainable Development
141 Goal (SDG) 16.2 and a public health priority of the World Health Organization (United Nations
142 General Assembly, 2015; World Health Assembly, 1996). Violence prevention and response in
143 humanitarian emergencies—caused by natural disasters, war, and mass population movement—is
144 key in achieving this goal. The United Nations Office for the Coordination of Humanitarian Affairs
145 (OCHA) estimated that more than 1% of the world’s population was affected by a humanitarian
146 emergency in 2017 (United Nations Office for the Coordination of Humanitarian Affairs, 2019).
147 Between 2008 and 2016, natural disasters displaced an average 25.3 million people per year, with
148 the largest proportion of people displaced within the borders of their home country (International
149 Displacement Monitoring Centre, 2017). These numbers are projected to increase in the future,
150 due to uneven population growth in the most affected areas and the increasing effect of climate
151 change on humanity (Peduzzi, Dao, Herold, & Mouton, 2009). Children are overrepresented in
152 humanitarian emergencies, representing 48 million of those who required assistance in 2017
153 (United Nations Children’s Fund, 2018). Most evidence that documents the relationship between
154 widescale catastrophic events and violence against children, however, comes from situations of
155 armed conflict, and less is known about possible differences in violence patterns after exposure
156 and displacement from natural disasters (Catani et al., 2010; Catani, Schauer, & Neuner, 2008;
157 Rubenstein, Lu, MacFarlane, & Stark, 2017; Stark, Warner, Lehmann, Boothby, & Ager, 2013).

158

159 Numerous methodological challenges exist in measurement of violence after natural disasters and
160 internal displacement. Prediction of the exact location and occurrence of natural disasters is
161 difficult (Gerstenberger, Wiemer, Jones, & Reasenber, 2005; Titov et al., 2005). Studies
162 commonly rely on the chance existence of prior data collection with the affected population and
163 less commonly, on a cross-sectional survey afterwards (Chen, Halliday, & Fan, 2016; Kolbe et al.,
164 2010). The former is inconsistently available, and the latter suffers from a lack of temporality in
165 understanding cause and effect, as correlations do not necessarily imply causation (Aldrich, 1995).
166 Observational studies used to study violence after population-based displacement events typically
167 face threats to validity because of confounding. The isolation of effect is complicated by an array
168 of social, economic, and behavioral factors that are correlated with violence outcomes (Doidge,
169 Higgins, Delfabbro, & Segal, 2017; Dong et al., 2004; Maguire-Jack & Font, 2017). Internal
170 displacement is usually influenced by having a lack of economic means, as the poorest of the poor
171 are both more likely to be displaced within their own country and spatially concentrate in remote
172 areas where instability and economic or environmental shocks frequently lead to displacement
173 (Cohen & Deng, 1998). In contrast, rapid-onset natural disasters like the Haitian earthquake act as
174 natural experiments in that for a brief moment, they quasi-randomly assign a large segment of the
175 population to internal displacement (Zubizarreta, Small, & Rosenbaum, 2014). The resulting
176 estimates are less influenced by other factors that would typically confound observational studies
177 (Zubizarreta, Cerdá, & Rosenbaum, 2013).

178

179 This study aims to analyze the effect of internal displacement from the Haitian earthquake on long-
180 term physical, emotional, and sexual violence against girls and boys. It secondarily lays out a

181 framework for certain data structures and analysis techniques to establish an empirical evidence
182 base for violence against children after rapid-onset natural disasters.

183

184 **Methods**

185 We used matching methods to pair individuals within exposed (displaced) and comparison groups
186 (non-displaced) who were similar on all observed pre-earthquake covariates within a nationally
187 representative cross-sectional survey for Haiti. We subsequently evaluated the relationship
188 between internal displacement from the earthquake and physical, emotional, and sexual violence
189 that occurred after the earthquake by gender. Our estimate of internal displacement from the
190 earthquake targeted the average treatment effect on the treated (ATT), which constitutes the
191 potential difference in violence outcomes for displaced people if they had not been displaced. The
192 matching procedure was designed to find suitable individuals in the comparison group who would
193 have survived the earthquake and not had the economic means or social connections to move
194 outside of the country.

195

196 **INSERT FIGURE 1**

197

198 **Data source**

199 The Haitian Violence Against Children Survey (VACS) is a nationally representative household
200 survey administered in Haitian Kreyol in 2012—two years following the Haitian earthquake.
201 Sampling methods stratified girls and boys into different clusters, providing representative
202 estimates by gender and a sample of IDPs in communities and camp settings. In total, 1,457 girls
203 and 1,459 boys completed the survey, with individual response rates of 93.1% for girls and 88.5%

204 for boys. Data collectors obtained informed consent from caregivers and assent from the child
205 respondents. The data collection methods and protocols are described in greater detail in the VACS
206 final report (Centers for Disease Control and Prevention, Interuniversity Institute for Research and
207 Development, & Comité de Coordination, 2014).

208

209 **Variable description**

210 The binary exposure variable was captured by asking respondents if they had moved or changed
211 households as a result of the earthquake. Earthquake exposure is decomposed into one element—
212 internal displacement to any location, including both camps and communities. Binary violence
213 outcomes were measured as experiences in the past 12 months. Since the survey was conducted
214 from April-June 2012, these violence outcomes had to occur after the earthquake. As per the
215 conventions of the VACS, the forms of violence were: (1) physical violence from parents,
216 caregivers, adult relatives, or other adult household members; (2) physical violence from authority
217 figures in the community; (3) emotional violence from parents, caregivers, adult relatives, or other
218 adult household members; and (4) sexual violence from anyone (see Supplementary File 1)
219 (Centers for Disease Control and Prevention et al., 2014).

220

221 We constructed pre-earthquake covariates for matching the exposed and comparison groups.
222 Limiting to pre-earthquake covariates importantly reduced the potential that measured
223 characteristics used in matching were derived from experiences during or after the earthquake. We
224 selected covariates in the survey that most strongly would confound the association between
225 earthquake exposure and violence against children and constructed dummy variables by
226 subtracting the respondent's current age at the date of the survey with the age of occurrence. A

227 timeframe of three or more years was deemed as occurring before the earthquake. We specifically
228 constructed covariates on experiences of physical, emotional, and sexual violence before the
229 earthquake to minimize possible confounding from past events (see Supplementary File 2).

230

231 **Study design and matching**

232 We restricted our analysis to the sample of 13 to 17 years old girls ($n = 635$) and boys ($n = 758$) in
233 each survey. As part of our study design, we assessed whether or not we had sufficient statistical
234 power to reliably detect the size of effects that we anticipated in our analysis (see Supplementary
235 Files 3-5). Data were pre-processed using the propensity score to trim individuals outside of the
236 area of common support—the area of overlap in which the exposed group has candidate
237 counterfactuals for matching (Rosenbaum, 2010).

238

239 We considered matching methods within the trimmed dataset to minimize the standardized mean
240 difference (SMD) of the pre-earthquake covariates between exposed and comparison groups,
241 blinding ourselves to violence outcomes until after settling on a particular matching design. We
242 prioritized matches that yielded a SMD within the range of +/- 0.10 (10 percent) (see
243 Supplementary File 6) (Normand et al., 2001). In recognizing the importance of certain covariates
244 in violence occurrence, we additionally prioritized the matching method that most reduced SMD
245 for pre-earthquake sexual violence for girls and physical violence in households for boys. We
246 chose these covariates, because they have been shown to be prevalent gendered forms of violence
247 in past studies (Stoltenborgh, van IJzendoorn, Euser, & Bakermans-Kranenburg, 2011), were
248 highlighted as most important for girls and boys in Haiti during a qualitative pre-study of the
249 VACS (Centers for Disease Control and Prevention & Interuniversity Institute for Research and

250 Development, 2011), and would likely be the most predictive of subsequent violence following
251 the earthquake. Using the outlined balance criteria, we selected a 1:1 Euclidean distance match
252 with a 0.2 caliper for both girls and boys. We implemented optimal matching using the `optmatch`
253 package (Hansen & Klopfer, 2006).

254

255 INSERT FIGURES 2 & 3

256

257 We used decision trees to articulate the implicit exclusion criteria that came about from trimming
258 and optimally matching, which is analogous to describing inclusion and exclusion criteria in a
259 randomized controlled trial (see Supplementary Files 7-8) (Traskin & Small, 2011).

260

261 **Statistical analysis**

262 We evaluated 95% confidence intervals (CI) from McNemar's exact test for matched pairs to
263 determine differences in violence outcomes after internal displacement from the earthquake and
264 considered *p*-values of less than 0.05 significant. In addition, we applied a multivariate analysis of
265 covariance (MANCOVA) with a Pillai test to explore patterns of violence. Differences in patterns
266 of missingness between exposed and comparison groups were assessed by using chained Fisher's
267 exact tests. We conducted gamma sensitivity analysis to determine the potential for unobserved
268 confounders to alter our observed results (Rosenbaum, 2010). We additionally ran power
269 calculations to determine if the sample size of the matched pairs was sufficient to detect changes
270 in violence after the earthquake, based upon pre-earthquake estimates of violence in Haiti. The
271 data were cleaned in Stata 15 and matched and analyzed in R v.3.3.3 (R Core Team, 2017;
272 StataCorp, 2017).

273

274

Results

275 Description of the study population

276 The majority of children living in camps in the study population were displaced by the Haitian
277 earthquake. As is the case for most IDPs globally, a sizable percentage of displaced children lived
278 outside of camps at the time of the survey (girls: 21.9% and boys: 20.8%) (United Nations &
279 European Union, 2018). Approximately half of girls (46.9%) and boys (54.1%) within the original
280 sample experienced some form of violence before the earthquake. The absolute number of violent
281 experiences by gender was similar after internal displacement from the earthquake (girls: 52.6%
282 and boys: 47.4%). The matched pairs mirrored the original sample, exhibiting high levels of
283 violence before and after the earthquake among girls and boys.

284 Table 1. Prevalence of violence before and after the Haitian earthquake in the matched pairs of girls and boys

	Girls				Boys			
	Before		After		Before		After	
	Estimate (n)	Percentage (%)						
Violence Type								
Physical violence from adults in household	120	39.2%	109	35.6%	151	43.9%	100	29.1%
Physical violence from authority figures	34	11.1%	41	13.4%	36	10.5%	59	17.2%

Emotional violence from adults in household	52	17.0%	94	30.7%	62	18.0%	62	18.0%
Sexual violence from anyone	31	10.1%	65	21.2%	24	7.0%	34	9.9%
Any form of violence	151	49.3%	172	56.2%	182	52.9%	157	45.6%

286 **Results of study design**

287 The matching yielded 153 pairs of girls and 172 pairs of boys (see Supplementary Files 9-10). Our
288 analysis was sufficiently powered to detect changes comparable to Demographic Health Survey
289 (DHS) estimates on national violence prevalence in Haiti before the earthquake (Cohen's $h = 0.20$)
290 (Cayemittes, Placide, Barrère, Mariko, & Sévère, 2001; Cayemittes et al., 2007).

291

292 **Violence against girls after internal displacement**

293 The odds ratios associating internal displacement from the earthquake with long-term physical,
294 emotional, or sexual violence were near null for girls. Multivariate analysis combining violence
295 outcomes followed a similar pattern, with no appreciable difference between exposed and
296 comparison groups ($F = 2.11$, $df = 4$, 294 , $p = 0.080$). The exposed and comparison groups did not
297 have substantial differences in their patterns of missingness for any form of violence. Sensitivity
298 analysis showed that sexual violence outcomes were highly sensitive to possible bias from
299 unmeasured confounding (the presence of $\Gamma = 1.1$ magnitude of bias could give rise to a connection
300 between internal displacement and violence; $p = 0.048$) (see Supplementary File 11, Table S6).

301

302 **Violence against boys after internal displacement**

303 The odds ratios associating internal displacement from the earthquake with long-term physical,
304 emotional, or sexual violence were similarly near null for boys. As in the case of girls, multivariate
305 analysis did not exhibit a different pattern in violence outcomes when combined ($F = 0.98$, $df = 4$,
306 324 , $p = 0.417$). The exposed and comparison groups likewise did not show evidence of differences
307 in their patterns of missingness. Sensitivity analysis illustrated that physical violence perpetrated
308 by authority figures was moderately sensitive to possible bias from unmeasured confounding (the

309 presence of $I = 1.2$ magnitude of bias could give rise to a connection between internal
 310 displacement and violence; $p = 0.045$) (see Supplementary File 11, Table S7).

311
 312 Table 2. Differences in violence outcomes after the Haitian earthquake for displaced and non-
 313 displaced girls and boys

Violence Type	Girls			Boys		
	Estimate	95% CI	<i>p</i> -value	Estimate	95% CI	<i>p</i> -value
Physical violence by adults in household	0.90	0.52-1.56	0.795	1.35	0.83-2.23	0.242
Physical violence by authority figures	1.67	0.85-3.40	0.154	0.87	0.38-1.95	0.851
Emotional violence by adults in household	1.11	0.57-2.17	0.875	1.39	0.79-2.49	0.281
Sexual violence by anyone	1.29	0.60-2.79	0.597	0.57	0.29-1.09	0.096
Any form of violence	0.84	0.52-1.33	0.500	1.03	0.61-1.73	1.000

314 Estimate = odds ratio; rounded to two decimal places

315
 316 **Discussion**
 317 We found that internal displacement was not associated with past-12 month physical, emotional,
 318 or sexual violence against girls or boys within affected households and caregiving networks two
 319 years after the earthquake in Haiti. Our findings present a hopeful picture that internal

320 displacement from the earthquake was not a driving factor of long-term violence against children
321 in Haitian society. The VACS sampling structure and our implicit inclusion criteria suggest that
322 we are able to generalize to all Haitian children who would have been internally displaced by the
323 earthquake and who did not have the economic means or social connections to move elsewhere.
324 The results therefore provide representative estimates for the affected population of children that
325 remained in Haiti after the earthquake. Sensitivity analysis indicates that the results are resistant
326 to high levels of possible unobserved biases, with the exception of sexual violence against girls
327 and physical violence perpetrated by authority figures against boys. We can have confidence in
328 the majority of our findings, but we must interpret the results for these two gendered forms of
329 violence with caution in light of the prospect that bias exists and was not measured in the survey.

330

331 Prior studies that investigate violence against children after disaster and displacement exposure
332 have had mixed findings. A recent meta-analysis found that natural disasters were not associated
333 with physical, emotional, and sexual violence against children (Cerna-Turoff, Fischer, Mayhew,
334 & Devries, 2019). Specific analyses that isolated the mediated effect of internal displacement as
335 distinct from natural disaster exposure were noticeably absent. The most relevant included study
336 on Hurricane Ike in the United States concluded that boys who were not evacuated were more
337 likely to perpetrate physical dating violence (aOR 3.19, 95% CI 1.50-6.80, $p < 0.01$) and perpetrate
338 or be victims of sexual violence (perpetration: aOR 3.73, 95% CI 1.50-9.28, $p < 0.01$;
339 victimization: aOR 2.47, 95% CI 1.17-5.23, $p < 0.05$) (Temple et al., 2011). Among adult women,
340 residing in an IDP camp increased the probability of physical and sexual violence but not general
341 displacement due to property destruction or loss from the Haitian earthquake (Weitzman &
342 Behrman, 2016).

343

344 Temporal trends are similarly unclear. Two studies from the United States point to an initial
345 increase in the first six months after several natural disasters, which decreases over time (Curtis,
346 Miller, & Berry, 2000; Keenan, Marshall, Nocera, & Runyan, 2004). Kolbe (2010) identified a
347 large number of new cases of sexual violence against girls in Port-au-Prince in the six weeks after
348 the Haitian earthquake, and Weitzman and Behrman (2016) found that the probability of physical
349 and sexual intimate partner violence increased among adult Haitian women in the two years
350 following the earthquake. In contrast, a study on physical, emotional, and sexual violence against
351 internally displaced girls in camps one to three years after the earthquake in Haiti found no
352 association, although the study faced several methodological challenges arising from limited
353 sample sizes ($N = 78$) and incomplete use of validated violence scales for measurement (Sloand et
354 al., 2017). The 2012 DHS similarly found a lower prevalence of physical and sexual intimate
355 partner violence against internally displaced adolescent girls in camps than the general population
356 after the earthquake (Cayemittes et al., 2013).

357

358 **Limitations and strengths**

359 Our findings must be considered in conjunction with the study's limitations. We did not have in-
360 depth information to decompose earthquake exposure into other distinct elements and test their
361 effect on violence. Ideally, we would have had access to the subgroup of people who had been
362 displaced to a camp or informal settlement but currently resided in the wider community. The
363 CDC's initial analysis reported that Haitian girls and young women who lived in the camps at the
364 time of the survey had a higher probability of experiencing post-earthquake sexual violence
365 (Centers for Disease Control and Prevention & Interuniversity Institute for Research and

366 Development, 2011). Their analysis did not adjust for pre-earthquake violence, measure the effect
367 of displacement on violence, or compare the estimates with the population of IDPs who no longer
368 lived in camps. A related limitation is that we did not have a measure of earthquake intensity that
369 could have been used to partition exposure into any unexplained deviation from randomness.
370 While an instrumental variable analysis would have offered an additional mechanism to reduce
371 possible biases in exposure, displacement as a clear binary question is less sensitive to unmeasured
372 biases than incremental changes in earthquake intensity (Rosenbaum, 2004). This point may be
373 salient in terms of a related national study on intimate partner violence against adult women in
374 Haiti, which used three gradations of earthquake intensity to determine exposure (Weitzman &
375 Behrman, 2016).

376

377 We could not include certain covariates in the matching metric that are typically associated with
378 violence in other settings, namely pre-earthquake place of residence and socio-economic status
379 (Willman & Marcelin, 2010). Geographic residence in where the children lived before the
380 earthquake may have biased this study's conclusions towards the null. Available data indicates
381 that Port-au-Prince had a lower corporal punishment prevalence than in rural or other urban areas
382 of Haiti before the earthquake (Cayemittes et al., 2001). In contrast, poverty would have biased
383 away from the null in this sample, given that poverty is a risk factor for violence (Maguire-Jack &
384 Font, 2017). The temporary and powerful mass disruption of the Haitian earthquake on the
385 population-level likely helped to reduce both of their influences in this analysis (Zubizarreta et al.,
386 2014).

387

388 We likewise did not have sufficient information to match on frequency and severity of violence
389 before the earthquake or protective factors. Overall characteristics for girls and boys who
390 experienced violence once may differ from those who frequently experience violence. Matching
391 methods, however, are not meant to pinpoint the exact same type of person for each characteristic
392 but rather, to construct “profiles of risk” for similar people across exposed and comparison groups
393 (Rosenbaum & Rubin, 1983). We additionally were able to determine that no child was in a
394 marriage-like relationship before the earthquake, reducing the possibility of ongoing intimate
395 partner violence. Certain groups in Haiti may be resilient to the stress-trauma pathway and exhibit
396 low rates of violence against children after a large-scale disaster and internal displacement. In other
397 contexts, people who reported high social cohesion and a lack of racial discrimination before
398 Hurricane Katrina and high social capital before a natural disaster in Japan had lower trauma
399 responses (Tsuchiya et al., 2017; Weems et al., 2007). These covariates and other possibly relevant
400 characteristics were not measured in the VACS, which has the potential to bias estimates. We
401 tested the strength of our findings in sensitivity analysis for this reason. Apart from physical
402 violence committed by authority figures against boys and sexual violence against girls, a high
403 amount of possible bias introduced by unobserved confounders would be needed to change our
404 study results.

405

406 We could not track short-term temporal changes in violence after the earthquake. Data collection
407 for the Haiti VACS was conducted 14 to 16 months after the earthquake. The violence measures
408 were assessed for the time period of 12 months before the survey (Centers for Disease Control and
409 Prevention et al., 2014). Therefore, the immediate two to four months following the earthquake
410 were not included in this analysis. Violence may have increased in the initial aftermath of the

411 earthquake, but this study was designed to understand violence sustained over a different
412 timescale.

413

414 The current study is not well-designed to identify issues arising from spillover effects which
415 violate the Stable Unit Treatment Value Assumption (SUTVA). An example would be if the influx
416 of internally displaced children to a non-earthquake region increased the rates of violence against
417 non-displaced girls and boys. This possibility would tend to bias the current study design toward
418 finding false null results. Spillover effects are quite likely in rapid-onset natural disasters but
419 require measurement of spillover pathways not collected in this survey and sophisticated methods
420 that to our knowledge have not been implemented in population-based surveys in humanitarian
421 emergencies.

422

423 This study had multiple strengths notwithstanding its limitations. We had access to a nationally
424 representative survey that extensively measured physical, emotional, and sexual violence against
425 children, with the appropriate data structure to create pre-earthquake covariates and identify long-
426 term gendered effects. We then applied an experimental approach to reduce potential biases. The
427 sample size used for our analysis was sufficiently powered to detect small changes in violence
428 outcomes (Cohen's $h = 0.20$). The survey question linked to the exposure variable—self-reported
429 internal displacement because of the earthquake—was not likely affected by recall bias among the
430 sampled children. Most IDPs in Haiti furthermore remained in country after the earthquake
431 (Bengtsson, Lu, Thorson, Garfield, & von Schreeb, 2011). Statistical analysis did not indicate that
432 missingness in children's response patterns was likely to change our findings. In addition, the
433 discarded observations from the comparison group did not appear to contain an important

434 subpopulation of children who were at elevated risk for violence. Multivariate analysis further did
435 not find that physical, emotional, and sexual violence covaried, which suggests that internal
436 displacement did not act in a joint manner on violence outcomes, apart from each individual effect.

437

438

Conclusions and future implications

439 This study contributes to increased knowledge of violence within disaster- and displacement-
440 affected populations and can inform policies and service provision towards global priorities to end
441 violence against children. Current global operating protocols in emergencies call for initiation of
442 violence prevention and response services months after the disaster event and displacement, based
443 upon a theory of sustained increases in violence (Inter-Agency Standing Committee, 2006). In
444 settings with “acute on chronic” underdevelopment and repeat episodic natural disasters, like Haiti,
445 the affected population may possess a high-level of resilience to recover from traumatic
446 experiences quickly (Gabrielli & Gill, 2014). We may be missing the optimal window of time to
447 respond to disaster and displacement exposure with targeted interventions to protect children.

448

449 The experience of internal displacement on the population level may be less important in
450 determining long-term violence outcomes than the location of displacement. Internal displacement
451 to a camp or informal settlement particularly may exacerbate risk, as indicated in a study of the
452 impact of the Haitian earthquake on violence against adult women (Weitzman & Behrman, 2016).
453 Better documentation of changes in violence over time, protective factors, subgroup differences,
454 and other aspects of the exposure and displacement experience would provide a more nuanced
455 understanding of violence patterns among children. Further study is merited specifically on sexual
456 violence against girls and physical violence perpetrated by authority figures against boys due to

457 their sensitivity to potential unmeasured biases. Moreover, internal displacement due to natural
458 disasters may affect populations differently than other humanitarian emergencies, such as armed
459 conflict, and violence patterns may diverge as a result (Catani, Jacob, et al., 2008; Norris et al.,
460 2002; Rubenstein et al., 2017). Data collection that can account for pre-disaster characteristics
461 would aid in confirming if these dynamics are consistent across disaster settings and in countries
462 that experience overlapping cycles of man-made violence and political instability (Hallward,
463 2010).

464

465 Natural disasters are often unpredictable events and therefore, present challenges in designing
466 studies which include a baseline when one does not know when an event will occur and who will
467 be affected (Gerstenberger et al., 2005; Titov et al., 2005). The analysis of a single, representative
468 cross-sectional survey of the affected and unaffected population is logistically more feasible than
469 pre-post studies, given the costs and time required to trace the affected population. This approach
470 is especially relevant to low- and middle-income settings where infrastructure and surveillance
471 systems are incomplete or weak (Galea, Maxwell, & Norris, 2008). Rapid-onset natural disasters
472 are specifically suited for this method, because they act as population randomizers, and unlike
473 armed conflict, may exhibit less strong spatial patterns (Guo, Lu, Doñate, & Johnson, 2017). By
474 creating pre-earthquake covariates and analyzing data with matching methods, we gained some of
475 the benefits of pre-post design in a low-resource setting and reduced the confounding inherent in
476 observational studies (Rubin, 2005).

477

478

Figure captions

479 Fig 1. Mapping of shake intensity by communes. Data source is the United States Geological
480 Survey and map produced in the `mapview` package (Appelhans, Detsch, Reudenbach, &
481 Woellauer, 2018; United States Geological Survey, n.d.)

482 Fig 2. Love plot of pre-earthquake covariate balance before and after matching for girls

483 Fig 3. Love plot of pre-earthquake covariate balance before and after matching for boys

484

485

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687

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694

695 **Declaration of interest**

696 None.