

1 Supplement for “Serostatus Testing & Dengue Vaccine Cost-Benefit Thresholds”

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80 population heterogeneity in exposure risk). 15

S1. ASSUMPTIONS

81

82 Based on our understanding of how CYD-TDV works in transmission settings of interest, we model the expected
 83 life time benefits and costs of vaccination interventions versus status quo on an individual basis. This framing implies
 84 that **interventions have no effect beyond the individual**, which effectively ignores any changes to transmission.
 85 Similarly, all **benefits and cost are on the margin**, so for example expenditure on testing and vaccination is
 86 independent of coverage level. However, economies of scale could be incorporated by, say, adjusting the vaccine or
 87 test prices to account for lower unit costs.

88 Our model is on a **discrete, annual time scale**. Practically, this means representing all dengue natural history
 89 effects as discrete-year effects (*e.g.* transient cross-serotype immunity). Likewise, interventions only occur once a
 90 year, and all data is interpreted at this yearly scale.

91 We assume **dengue is an environmental risk**, rather than a dynamically spread infection, with multiple, but
 92 indistinguishable, serotypes. Combined with the annual time scale, this means we represent dengue exposure risk as a
 93 probability per year-of-life effect; see SI Section *Exposure Model* for details. Since we assume the **dengue serotypes**
 94 **are indistinguishable**, they have identical infection probability, disease risk, circulation probability, *etc.* We assume
 95 that any particular serotype is life-time immunizing against that serotype, and that infection in one year precludes
 96 infection by any serotype in the next year. Finally, we assume no direct age-dependence for any aspect of dengue
 97 disease.

98 Using this model of dengue exposure, we pre-compute sample life trajectories for 1 million individuals (to ensure
 99 relatively smooth estimates), all living to age 70. We did not explicitly explore sensitivity to this life expectancy
 100 (either mean or distribution), but the qualitative trends can be determined from reasoning: increase in life expectancy
 101 increases the intervention value (and *vice versa*), as more people are likely to have multiple life infections, and any
 102 increase in single infections is well outside the test and vaccinate window. However, this effect applies to such a small
 103 proportion—those with a single infection by or during the test-then-vaccinate window, on the margin of having one
 104 or two total life time infections that gain or lose an infection by extending or shortening life decades later—and so we
 105 expect it to have little impact on any of these assessments.

106 For dengue disease natural history, we assume only the first and second infections have potentially associated
 107 disease.

108 We assume that, aside from infection-history-derived disease outcomes, vaccination has no impact on infections
 109 over life. That is, the transient immunity observed during CYD-TDV trials does not change the number of lifetime
 110 infections any individual experiences. We justify this based on the relatively short duration of immunity compared
 111 to life expectancy. By making this assumption, we can use pre-sampled life infection trajectories without having to
 112 perturb them by vaccination. We also assume the vaccination perfectly replaces one natural infection; this is equivalent
 113 to assuming it is perfectly efficacious against disease in seropositive recipients. Estimates indicate that seropositive
 114 efficacy is more like 70-80% [1]. We used an optimistic assumption because we are computing a threshold. Additionally,
 115 assuming a leaky vaccine would be incompatible with pre-computing life histories. However, the derivation that follows
 116 could accommodate a non-leaky, imperfect vaccine efficacy.

117 Finally, we assume tests are perfectly sensitive and specific. This provides an upper bound on benefit, which can
 118 be used to preclude combinations of strategy and price for vaccines and tests, irrespective of test performance. The
 119 detailed cost performance of real tests will vary by setting, test properties, and intervention approach. We assume
 120 any extant individual test history (*e.g.* previous lab confirmed dengue infection) is irrelevant to this strategy on a
 121 longer term basis, and so we ignore that testing; going forward, any non-vaccine related testing could be reasonably
 122 considered part of our cost estimate.

123 On balance, these assumptions tend to overvalue the intervention. We judge that the net effect is likely small
 124 relative to estimating limiting space of a test-then-vaccinate regimen for the yes or no decision many settings now
 125 face for CYD-TDV. For those settings where the prices are in (or close enough to) the beneficial space, this type of
 126 analysis should be revisited with more modelling detail.

127

S2. COST-BENEFIT EQUATIONS

128

A. Definitions

129 Our model represents the following costs:

- 130 • F, S : the average individual costs of first-like and second-like infections, respectively. These should be based on
 131 local economic data, *e.g.* hospital costs, lost productivity, willingness to pay, but like most economic decisions
 132 will have a subjective element

- 133 • V, T : the unit cost of vaccine and testing regimens; these values are to be constrained, or have proposed prices
134 to be evaluated against those constraints
- 135 • $\nu = \frac{V}{S}, \tau = \frac{T}{S}$: the costs of the vaccine and test as a fraction of the cost of second-like infection; deriving equations
136 in these terms allows general results
- 137 • C_X , the total individual cost for scenario X (no intervention, vaccination without testing, *etc.*)
- 138 • Δ_X , the difference between status quo (C_0 ; no vaccination) and intervention X : $\Delta_X = C_0 - C_X$. For an
139 intervention to have net benefit, $\Delta_X \geq 0$ must be true

140 Note, that these costs may represent any perspective (individual, societal, *etc.*), but should be from a consistent
141 perspective.

142 The population is stratified by total lifetime infections, as well as number of infections at the routine testing age
143 (or vaccination age, for the comparison scenario without testing). These proportions in turn determine the costs and
144 benefits of a particular vaccination and testing strategy. We represent these proportions with the following general
145 constructions:

- 146 • \mathcal{P}_X : proportion having X infections during life; this value is age independent
- 147 • ${}_N\mathcal{P}_X\{A\}$, proportion having N infections prior to consideration for vaccination at age A , and X total lifetime
148 infections
- 149 • N and X can include modifiers, like X^+ to mean X or more, or have \forall to mean any number of total infections
150 (by definition, $X \geq N$).
- 151 • $\mathcal{C}\{A\}$, the conditional probability of seroconverting between age A and $A + 1$, given seronegativity at age A .

152 In general, we drop A , as it is the same for all terms in most equations. We only use it explicitly for multiple testing
153 scenarios. Here are some example proportions which are relevant to our final derivation results:

154 The proportion of individuals that

- 155 • \mathcal{P}_0 : ... are lifetime seronegative
- 156 • \mathcal{P}_{1+} : ... have one or more lifetime infections
- 157 • ${}_0\mathcal{P}_{2+}, {}_1\mathcal{P}_{2+}$: ... have no or one infection, respectively, prior to routine age, and have two or more infections over
158 life
- 159 • ${}_2\mathcal{P}_\forall$: ... have two or more infections at the time of vaccination
- 160 • ${}_1\mathcal{P}_\forall, {}_{1+}\mathcal{P}_\forall$: ... have exactly one or at least one infection at the time of vaccination

161 In a later section, we will define an exposure model to estimate these probabilities, but we may take the specifics
162 of that for granted while we layout the cost relationships. Because the cost model is defined only in terms of the
163 probabilities, we could replace the exposure model with another approach that provides the same probabilities, as
164 long as it addressed other pertinent assumptions, like the independence of individuals.

165 B. Status Quo and Vaccination Only Costs (C_0, C_V)

166 Without vaccination, an individual's expected lifetime dengue burden would be:

$$C_0 = (1 - \mathcal{P}_0)F + \mathcal{P}_{2+}S \quad (1)$$

167 Note that the proportions are all independent of age, and only concern lifetime outcomes. When considering
168 interventions, age becomes a factor (implicit with the ${}_X\mathcal{P}_N$ terms, which are a function of A). With universal
169 vaccination (*i.e.*, irrespective of serostatus), the cost would be:

$$C_V = {}_1\mathcal{P}_\forall F + {}_{2+}\mathcal{P}_\forall(F + S) + V + {}_0\mathcal{P}_{1+}S \quad (2)$$

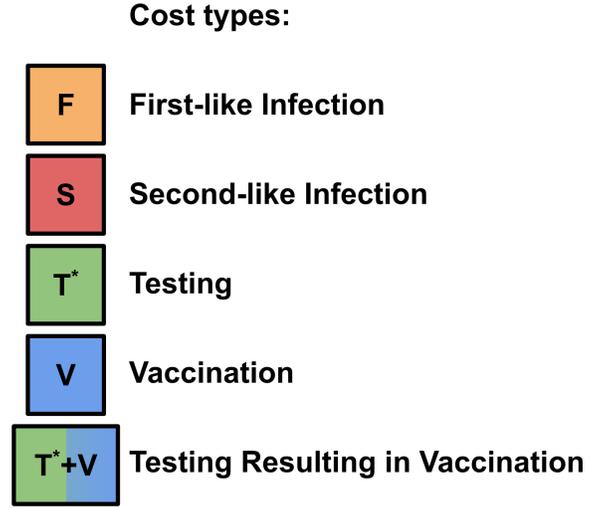


FIG. S1. **Diagram Key.** Legend for the diagrams showing life time trajectories for various intervention scenarios.

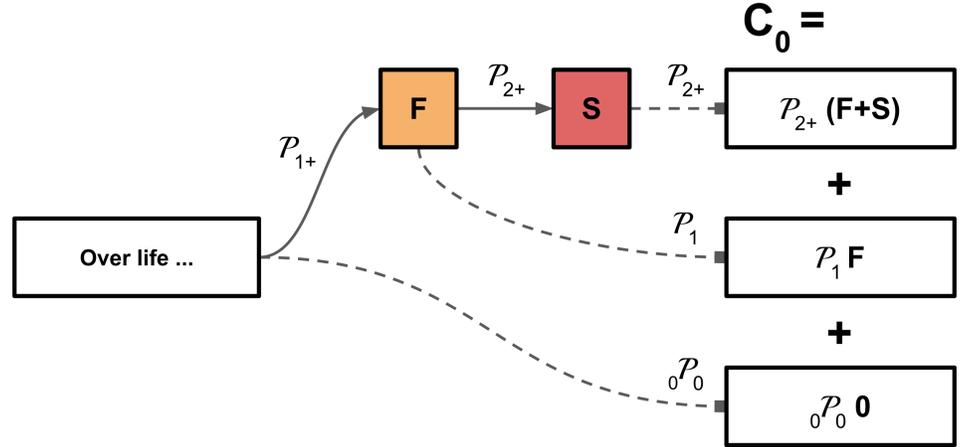


FIG. S2. **Outcomes for Status Quo.** Pathways show the probability of various life trajectories and ultimate costs for Eq. 1.

170

C. Model with Free Testing

171 To begin understanding the cost model with testing, we imagine combining the vaccine with a free test (denoted
 172 V'); this assumption is obviously ludicrous—the test will cost resources to produce and use—but provides insight
 173 about the limits of vaccine value for an epidemiological setting.

174 We assume this imaginary test identifies the *number* of past dengue infections, so we refer to it as the *ordinal*
 175 test. This test enables us to only vaccinate seropositives that *could* have another disease-bearing infection, and avoid
 176 vaccinating seronegatives until they seroconvert. We use the test every year from the routine testing age, until the
 177 individual receives the vaccine. The cost for this hypothetical regimen is:

$$C_{V'} = (1 - \mathcal{P}_0)F + (1 - \mathcal{P}_0 - {}_{2+}\mathcal{P}_V)V' + {}_{2+}\mathcal{P}_V S \quad (3)$$

178 Note that V' vaccination only occurs for seropositives that have only had one infection and individuals that se-
 179 roconvert after the routine testing age; because we assume no benefit against third and later infections, we do not
 180 administer vaccine to those with multiple past infections. If we have a test that only detects seropositivity but not
 181 detailed infection history, which we refer to as a *binary* test, then costs would be:

$$C_{V^\dagger} = (1 - \mathcal{P}_0)F + (1 - \mathcal{P}_0)V^\dagger + {}_{2+}\mathcal{P}_V S \quad (4)$$

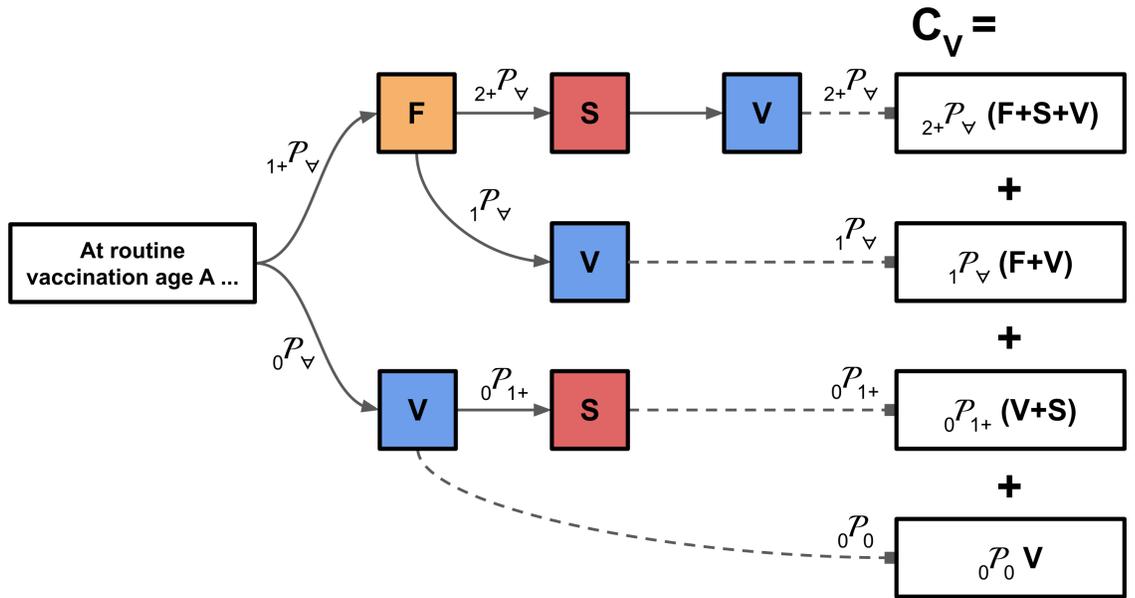


FIG. S3. **Outcomes for Unconditional Vaccination.** Pathways show the probability of various life trajectories and ultimate costs for Eq. 2. Compared to Fig. S2, individuals that will experience infections over life are now either vaccinated before or after their first infections, meaning they may have a first-like infection converted to a second-like infection, or possibly avoid second-like infections. The vaccination may also be wasted, if given to individuals having already experienced multiple infections.

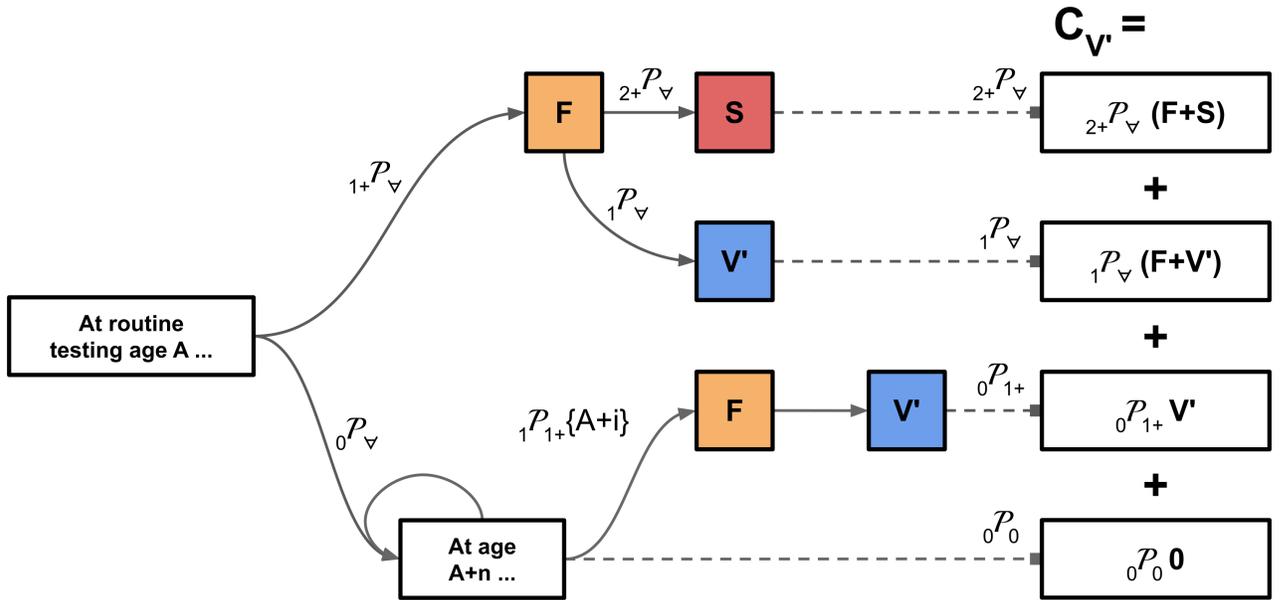


FIG. S4. **Outcomes for Vaccination with a Free Ordinal Test.** Pathways show the probability of various life trajectories and ultimate costs for Eq. 3. Compared to Fig. S3, we never induced second-like infections, and indeed avoid all of them that would have occurred after the routine testing age.

182 which we distinguish by the V^\dagger term. The only change is on the coefficient for the V term (and the $\nu = \frac{V}{S}$ terms,
183 when we switch to non-dimensional form); this result will be consistent as we consider more complicated models.

184 D. Cost-Benefit Constraint Equations for Vaccination-Only and Free Testing Models (C_V , $C_{V'}$, & C_{V^\dagger})

185 If we compare all these testing scenarios to non-vaccination:

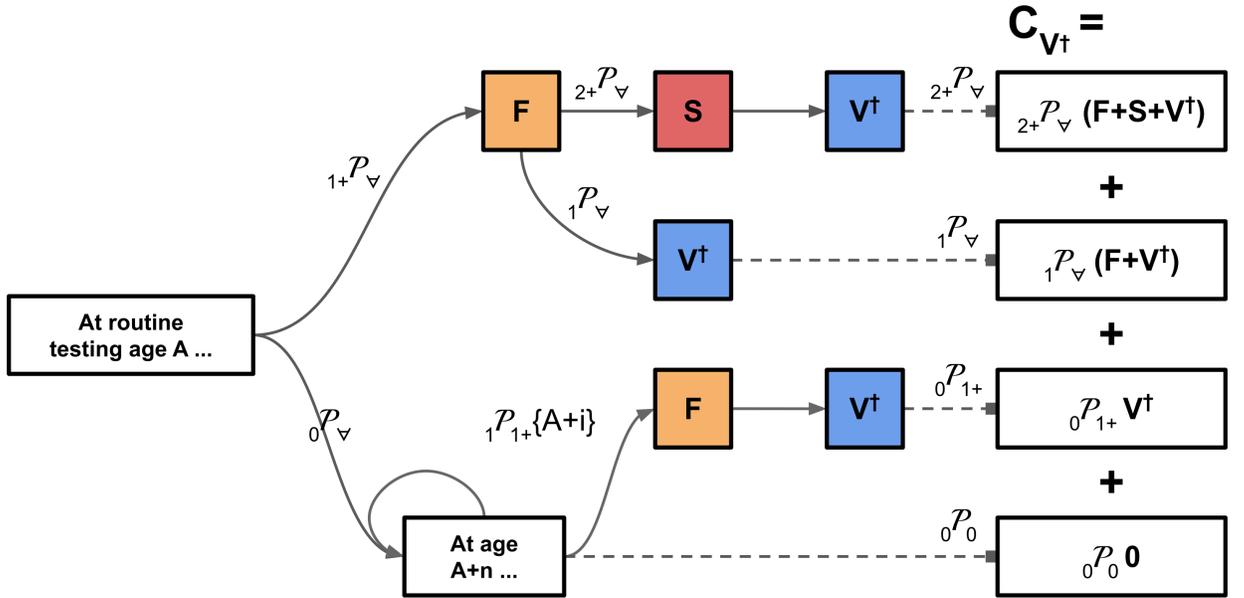


FIG. S5. **Outcomes for Vaccination with a Free Binary Test.** Pathways show the probability of various life trajectories and ultimate costs for Eq. 4. Note that the only distinction from Fig. S4 is the additional vaccination along the uppermost path, the trajectory corresponding to two or more infections prior to the routine vaccination age. While only a single term, this can make vaccination substantially more expensive in settings where multiple infections are frequent prior to the routine testing age.

$$\begin{aligned} \Delta_V = C_0 - C_V &= (1 - \mathcal{P}_0 - {}_1\mathcal{P}_v - {}_2+\mathcal{P}_v)F + (\mathcal{P}_{2+} - {}_0\mathcal{P}_{1+} - {}_2+\mathcal{P}_v)S - V \\ &= (\mathcal{P}_{1+} - {}_1\mathcal{P}_v - {}_2+\mathcal{P}_v)F + ({}_0\mathcal{P}_{2+} - {}_0\mathcal{P}_{1+} + {}_1\mathcal{P}_{2+} + {}_2+\mathcal{P}_{2+} - {}_2+\mathcal{P}_v)S - V \\ &= {}_0\mathcal{P}_{1+}F + ({}_1\mathcal{P}_{2+} - {}_0\mathcal{P}_1)S - V \end{aligned} \quad (5)$$

$$\Delta_{V'} = (\mathcal{P}_{2+} - {}_2+\mathcal{P}_v)S - (1 - \mathcal{P}_0 - {}_2+\mathcal{P}_v)V' \quad (6)$$

$$\Delta_{V^\dagger} = (\mathcal{P}_{2+} - {}_2+\mathcal{P}_v)S - (1 - \mathcal{P}_0)V^\dagger \quad (7)$$

186 Recall that by definition, vaccination without testing is only net beneficial if $\Delta_V \geq 0$. Testing can expand the region
187 of vaccine benefit, but only if $\Delta_{V'} \geq 0$ or $\Delta_{V^\dagger} \geq 0$, depending on the test mechanism. For routine vaccination at a
188 certain age to be net beneficial after the introduction of any real testing cost, the vaccine regimen cost must obey:

$$\nu' < \frac{\mathcal{P}_{2+} - {}_2+\mathcal{P}_v}{1 - \mathcal{P}_0 - {}_2+\mathcal{P}_v} \quad (8)$$

$$\nu^\dagger < \frac{\mathcal{P}_{2+} - {}_2+\mathcal{P}_v}{1 - \mathcal{P}_0} \quad (9)$$

189 Note that combined with either of the free tests, the constraint on vaccine cost depends only on three generic
190 parameters of the context (the average cost of second-like infections (S , implicit within ν), and the lifetime probabilities
191 of 0 or 2^+ infections) and one parameter of the intervention (*i.e.*, age A of consideration for vaccination, which
192 determines the fraction of the population that has had two or more infections). As the routine
193 age is shifted earlier, the probability of two infections prior to consideration for vaccination becomes vanishingly small,
194 *i.e.* ${}_2+\mathcal{P}_v \rightarrow 0$. Therefore, at a young enough vaccine age the constraint equations converge to the ratio

$$\nu' = \nu^\dagger \leq \frac{\mathcal{P}_{2+}}{1 - \mathcal{P}_0} \quad (10)$$

195 or the fraction of infectees that will have multiple infections out of all infectees. We could have arrived at this result
196 reasoning from event probabilities: an individual only pays the vaccine cost if they experience at least one infections

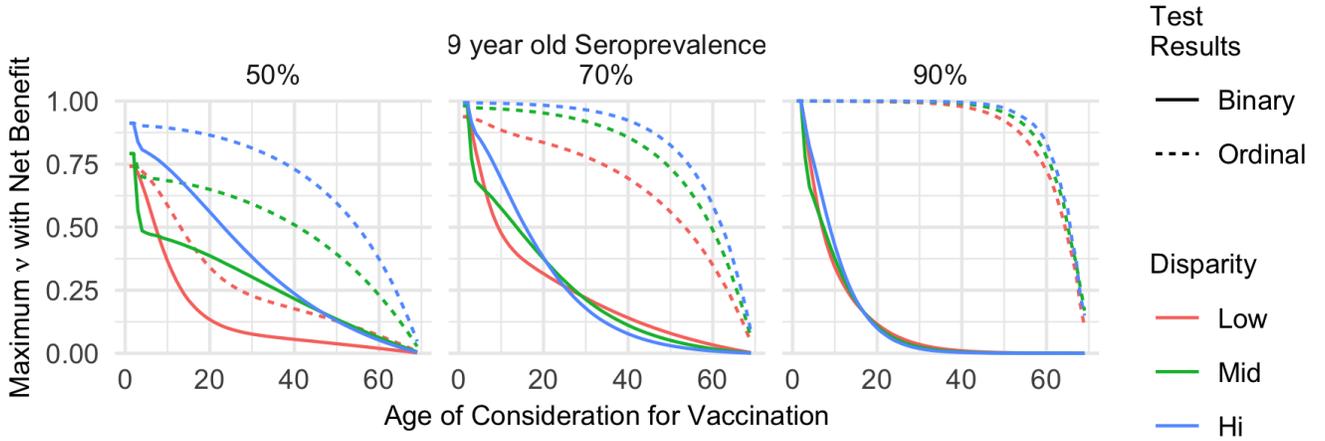


FIG. S6. **Maximum Vaccine Cost Fraction Allowing Net Benefit.** The approximate maxima for ν based on Eq. 8-9, including their convergence at younger ages. These curves represent the thresholds for vaccine costs given *real* tests, *i.e.* non-free with less than perfect sensitivity and specificity. Note that these curves only represent the restriction on ν for *some* benefit; the amount of benefit is generally shrinking faster than the cap on vaccine cost, particularly for the ordinal test. We facet by epidemiological parameters here: seroprevalence in 9-year-olds (a surrogate for dengue transmission level) and disparity (a measure of risk heterogeneity).

197 $(1 - \mathcal{P}_0)$, and only individuals that experience multiple infections benefit (by avoiding S). Therefore, the maximum
 198 value of V is the conditional probability of experiencing S given any infections.

199 Note that $\nu' \geq \nu^\dagger$ for any routine testing age, but the value of ν^\dagger grows faster with decreasing age (as it must for
 200 them to converge).

201

E. Models with Non-Free Testing ($C_{VT'}$, C_{VT^\dagger})

202 If we now acknowledge the cost of testing, T , and consider only a single test (instead of potentially testing several
 203 times), the individual costs are:

$$C_{VT'} = T' + (1 - \mathcal{P}_0)F + ({}_0\mathcal{P}_{2^+} + {}_{2^+}\mathcal{P}_V)S + {}_1\mathcal{P}_V V \quad (11)$$

$$C_{VT^\dagger} = T^\dagger + (1 - \mathcal{P}_0)F + ({}_0\mathcal{P}_{2^+} + {}_{2^+}\mathcal{P}_V)S + {}_1\mathcal{P}_V V \quad (12)$$

204 and the intervention benefits are:

$$\begin{aligned} \Delta_{VT'} &= (\mathcal{P}_{2^+} - {}_0\mathcal{P}_{2^+} - {}_{2^+}\mathcal{P}_V)S - T' - {}_1\mathcal{P}_V V \\ &= \mathcal{P}_{2^+1}\mathcal{P}_{2^+}S - T' - {}_1\mathcal{P}_V V \end{aligned} \quad (13)$$

$$\Delta_{VT^\dagger} = \mathcal{P}_{2^+1}\mathcal{P}_{2^+}S - T^\dagger - {}_1\mathcal{P}_V V \quad (14)$$

205 Which makes our non-free, single testing constraints for cost effectiveness:

$$\tau' + {}_1\mathcal{P}_V \nu \leq \mathcal{P}_{2^+1}\mathcal{P}_{2^+} \quad (15)$$

$$\tau^\dagger + {}_1\mathcal{P}_V \nu \leq \mathcal{P}_{2^+1}\mathcal{P}_{2^+} \quad (16)$$

206 Notably, there is still no dependence on the cost of first infections (F).

207

F. Multiple Testing ($C_{VLT'}$, C_{VLT^\dagger})

208 Now consider if we allow testing every year, up to a maximum of L tests (including the first). Like the free test
 209 intervention, we may identify people who seroconvert after initial consideration for vaccination, though since there is
 210 a test cost, there will be diminishing returns (and eventually losses) with repeat testing.

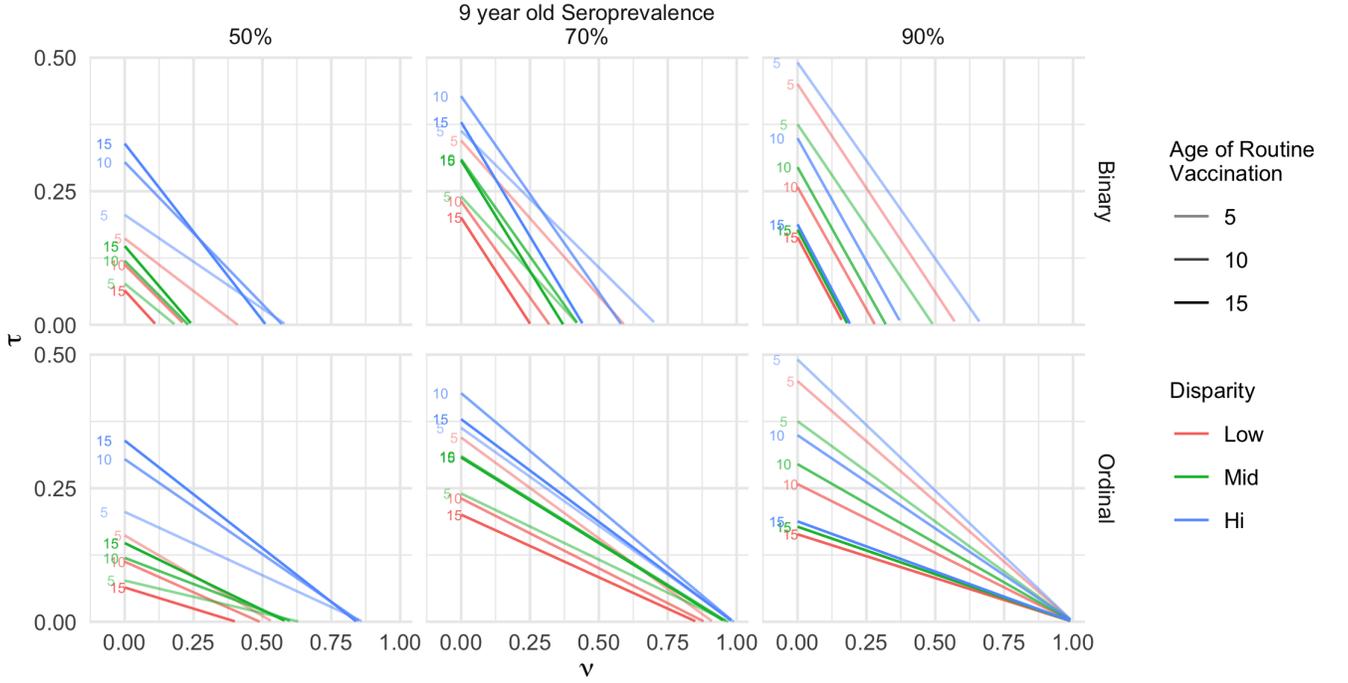


FIG. S7. **Vaccine and Single Test Cost Fraction Boundary.** The approximate boundary for ν and τ based on Eq. 15-16. To be net beneficial, the costs of testing and vaccine (under a single test strategy) must result in a point below the relevant line (based on routine vaccination age and context). We show how the threshold line shifts with age. We facet by environmental sensitivity parameters here: seroprevalence in 9-year-olds (a surrogate for dengue transmission level) and disparity (a measure of risk heterogeneity).

211 Relative to C_{VT} or C_{VT^+} , lifetime seronegatives (\mathcal{P}_0) will pay for $L - 1$ additional tests. The fraction of the
 212 population in ${}_0\mathcal{P}_{1+}$ will pay for some number of additional tests, and possibly vaccination (*i.e.*, for the proportion of
 213 individuals seroconverting during the testing window). This means costs for this strategy will change some ${}_0\mathcal{P}_{1+}\{A\}$
 214 (*i.e.*, seronegative at A , the routine initial testing age) to ${}_1\mathcal{P}_{1+}\{A+n\}$ at some later age prior to $A + L - 1$ (*i.e.*,
 215 seroconverting during the testing window, and therefore vaccinated).

216 Relative to single testing, there will instead be some average number of tests, $\langle n(A, L) \rangle$, and additional vaccination
 217 based on the amount of seroconversion during the testing period, thus increasing the cost of the intervention. In
 218 return, there will be a reduction in the proportion incurring S . Relative to indefinite free testing, this reduction will
 219 be incomplete, as some proportion will pass the entire L years of the testing period without seroconversion, but still
 220 suffer multiple future infections.

221 The average number of tests is related to the proportion of people already seropositive at the start of the testing
 222 interval—who will get only one test—and the annual seroconversion probability each year through the end of the
 223 testing interval. Note that the average number of test is identical for both binary and ordinal tests. Under our
 224 assumed perfect testing model and at-most one infection per testing interval, the distinctions between outcomes in
 225 the first test only influences vaccination at that time, and further testing is only identifying initial seroconversions.

$$\begin{aligned}
 \langle n(A, L) \rangle = & 1 + {}_0\mathcal{P}_{1+}\{A\} \left(\mathcal{C}\{A\} + 2 \frac{{}_0\mathcal{P}_{1+}\{A+1\}}{{}_0\mathcal{P}_{1+}\{A\}} \mathcal{C}\{A+1\} + 3 \frac{{}_0\mathcal{P}_{1+}\{A+2\}}{{}_0\mathcal{P}_{1+}\{A\}} \mathcal{C}\{A+2\} + \dots \right. \\
 & \left. \dots + (L-1) \frac{{}_0\mathcal{P}_{1+}\{A+L-2\}}{{}_0\mathcal{P}_{1+}\{A\}} \mathcal{C}\{A+L-2\} \right) + (\mathcal{P}_0 + {}_0\mathcal{P}_{1+}\{A+L-1\})(L-1)
 \end{aligned} \tag{17}$$

226 which can be conveniently expressed as a summation of terms

$$\langle n(A, L) \rangle = 1 + {}_0\mathcal{P}_v\{A+L-1\}(L-1) + \sum_{i=0}^{L-2} \mathcal{C}\{A+i\} {}_0\mathcal{P}_{1+}\{A+i\}(i+1) \tag{18}$$

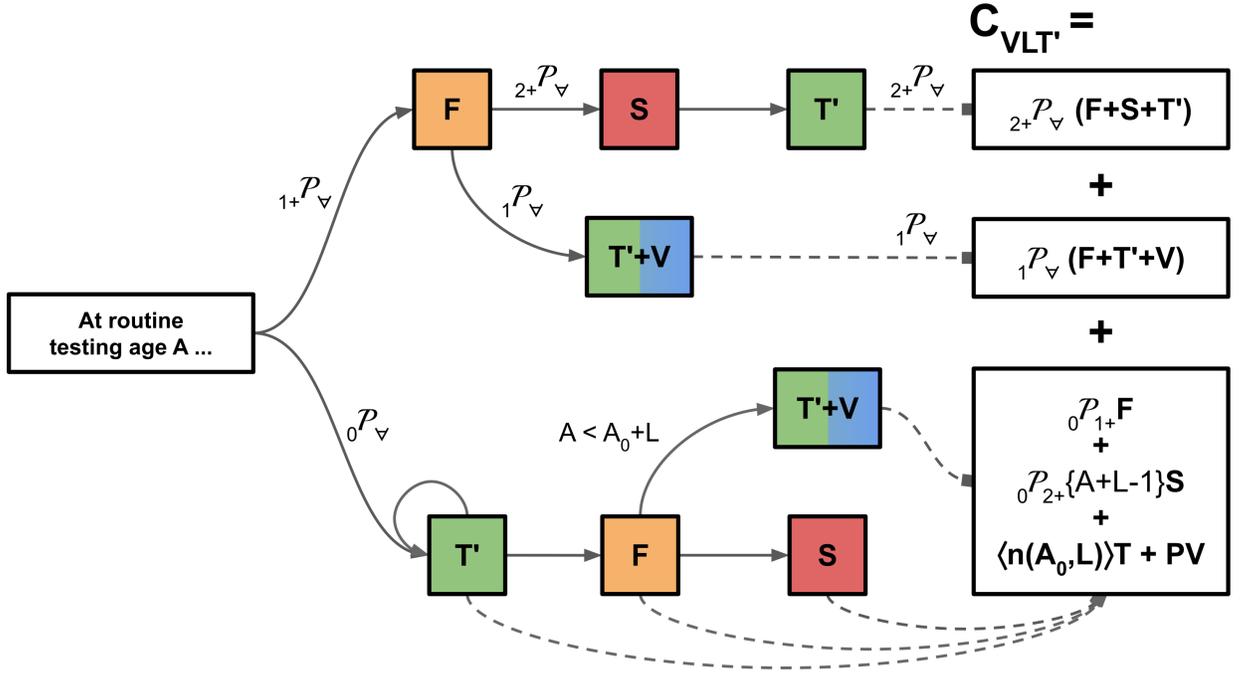


FIG. S8. **Outcomes for Vaccination with Multiple Ordinal Tests.** Pathways show the probability of various life trajectories and ultimate costs for Eq. 21, assuming the ordinal test. Unlike to single testing, both the upper (corresponding to infection(s) prior to testing) and lower (an initial non-passing test followed by any future trajectory) paths are allowed.

note that for the case $L = 1$, the summation term is empty, and therefore $\langle n(A, 1) \rangle = 1$, recovering the single test case.

Vaccination can also be considered in terms of this series, but since we are unconcerned about whether they are vaccinated due to the first test or the last, only if it happens during the interval, we can instead consider the complementary proportion of the un-vaccinated. For all versions of the test, those that are still seronegative at the end of the testing period, ${}_0\mathcal{P}_V\{A+L-1\}$, will not be vaccinated. For the ordinal test, those with two infections at the start of routine testing will also not be vaccinated; *n.b.*, these categories are mutually exclusive. Therefore, the proportion of the population vaccinated, P_V is:

$$P_V^\dagger\{A+L-1\} = 1 - {}_0\mathcal{P}_V\{A+L-1\} \quad (19)$$

$$P_V'\{A+L-1\} = P_V^\dagger\{A+L-1\} - {}_2\mathcal{P}_V\{A\} \quad (20)$$

Revisiting our net cost equations, using T^* to mean either test scenario:

$$C_{VLT^*} = \langle n(A, L) \rangle T^* + (1 - \mathcal{P}_0)F + ({}_0\mathcal{P}_{2^+}\{A+L-1\} + {}_2\mathcal{P}_V)S + P_{V^*}\{A+L-1\}V \quad (21)$$

Meaning the benefit constraint equations are:

$$\Delta_{VLT^*} = (\mathcal{P}_{2^+} - {}_0\mathcal{P}_{2^+}\{A+L-1\} + {}_2\mathcal{P}_V)S - \langle n(A, L) \rangle T^* - P_V^*\{A+L-1\}V \quad (22)$$

Again, the cost of first-like infections, F , does not appear anywhere in the constraints. We can also divide by S and again work in the non-dimensional terms to find the net benefit constraint relationship:

$$\langle n(A, L) \rangle \tau^* - P_V^*\{A+L-1\} \nu \leq (\mathcal{P}_{2^+} - {}_0\mathcal{P}_{2^+}\{A+L-1\} + {}_2\mathcal{P}_V) \quad (23)$$

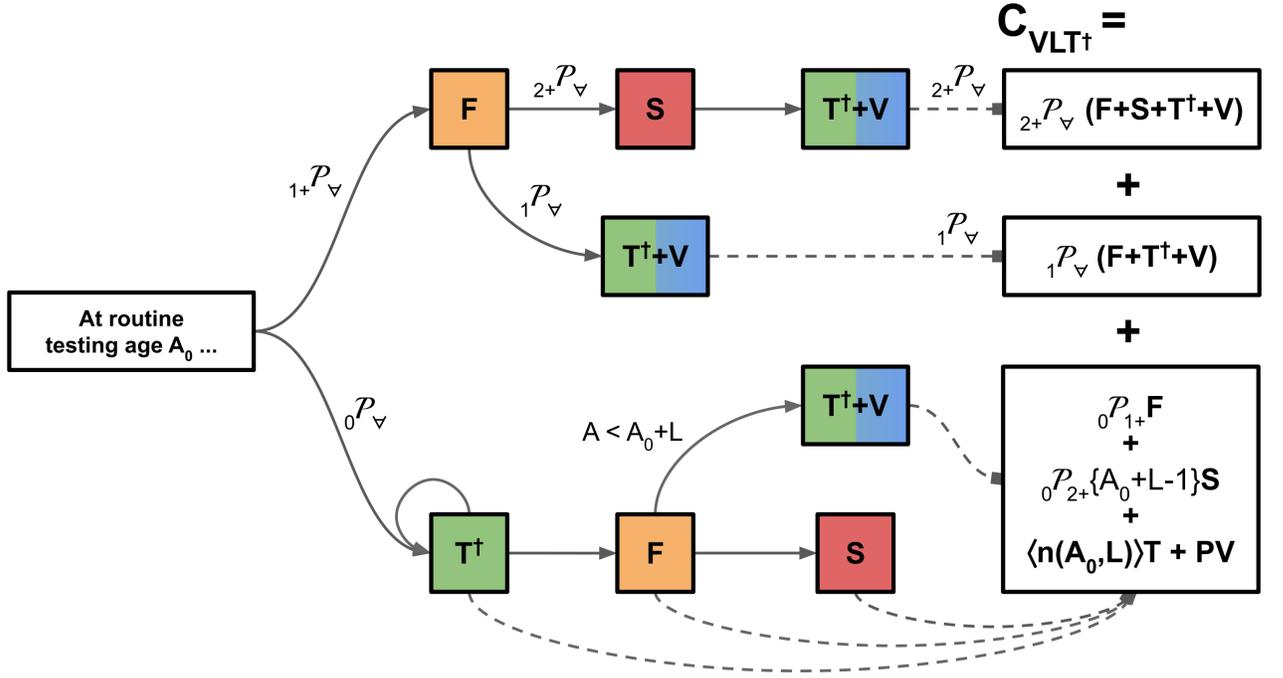


FIG. S9. **Outcomes for Vaccination with Multiple Binary Tests.** Pathways show the probability of various life trajectories and ultimate costs for Eq. 21, assuming the binary test. Again, note that the only distinction from Fig. S8 is the additional vaccination along the uppermost path, the trajectory corresponding to two or more infections prior to the routine vaccination age.

239

G. Comparison to Vaccination Without Testing

240 As a final step, we compare the benefits of vaccination with and without testing. This allows us to determine
 241 where adding testing of makes vaccination a more cost-effective intervention. Recalling C_V and C_{VLT^*} from Eq. 2
 242 and Eq. 21, respectively, we can write:

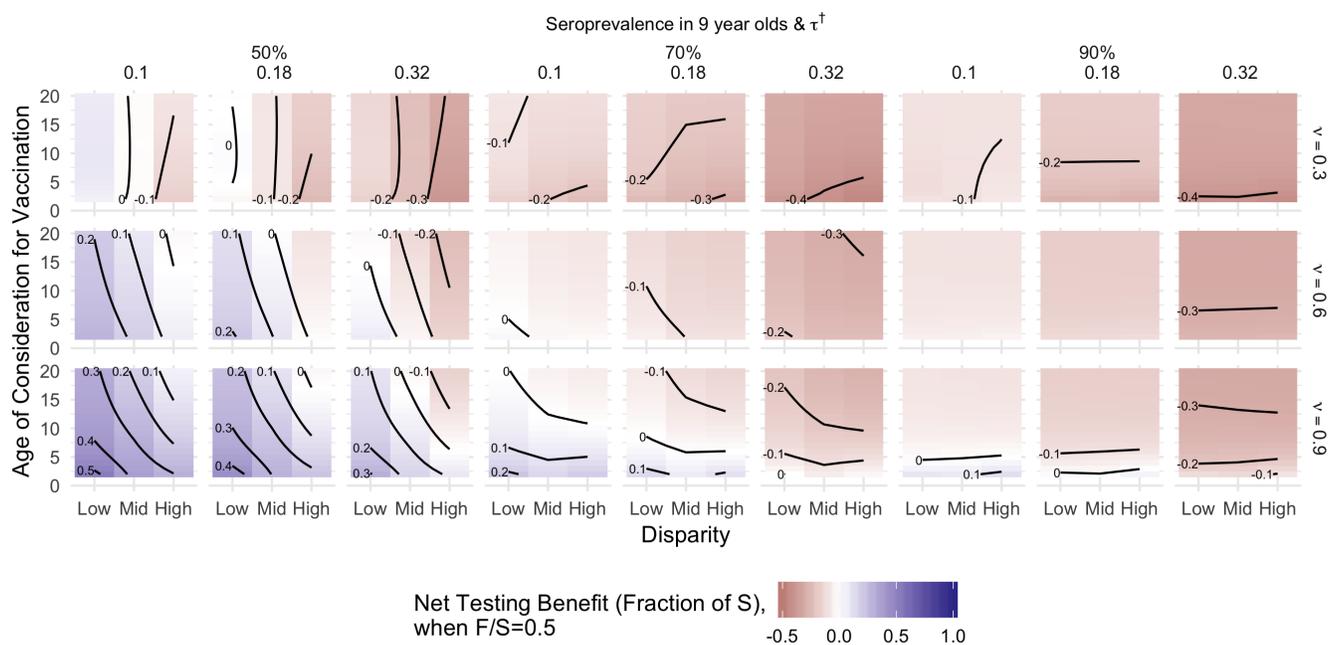
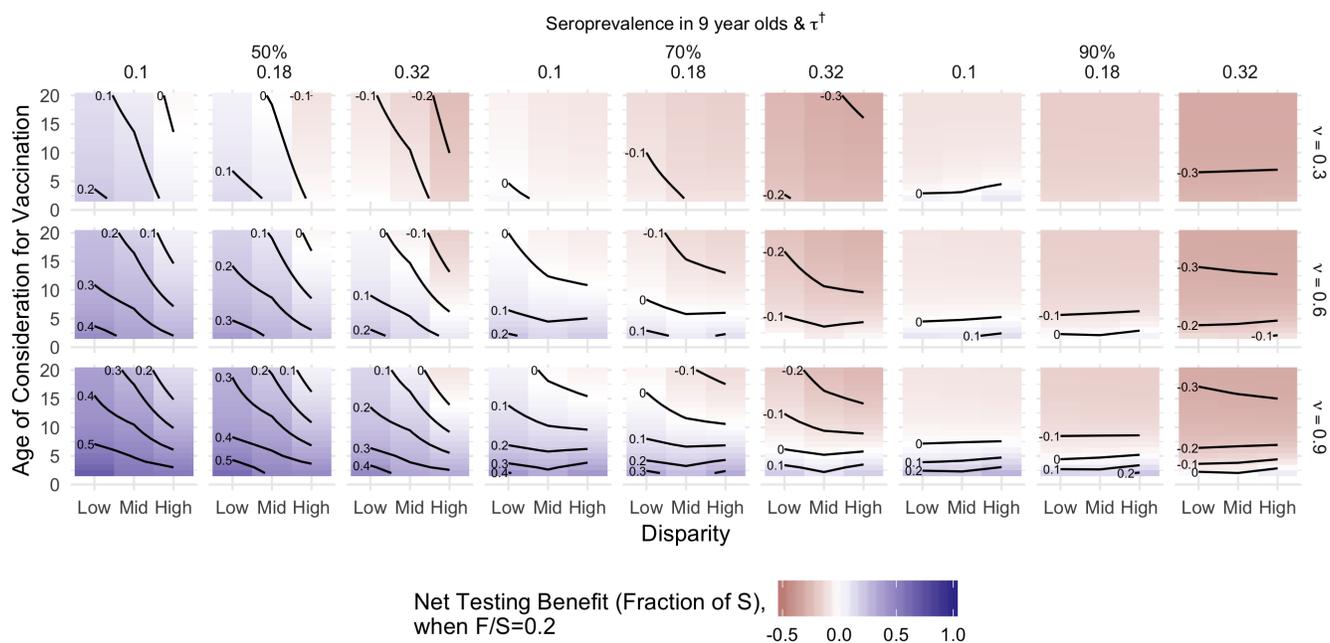
$$C_V - C_{VLT^*} = (1 + \mathcal{P}_v - 1 + \mathcal{P}_0)F + ({}_0\mathcal{P}_{1+} - {}_0\mathcal{P}_{2+}\{A+L-1\})S - \langle n(A, L) \rangle T^* + (1 - P_{V^*}\{A+L-1\})V \quad (24)$$

243 For testing to improve the intervention cost-benefit, $C_V - C_{VLT^*} \geq 0$. Imposing that constraint, and re-arranging
 244 to match the form of Eq. 23:

$$\langle n(A, L) \rangle \tau^* + (1 - P_{V^*}\{A+L-1\})\nu \leq ({}_0\mathcal{P}_{1+} - {}_0\mathcal{P}_{2+}\{A+L-1\}) - {}_0\mathcal{P}_{1+} \frac{F}{S} \quad (25)$$

245 Some notable results are derivable directly from this relationship. Increasing vaccine cost increases the advantage of
 246 the testing intervention, which is intuitive: testing decreases vaccination rates, so the more expensive the vaccine, the
 247 more cost avoided by testing. The other direct result is that increasing costs of first-like infections *decreases* the value
 248 of testing. This outcome can be understood by thinking about what the vaccine does in practice, namely preventing
 249 one of F or S , and what testing does, namely shifting prevention from F to S . As F and S become closer, the value
 250 to preventing S instead of F decreases, and thus so does value to testing. As explained in the **Context Economics**
 251 section, we consider a pessimistic $\frac{F}{S} = 0.5$.

252 In general, we see that including testing makes the intervention more beneficial in lower transmission settings, while
 253 it is wasted in high transmission settings. In high transmission settings, we already have a lot of information about
 254 serostatus, so testing provides little benefit.

FIG. S10. Vaccination Benefit with and without Testing, $F/S=0.5$.FIG. S11. Vaccination Benefit with and without Testing, $F/S=0.2$.

S3. EXPOSURE MODEL

255

256 As stated earlier, we assume there is a constant force-of-exposure. However, we also divide individuals into low or
257 high risk sub-populations, so there are three fundamental parameters in the model: the two exposure risks and the
258 proportion in high versus low risk.

259 For natural history of the infection, we assume that exposure to a serotype leads to infection if the individual
260 has not been exposed to that serotype previously. We also assume that infection in one year precludes infection in
261 the following year, consistent with empirical observations of temporary cross immunity. Finally, we assume only one
262 dengue serotype is circulating in any given year. In the next section we consider relaxing these assumptions, and
263 justify our decision to use these simpler assumptions.

264 We can use age-seroprevalence data to fit this force-of-exposure model. These data allow us to estimate the
265 probability that individuals have experienced at least one infection by a given age. If we define the constant exposure
266 forces (and complementary avoidance probabilities) $f_X = 1 - s_X$ and probability of being in the high risk group as ρ_H ,
267 the probability of being seropositive at age A is:

$$\mathbf{P}(+|A) = \rho_H(1 - s_H^A) + (1 - \rho_H)(1 - s_L^A) \quad (26)$$

268 Once we fit that model to specific age-seroprevalence data, we can then use simulation to estimate the relevant $N\mathcal{P}_X$
269 terms for that context, and thus the τ - ν constraints for that setting. While analytical derivation of these is probably
270 feasible for this set of assumptions, simulation affords us the ability to consider a broader range of assumptions in the
271 future without having to re-derive outcomes.

272 Recall that for the particulars of our cost model, we need the following probabilities:

273 The proportion of individuals that

- 274 • ${}_0\mathcal{P}_\forall$: ... are seronegative at a given age
- 275 • \mathcal{P}_{1+} : ... have one or more lifetime infections
- 276 • ${}_0\mathcal{P}_{2+}, {}_1\mathcal{P}_{2+}$: ... have no or one infection, respectively, prior to consideration for vaccination, and have two or
277 more infections over life
- 278 • ${}_{2+}\mathcal{P}_\forall$: ... have two or more infections at the time of vaccination
- 279 • ${}_1\mathcal{P}_\forall, {}_{1+}\mathcal{P}_\forall$: ... have exactly one or at least one infection at the time of vaccination; and
- 280 • $\mathcal{C}\{A\}$, the conditional probability of seroconverting between age A and $A + 1$, given seronegativity at age A .

281

A. Alternative Exposure Models

282 If there were more readily available data, it might be worth considering models incorporating (1) maternally-derived
283 temporary immunity and (2) multiple serotype circulation.

284 To include maternal immunity, we would need to consider the probability of having a previously infected mother as
285 part of the first exposure year. If an individual had an exposed mother, we would simply assume one the individuals
286 was immune in that year, meaning one fewer exposure year in their life history. We could derive the probability of
287 maternal exposure with maternal-age distribution and the age-seroprevalence from the fit, meaning we do not actually
288 need an additional parameter, just more data.

289 For multiple simultaneously circulating serotypes, we considered allowing up-to four exposures per year, probabilis-
290 tically ordered by relative weights. Since all the fitting is by first exposure, we could fit any one parameter model of
291 exposure (*e.g.* a four-trial binomial draw). This would tend to drive up second infection rates, perhaps more accu-
292 rately representing high risk populations in hyper-endemic locales where all four dengue serotypes routinely circulate.
293 This would in turn tend to drive the benefit curve towards earlier consideration for routine vaccination. However, the
294 data to correctly parametrise and constrain such a model (for example, age sero-ordinality surveys) is more detailed
295 and does not appear to be commonly available.

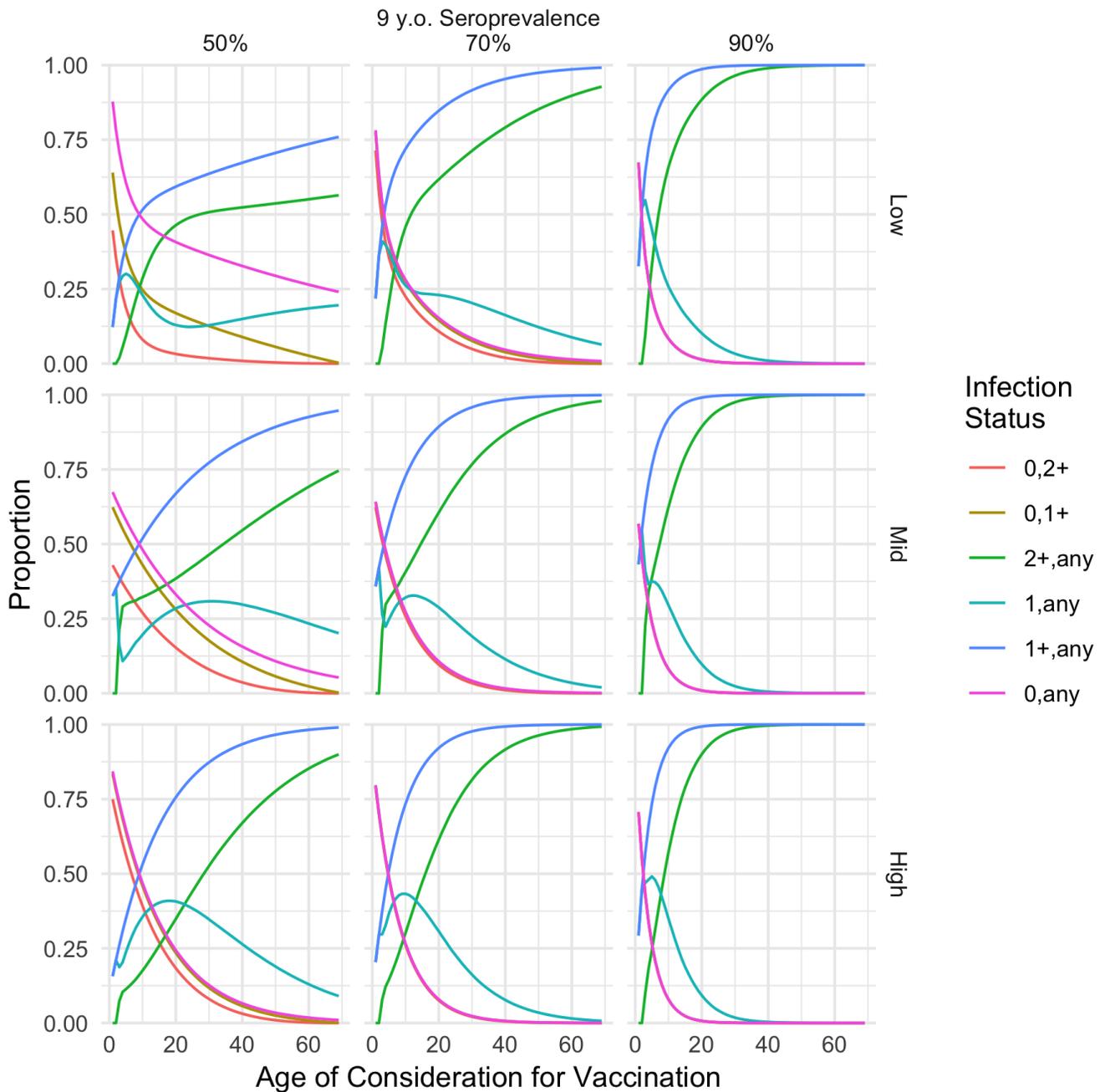


FIG. S12. **Trends by Age of Relevant Population States.** Assorted $\mathcal{N}\mathcal{P}_X$ trend lines, faceted by seroprevalence in 9-year-olds (a surrogate for overall force of exposure) and disparity level (a representation of population heterogeneity in exposure risk).

296

B. Fitting Force of Exposure Model

297 To fit the force of exposure model, we change parameters to break the parameter symmetry (*i.e.*, that either
 298 risk category could be higher or lower) and express them parameters so that they have infinite domain, which is
 299 preferable for most numerical solvers. The detailed code can be found in the `denvax` package source repository,
 300 https://gitlab.com/cabp_LSHTM/denvax, particularly the `serofit` function and utility functions earlier in the
 301 `R/main.R` script.

302

C. Serosurvey Fits

303 Using this model, we fit to two data sets, Table S1 from the CYD14 trial [2] (reference Table 4; there is a discrepancy
 304 for Indonesia in 13-16 year-olds category when compared to reference Table 1; we have assumed Table 4 is correct) and
 305 Table S2 from work in Peru [3] (sample numbers read from reference Fig. 2 and percentiles extracted from reference
 306 Fig. 3 using a webtool [4]):

Country	Age	N	S ⁺
Indonesia	2-4	87	50
	5-8	90	76
	9-12	116	101
	13-16	57	56
Malaysia	2-4	68	22
	5-8	82	28
	9-12	102	59
	13-16	48	32
Philippines	2-4	150	87
	5-8	167	125
	9-12	157	139
	13-16	128	119
Thailand	2-4	68	33
	5-8	118	73
	9-12	117	93
	13-16	38	32
Vietnam	2-4	97	43
	5-8	142	70
	9-12	145	90
	13-16	22	17

TABLE S1. Data from L’Azou (2016).

Country	Age	N	S ⁺
Peru	5	261	146
	6	307	209
	7	340	244
	8	368	269
	9	353	273
	10	340	273
	11	327	272
	12	320	257
	13	267	222
	14	238	209
	15	205	187
	16	147	135
	17	98	93

TABLE S2. Data from Morrison (2010).

307 For our practical examples, we use Peru, a Latin American setting, and Malaysia, a Southeast Asian setting. The
 308 Peru data is ideal for the maximum-likelihood style approach we used. The aggregated CYD-14 data, on the other
 309 hand, throws away information and complicates model fitting, whether using maximum likelihood or other approaches.
 310 While the data may “look better” after aggregation, this mode of reporting in fact impedes further study.

311

D. Context Sensitivities

312 For our parameter sensitivities, we considered three dimensions: *high risk fraction*, *aggregate force of exposure*, and
 313 *disparity*. In keeping with past analyses, we distinguish aggregate force of exposure by seroprevalence at a particular
 314 age. We define disparity by the seroprevalence odds ratio (OR) between the high and low risk groups. These values
 315 are governed by:

$$\begin{aligned}
\mathbf{S}^+(A) &= \rho_H \mathbf{S}_H^+(A) + (1 - \rho_H) \mathbf{S}_L^+(A) \\
&= \rho_H (1 - (1 - f_H)^A) + (1 - \rho_H) (1 - (1 - f_L)^A) \\
&= \rho_H (1 - s_H^A) + (1 - \rho_H) (1 - s_L^A)
\end{aligned} \tag{27}$$

$$\text{OR} = \frac{\mathbf{S}_H^+(A)(1 - \mathbf{S}_L^+(A))}{\mathbf{S}_L^+(A)(1 - \mathbf{S}_H^+(A))} = \frac{(1 - s_H^A)s_L^A}{(1 - s_L^A)s_H^A} \tag{28}$$

316 Therefore, for a particular $\{\rho_H, \mathbf{S}^+ = 1 - \mathbf{S}^-, \text{OR}\}$, where the dependence on A is implicit, we can identify $\alpha = s_H^A$
317 and $\beta = s_L^A$:

$$\begin{aligned}
\alpha &= 1 - \frac{\mathbf{S}^+}{\rho_H} + \frac{1 - \rho_H}{\rho_H} - \frac{1 - \rho_H}{\rho_H} \beta \\
&= \frac{1 - \mathbf{S}^+}{\rho_H} - \frac{1 - \rho_H}{\rho_H} \beta = \frac{\mathbf{S}^-}{\rho_H} - \frac{1 - \rho_H}{\rho_H} \beta \\
\text{OR} &= \frac{(1 - \alpha)\beta}{(1 - \beta)\alpha} \\
0 &= \text{OR}\alpha + (1 - \text{OR})\alpha\beta - \beta \\
&= \text{OR}(\mathbf{S}^- - (1 - \rho_H)\beta) + (1 - \text{OR})(\mathbf{S}^- - (1 - \rho_H)\beta)\beta - \rho_H\beta \\
&= \text{OR}\mathbf{S}^- + ((1 - \text{OR})(\mathbf{S}^- - \rho_H) - \text{OR})\beta - (1 - \text{OR})(1 - \rho_H)\beta^2
\end{aligned} \tag{29}$$

$$\tag{30}$$

318 where Eq. 30 can be solved with the quadratic equation, subject to $\beta \in (0, 1)$, and substituted into Eq. 29.

319 We are being somewhat circular with the concept of disparity, in that we fit the serosurvey data assuming a two-
320 group risk model, and then calculated the OR, rather than identifying mechanistically a useful distinguishing variable
321 (*e.g.* living with air conditioning versus not) and measuring OR. More targeted deployments may require finding such
322 a factor that predicts similar disparity as the serosurvey approach.

323 For our practical examples, the model parameters are:

Country	\mathbf{S}^+	ρ_H	$\log_{10}(\text{OR})$
Indonesia	0.8806040	0.4668654	1.777455
Malaysia	0.5072371	0.1401441	31.251271
Philippines	0.8419634	0.3162632	26.265428
Thailand	0.7342112	0.2579016	28.575004
Vietnam	0.5870976	0.3116296	32.564970
Peru	0.7684385	0.2940011	1.566647

TABLE S3. Fits to seroprevalence data.

324 For our sensitivity study, we used the following three by three grid for force of exposure and combined high risk
325 fraction and disparity:

Force of Exposure	Risk Fraction & Disparity		
	Low	Medium	High
Low	{0.5, 0.5, 2}	{0.5, 0.3, 17.5}	{0.5, 0.1, 33}
Medium	{0.7, 0.5, 2}	{0.7, 0.3, 17.5}	{0.7, 0.1, 33}
High	{0.9, 0.5, 2}	{0.9, 0.3, 17.5}	{0.9, 0.1, 33}

TABLE S4. Sensitivity categories, \mathbf{S}^+ , ρ_H , \log_{10} OR.

326

S4. CONTEXT ECONOMICS

327 For the two practical examples in the main text, Peru and Malaysia, we use previously assumed cost data for S ,
328 V , and T to set ν and τ [5].

329 The generally assumed cost of a vaccine regimen is 78 USD (21 USD per dose, for a three dose course, plus 15 USD
330 for additional logistical costs), and test costs have been suggested at approximately 5 USD.

³³¹ For Malaysia, if we assume the public payer perspective, the benefit of preventing a second-like infection is roughly
³³² $S = 86$ USD, therefore $\nu \approx .9$ and $\tau \approx .2$. For Peru, if we assume the public payer perspective, $S = 223$, $\nu \approx .3$ and
³³³ $\tau \approx .1$.

³³⁴ In these two settings, we also estimated $\frac{F}{S}$; in Malaysia, it is $\frac{F}{S} \approx 0.2$ while in Peru $\frac{F}{S} \approx 0.4$. For our comparison to
³³⁵ vaccination without testing, we assumed $\frac{F}{S} = 0.5$ as an upper limit.

-
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