1	Human health as a motivator for climate change mitigation:	
2	results from four European high-income countries	
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36 Abstract

37 Invoking health benefits to promote climate-friendly household behavior has three unique 38 advantages: (i) health co-benefits accrue directly to the acting individual, they are "private 39 goods" rather than public ones; (ii) the evidence base and magnitude of health co-benefits is 40 well-established; and (iii) the idea of a healthy life-style is well-engrained in public discourse, 41 much more so than that of climate-friendly life-style. In previous research assessing the 42 influence of information on health effects on people's motivation to adopt mitigation actions, 43 health co-benefits for the individual were typically confounded with collective health co-44 benefits, for example from pollution reduction. The present research aims to overcome this 45 limitation by providing information on individual health co-benefits that are unconditional on the 46 actions of others (direct health co-benefits). We report effects of this kind of health information 47 on stated preferences to adopt mitigation actions as well as on simulation-based carbon emission 48 reductions in an experimental setting among 308 households in 4 mid-size case-study cities in 4 49 European high-income countries: France, Germany, Norway and Sweden. For each mitigation 50 action from the sectors food, housing, and mobility, half of the sample received the amount of 51 CO_2 equivalents (CO_2 -eq) saved and the financial costs or savings the respective action 52 generated. The other half additionally received information on direct health co-benefits, where 53 applicable. For households receiving information on direct health co-benefits, we find a higher 54 mean willingness to adopt food and housing actions, and a greater proportion very willing to 55 adopt one or more mitigation actions (OR 1.86, 95% CI 1.1, 3.12); and a greater simulated 56 reduction in overall carbon footprint: difference in percent reduction -2.70%, (95% CI -5.34, -57 0.04) overall and -4.45%, (95% CI -8.26, -0.64) for food. Our study is the first to show that 58 providing information on strictly unconditional, individual health co-benefits can motivate 59 households in high-income countries to adopt mitigation actions. 60

61 Keywords: climate change, health co-benefits, mitigation, household preferences, health,
62 behavior

63

64 Introduction

65 Climate change has far-reaching effects on human health (Myers and Patz 2009, IPCC, 66 2014, Woodward et al. 2014). The human health effects of climate change have been recognized 67 as one of the greatest threats of mitigation failure (Costello et al., 2009; Patz, Campbell-68 Lendrum, Holloway, & Foley, 2005; Stern, 2007). However, with transformative mitigation 69 policies, there are also numerous opportunities for health gains, the so-called *health co-benefits*. 70 It is therefore no contradiction, when the latest Lancet report on climate change and health sees 71 climate change as potentially "the greatest global health opportunity of the 21st century" (Watts 72 et al., 2015). Health co-benefits arise when a policy or a behavior, which primarily addresses 73 mitigation, also generates health benefits. One example is using active transport such as cycling 74 or walking instead of fossil fuel powered cars, which not only saves emissions but also increases 75 cardiovascular fitness. Scholars have increasingly advocated the use of health co-benefits of 76 climate-friendly behavior as a motivator for adopting mitigation actions (Nisbet & Gick, 2008, 77 Myers et al. 2012, Sauerborn et al., 2009). Health nevertheless constitutes the most understudied 78 argument in existing climate communication studies, with the few existing studies yielding 79 partially promising, but inconsistent results (Bain et al., 2016; Maibach et al., 2010; Myers et al., 80 2012).

81 Unconditional, individual health co-benefits have three advantages for communication 82 and motivation:

(i) health co-benefits can accrue directly to the acting individual in addition to being
contingent on other individuals to join in climate-friendly behavior. Conversely, in order to
receive health co-benefits, individuals cannot "free ride". Health co-benefits are "private goods"
as well as public ones.

(ii) the evidence base for and magnitude of health co-benefits is well-established, and
can be gleaned from numerous epidemiological studies (e.g., Haines et al., 2007; Milner et al.,
2015; Wilkinson et al., 2009; Woodcock et al., 2009)

90 (iii) the idea of a healthy life-style is well-engrained in public discourse, much more so
91 than that of climate friendly life-style (Magnusson et al., 2003; DEFRA, 2002).

It may be argued that healthy life-styles, although widely talked about, are not necessarily implemented, and that therefore the health argument does not constitute a helpful addition to the climate change discourse. However, research on consumers` willingness to buy organic food, for example, showed that health impacts constitute a more important argument than environmental impacts – even though the health effects of organically grown food are ambiguous (Honkanen et al, 2006), and the health benefits for the individual consumer are far less supported than those for

98 the environment (Magnusson et al., 2003). Moreover, it was argued before that conceptualizing 99 environmental actions as health actions may be useful (cf. Staub & Leahy, 2014). Therefore, 100 effects on individual health may be a particularly compelling argument for individuals to 101 implement mitigation actions (see Staub & Leahy, 2014 for a similar argument with respect to 102 pro-environmental behavior more generally). From a policy angle, it is therefore surprising to 103 note that the individual health benefits accruing to climate friendly vior are not (yet) a focus of 104 communication to the public, and are not yet made salient enough in negotiations (Ganten, 105 Haines, & Souhami, 2010; Sauerborn, Kjellstrom, & Nilsson, 2009). Consequently, health co-106 benefits are still largely underestimated by the public (Maibach et al. 2010). Here we focus on 107 households as the target group for information on health co-benefits. Private households have 108 been identified as key actors in global climate change mitigation (Aall & Hille 2010, Dubois & 109 Ceron 2015, Sovacool 2014), and households in high-income countries influence up to 72% of 110 global greenhouse gas emissions (Hertwich and Peters 2009).

111 Previous studies have tested the effectiveness of framing actions to reduce emissions-112 mitigation actions—around their benefits on human health, and yielded inconsistent results. In a 113 national representative US sample, a health benefit frame (but, in contrast, not an environmental 114 risk or national security frame), elicited positive emotional responses such as hopefulness, which 115 are believed to be more in line with mitigation policy support than negative emotional responses 116 such as anger (Myers et al., 2012). In two large US samples, however, a health frame 117 emphasizing the benefits of combatting climate change on human health was ineffective in 118 fostering support for climate change mitigation (Bernauer & McGrath, 2016).

119 A shortcoming of previous studies that estimated the effectiveness of health co-benefits 120 on people's support for mitigation actions is their confounding with the collective action 121 *problem.* One example is cleaner air, as it is certainly a laudable public health goal—but only 122 achievable by collective action. An individual who chooses to reduce flying is helping to achieve 123 the goal of lower global emissions, but (may) only benefit personally from this action indirectly, 124 i.e. if others contribute as well. Bernauer and McGrath (2016; Supplementary material, p.6) use 125 the framing "using cleaner forms of energy—such as solar and wind power—will reduce air and 126 water pollution, thereby preventing many forms of illness.". Such a frame can only be expected 127 to induce behavioral change if the addressee believes that a sufficiently high portion of society 128 also contributes to the collective good of, for example, clean air. The mitigating actions of others 129 are fundamentally uncertain, however, so that people may choose to not contribute to the public 130 good of cleaner air, but to free-ride instead. A counter-example is eating less meat, which also 131 helps to achieve lower emissions and hereby contributes to a public good (Aston, Smith, &

Powles, 2012), but additionally reduces risks for cardiovascular disease, type 2 diabetes and
colorectal cancer, a direct and unconditional health co-benefit for the individual (Milner et al.,
2015).

Thus, we do not contend that the concept of individual and unconditional health cobenefits will solve the collective action problem - for health effects to achieve a demonstrable climate effect, a sufficiently high portion of society still needs to contribute. However, the fact that some health co-benefits are unconditional on the actions of others render them a particularly valuable tool as they help to preempt the debilitating effects of free-riding.

By confounding direct and unconditional health effects for the individual with common
public goods, the effectiveness of framing mitigation in terms of its co-benefits on health cannot
be clearly assessed. When studies fail to find a health framing effect (Bernauer & McGrath,
2016), it remains unclear whether people are unresponsive to health arguments, or unwilling to

144 invest in public goods. Conversely, when studies find a health framing effect (e.g., Myers et al.,

145 2012), it remains unclear whether it was the health argument that convinced them, or whether

they wanted—or felt obliged—to contribute to a common public good.

147 The present research therefore aims to overcome this limitation. As previous research could not

148 delineate whether it was health information that did or did not increase people's willingness to

adopt mitigation actions, or whether people did or did not want to invest into a common public

150 good, such as an improved climate, we deliver an estimate of the influence of *direct health co-*

benefits: health co-benefits that are unconditional on the actions of others to yield an effect onthe individual.

In four European high-income countries, we informed one half of participants about direct health co-benefits of adopting a range of actions that are beneficial for the individual adopting the action, irrespective of the action of others; the other half did not receive information on direct

health co-benefits. We assessed households' stated willingness to adopt the actions, and

157 calculated the resulting carbon footprint reduction.

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159 Methods

160 The study was based on the interdisciplinary research project *HOusehold Preferences for* 161 *reducing greenhouse gas Emission in four European high income countries* (HOPE), designed to 162 gather data relevant to household choices for greenhouse gas mitigation from selected case-study 163 cities in France, Germany, Norway and Sweden.

164

Sample size. Prior to data collection, we determined the sample size needed to detect an experimental effect of information on direct health co-benefits on the stated preferences to adopt mitigation actions as well as on simulation-based carbon emission reductions in a two-group between-subject design (health information given versus no health information given), assuming one-sided testing and a small- to medium-sized effect of d = .3 (based on common effect sizes found in framing studies, e.g. Bain et al. 2012), alpha = .05, and .8 power, and assuming no cluster effect. Based on these considerations, we aimed for a total sample-size of N = 278.

172 Recruitment and sampling. Households were recruited by sending invitation letters to a 173 random sample of inhabitants on the population registers of the target cities, supplemented by 174 others recruited through responses to media announcements. The sample was stratified by 175 country. Allocation to the group receiving information on direct health co-benefits health was 176 done randomly within each country among households agreeing to participate in the study. In 177 total, 309 households were recruited: 70 from the Communauté du Pays d'Aix (France), 107 178 from Mannheim (Germany), 58 from Bergen (Norway), and 74 from Umeå (Sweden). Roughly 179 half of households received information on health co-benefits (n=156), the other half did not 180 (n=152,). For one household, coding of which group it belonged to was lost; it was therefore 181 excluded from the analysis, rendering a final sample size of N=308. Interviews with households 182 were carried out between June and November 2016.

183 Ethics approval and consent to participate. All participants were given written 184 information about the study objectives and modalities (points of assessment, length of 185 questionnaires), data preparation and pseudonymized data storage, the expected amount of 186 commitment, the voluntary nature of participation, and their right to withdraw at any time. 187 Furthermore, participants were informed verbally about the study purpose and procedures and 188 were given the chance to ask questions. All participants provided written informed consent. All 189 countries assure that data processing and storage is done in line with European and national data 190 protection rules. Where necessary the study procedures were approved by an ethical committee. 191 In Norway the Norwegian Center for Research Data approved of the study (44003). In Germany 192 the Institutional Review Board of the Medical Faculty by the University of Heidelberg approved 193 of the study (S-611/2015). In Sweden the study was approved by the Regional Ethical Review 194 Board in Umeå (2015/357-31Ö). In France the project needed to fullfill the obligations of the 195 CNIL (Commission nationale informatique et libertés), no specific ethical approval was 196 necessary.

197 Data collection procedure: Stated preferences and simulated carbon reduction. In
198 the present study, we report on two different dependent variables: (i) the intention-based actions

- 199 where participants rated their preference to adopt each action using a 5-point Likert scale
- 200 (1=very willing to 5= not at all willing) in round 1; and (ii) the simulation-based actions of
- 201 achieved carbon reductions when participants were asked to select actions they would like to
- 202 implement to reduce their carbon footprint in a simulation game in round 2. The carbon emission
- 203 reduction was calculated by the FCS Tool based on selected mitigation actions.
- Households' initial footprint was calculated using a comprehensive on-line carbon footprint calculator (the Footprint Calculation and Simulation-Tool (FCS-Tool), Dubois *et al*, manuscript in preparation), which computed all greenhouse gas emissions in CO₂-equivalents (CO₂e) made by the household in one year (Wiedmann & Minx, 2008) under the headings:
- 208 *housing, mobility, food, and other consumption.* This provided the baseline estimate of CO₂e
- 209 emissions for each participating household.
- Additional socioeconomic variables of the households, such as age, income, education etc. wereassessed in this first step.
- 212 During a subsequent on-site visit approx. two to four weeks after assessment of a 213 household's initial carbon footprint, each participant was presented with their initial footprint 214 and asked to consider each of 65 household mitigation actions within their range (i.e. excluding 215 those that were not relevant for the households, or had already been implemented). In the first 216 round (rating of mitigation actions), participants rated their willingness to adopt these 217 mitigation actionss. In the second round (choosing and ranking in voluntary scenario), 218 participants chose the actions, they would actually like to implement, and received real-time 219 feedback on the simulated reduction achieved. If participants did not reach the aim to reduce 220 their carbon footprint by 50% in this round, they entered into a third round (choosing and 221 ranking in voluntary scenario), in which they chose additional actions they would be most likely 222 to implement if they were forced to reduce their carbon footprint by 50% until they reached the 223 requested reduction. Additional information on the data collection procedures is given in the 224 Study protocol of the HOPE study (Herrmann et al. 2017).
- 225

Experimental variation of information. Participants were given information specific to their household on the CO₂ reduction and the monetary costs or savings of adopting each possible mitigation action using action cards. For example, participants were asked to rate their preference to replace 60% of meat-based foods with vegetarian options. While all participants received information on the associated costs and CO₂e reduction, only participants in the experimental group additionally received information on direct health co-benefits. For example,

- 232 participants were given the additional information that an increase in vegetarian food intake
- would yield an appreciable *positive* health benefit for them (Fig. 1).

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Fig 1. Examples of action cards showing one mitigation action from the sector *housing* and one from the sector food.

- 238 Cards on the left panel were given to the experimental group (with health information), cards on
- the right panel to the control group (without health information). Both groups received
- 240 information on costs and CO₂ reduction.
- 241

Assessment of direct health co-benefits. The assessment of the likely health impact of
individual mitigation actions was based on semi-quantitative extrapolation of evidence (an
'expert judgement') from published modelling studies of low carbon interventions in electricity
production (Markandya et al., 2006), housing (Wilkinson et al., 2009), transport (Woodcook et
al., 2009) and food and agriculture (Milner et al., 2015) for similar European populations. Using
this evidence, we classified the likely impact on life expectancy using four categories: small
(estimated increase in life expectancy < 1 month); moderate (estimated increase in life

expectancy 1-3 months); substantial (>3 months increase), and negative (a decrease in life

250 expectancy). Eleven of the 65 mitigation actions were judged likely to have a direct positive 251 health impact for the individual (three housing actions, and four each for mobility and food), one 252 (reducing the indoor temperature thermostat setting by 3 degrees Celsius) to have a negative 253 health impact, and 53 to have no clearly positive or negative impact on the individual. No 254 mitigation action in the category of 'other consumption' was deemed likely to have appreciable 255 health impact (see Study Protocol for the detailed list of mitigation actions with health effects, 256 Hermann et al., 2017). 257 258 **Hypotheses**. Three hypotheses were *pre-registered* in the Study Protocol prior to data analysis 259 (Hermann et al., 2017). Specifically, we expected that participants who received additional 260 information on the direct health impact of adopting mitigation actions ('direct health co-261 benefits') would 262 263 1. report a higher stated willingness to implement mitigation actions with health co-264 benefits compared to households not receiving health information; 265 2. select higher numbers of mitigation actions with direct health co-benefits compared to 266 households not receiving health information (i.e., comparing households); and 267 3. select more actions with direct health co-benefits than actions without direct health 268 co-benefits (i.e., comparing actions). 269 270 Additionally, we investigate *exploratively*, 271 1. if the potential impact of information on direct health co-benefits varies by sector 272 (housing, food, mobility); and

- 2732. whether participants who received additional information on direct health co-benefits274achieve a higher household carbon footprint reduction in the simulation
- 275

276 Analysis

- 277 We present two sets of dependent variables:
- Stated preference to implement: Analysis of the *proportion* of participants rating each
 mitigation action with health impact as 'very willing' to implement it; and analysis of the
 mean willingness to adopt actions from one sector.
- (2) Simulated carbon footprint reduction: Analysis of the difference in carbon footprint
 achieved by the actions participants said they would implement in a voluntary scenario if
 asked to aim for a substantial reduction in their household's carbon emissions by 2030.

284 For the first analyses (1), we report the mean preference to adopt actions from each sector. 285 For a stricter measurement of preference, we additionally tabulate (by health information status) 286 the number and proportion of people very willing to implement each individual action if it is 287 relevant to them (i.e. excluding those for whom the action is not applicable or already 288 implemented), and the proportion very willing to adopt *any* one or more of the actions, together 289 with unadjusted and adjusted odds ratios as obtained from logistic regression analysis. The co-290 variates used in the adjusted model included: the initial carbon footprint, household type (family 291 with children (yes/no), no. of adults in working age, (no. of adults>=65 years), housing tenure 292 (ves/no), household income, whether a vegetarian household (ves/no), age of the house, city 293 district (self-report: city centre, suburb, or rural by tendency) and country. All quantitative 294 variables were z-standardized prior to inclusion into the regression model. For all preregistered 295 hypotheses we present one-sided testing results (where applicable), and present results both 296 without controlling for the list of covariates (as the specific list of co-variates was not pre-297 registered), and when including the full list of co-variates to investigate potential differences.

298 For the calculated changes in simulated household carbon footprints (2), we present 299 evidence (by health information status) of the baseline CO₂e footprint and the change when 300 asked to aim for substantial reduction by 2030, together with the percentage change, and the 301 difference in percentage reduction between those with and without evidence of health impact. 302 Regression-based estimates of the difference in the reduction in CO₂e were adjusted for the same 303 covariates as for analysis (1). For all exploratory hypotheses we present two-sided testing results, 304 together with the full list of covariates to investigate in-depth which households characteristics 305 shape willingness to adopt mitigation actions when presented with their health co-benefits, and 306 to control for baseline differences in the experimental and control group.

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309 Results

The characteristics of study households by health information status are summarized in Table 1. Those receiving health information were broadly similar to those not receiving it, but they were fewer families with children, resulting in a (marginally) lower initial carbon footprint per household. Importantly, however, the initial carbon footprint per capita did not differ between both groups, suggesting overall comparable lifestyles. **Table 1.** Household characteristics in groups with and without

information on the health impact of each mitigation action.

	Mean (standard deviation) or number <i>(percent)</i>		
	Group without information on health impacts (n=152)	Group with information on health impacts (n=156)	Comparison
Initial carbon footprint per <u>household</u> (tons CO₂e/year)	18,200 (10,640)	16,090 <i>(8,050)</i>	t(281.33)=1.96, <i>p</i> =.051
Initial Carbon footprint per <u>capita</u> (tons CO₂e/year)	10,549 <i>(5,325)</i>	10,120 (4,014)	t(306)=0.8, p=.43
Housing tenure Owner Tenant	96 (63.2%) 56 (36.8%)	95 <i>(60.9%)</i> 61 <i>(39.1%)</i>	<i>X</i> (1)=.17, <i>p</i> =.73
Household income (€/month)	3791 (2124)	3483 (1616)	t(306)=1.5, <i>p</i> =.15
Whether vegetarian household No Yes	135 (88.8%) 17 (11.2%)	146 (93.6%) 10 (6.4%)	X(1)=2.19, p=.16
Age of house Pre-1950 1950-1989 1990 or later	44 (29.0%) 72 (47.4%) 36 (23.7%)	33 (21.2%) 79 (50.6%) 44 (28.2%)	X(2)=2.64, p=.27
Location Rural/city outskirts Suburb Town/city centre	44 (28.9%) 62 (40.8%) 46 (30.3%)	36 (23.1%) 66 (42.3%) 54 (34.6%)	X(2)=1.5, p=.50
Country France Germany Norway Sweden	36 (23.7%) 52 (34.2%) 35 (23.0%) 29 (19.1%)	34 (21.8%) 55 (35.3%) 38 (24.4%) 29 (18.6%)	X(3)=.21, p=.98

316

317 **Pre-registered hypothesis 1: Preference to adopt mitigation actions across sectors**. We

318 compared the mean stated willingness to implement all feasible mitigation measures h were

319 direct health co-benefits exist. In line with our expectation, households who were given

information on direct health co-benefits reported a higher willingness to implement these
measures (M=2.70, SD=0.66) than households not provided with information on direct health
co-benefits (M=2.60, SD=0.70), t(300)=1.8, p=.03. Results do not change when including the
full list of co-variates, F(1)=3.4, p=.03.

324

325 Explorative analysis: Preference to adopt mitigation actions: Mean willingness per sector.

326 We conducted regression analyses assessing the impact of the health information on the mean 327 rated willingness to adopt mitigation actions for each of the three sectors for which health co-328 benefits exist: Food, housing and mobility. For this analysis, we controlled for the list of 329 covariates (Table 1) to assess the health impact over and above small baseline differences 330 between experimental and control group. Figure 2 shows that participants receiving information 331 on health co-benefits were more willing to implement actions in the areas food and housing. 332 Receiving information on health co-benefits did not change participants' willingness to 333 implement any changes in the sector mobility. These results suggest that the differences in 334 overall willingness to implement described above were driven entirely by the food and housing 335 sector.

336



- 338 Fig 2. Results of three regression models estimating the effect of informing participants
- 339 about health impacts on preference to implement mitigation actions in the sectors housing,
- 340 food, and mobility, respectively. The health effect on Housing (Panel A), Food, (Panel B),
- 341 Mobility (Panel C), each controlling for all co-variates entered at once. Results of 95% CI that do
- not include the null in bold.
- 343

344 Explorative analysis: Preference to adopt mitigation actions: Proportion of respondents

- 345 "very willing". To obtain a more differentiated picture of the impact of the health information,
- 346 we assessed its effect on each of the 12 mitigation actions with a known health effect separately,
- 347 and whether providing health information yielded a higher proportion of participants saying they
- 348 were "very willing" to implement each of these actions. We chose to focus on highest ratings of
- 349 stated willingness only, as participants with the highest intention to mitigate would realistically
- be the ones most likely to implement their intentions in the real world. Table 2 provides an
- 351 overview of the analyses separately for an unadjusted version, and a version adjusted for all
- 352 covariates. The overall proportion of households 'very willing' to adopt any one or more of the
- 353 mitigation actions with health impact was appreciably greater in those given health information
- 354 (adjusted odds ratio 1.86, 95% CI 1.10, 3.12).

Table 2. Numbers (%) and odds ratios (95% CI) for being very willing to implement each of the specified mitigation actions by whether information on health impact.

Mitigation action	Health impact	Number/denominator and (%) very willing to implement action		
		Without health information	With health information	
Housing				
Insulation of roof/attic	+	13/50 <i>(26.0%)</i>	19/49 <i>(38.8%)</i>	1
Insulation of walls	+	9/65 <i>(13.9%)</i>	13/62 (21.0%)	1
Improve windows (increase glazing of your windows)	+	16/61 (26.2%)	23/62 (37.1%)	1
Lower thermostat setting by 3°C	-	13/141 (9.2%)	13/139 (9.4%)	1
Any housing action	- to +	36/152† (23.7%)	49/156† <i>(31.4%)</i>	1
Dietary change				
Gradually give up on ready-made meals (e.g. frozen pizza, canned soups)	++	33/87 <i>(37.9%)</i>	32/74 (43.2%)	1
Eat 30% more vegetarian food (less meat and fish)	+++	35/138 (25.4%)	50/144 (34.7%)	1
Eat 60% more vegetarian food (less meat and fish)	+++	16/138 (11.6%)	18/143 <i>(12.6%)</i>	1
Stop eating meat (and fish?)	+++	2/138 (1.5%)	4/144 (2.1%)	1
Any dietary change	++ to +++	57/152† <i>(37.5%)</i>	68/156† <i>(43.6%)</i>	1.
<u>Travel/mobility</u>				
Shift more than 30% of car journeys to public transport	++	26/118 (22.0%)	27/136 (19.9%)	0
Shift to non-motorized transport (walk, bike) instead of public transport	+++	32/103 (31.1%)	41/113 (36.3%)	1
Decrease travel by cars public transport and other motorized vehicles by 30%.	++	17/146 (11.6%)	20/150 (13.3%)	1
Give up your car(s) and other motorized vehicle(s)	++	8/121 (6.6%)	5/134 <i>(3.7%)</i>	0
Any mobility action	++ to +++	58/152† (38.2%)	63/156† <i>(40.4%)</i>	1
Any of the above (i.e. any mitigation action with health impact)	- to +++	100/152†(65.8%)	121/156 ⁺ (77.6%)	1

* -- Initial carbon footprint, household type, tenure, household income, whether vegetarian household, age of house, urban-rural location, countries of the second second

[†] -- Denominator is all households, including those for whom individual actions are already implemented or not relevant

355 356 357 Table 3 examines the differences in the effect of being given health information in relation to 358 household type. There is no clear evidence of statistical interaction (effect modification) here 359 (p=0.12), but the point estimates vary – with households of working age adults without children 360 apparently showing much greater likelihood of being very willing to adopt any of the mitigation 361 actions with health impact (OR 3.12, 95% CI 1.45, 6.68), while the point estimate of the odds 362 ratios for households composed only of members over the age of 65 years was well below unity 363 (0.50, 95% CI 0.10, 2.44).

Table 3. Odds ratios (95% CI) for being very willing to implement any household mitigation action with health impact: ORs for those given health information *vs* those not given health information by household type

	Odds ratio (95% CI) for wi mitigation r	Test of statistical		
Household type	Unadjusted	Adjusted for all covariates*	interaction, adjusted model	
Families with children	1.24 (0.50, 3.05)	1.16 (0.44, 3.05)		
Working age adults, no children	3.05 (1.48, 6.28)	3.12 (1.45, 6.68)	p=0.12	
Adults >=65 years	0.89 (0.26, 3.04)	0.50 (0.10, 2.44)		

* -- Initial carbon footprint, household type, tenure, household income, whether vegetarian household, age of house, urban-rural location, country

365

366 Pre-registered hypothesis 2: Selection of mitigation actions with direct health co-benefits, 367 in 'voluntary' and 'forced' scenario. We investigated how many mitigation actions that exert a 368 direct health co-benefit were chosen in the 'voluntary' and the 'forced' scenario. In line with our 369 hypothesis, households given information on direct health co-benefits chose, in the voluntary 370 scenario, more actions that exert direct health co-benefits (M=2.00, SD=1.45) compared to 371 households not given this information (M=1.70, SD=1.23), both without the list of covariates, 372 t(306)=1.8, p=.04, and when controlling for the list of covariates, F(1)=5.61, p=.02. In the forced 373 scenario, there were no differences in the number of actions that exert direct health-co-benefits 374 chosen by households given health information (M=3.2, SD=1.6) compared to households not 375 given health information (M=3.20, SD=1.6), either without the list of covariates, t(306)=0.9, 376 p=.38, or when controlling for the list of covariates, F(1)=1.93, p=.17.

377

378 Pre-registered hypothesis 3: Comparison of actions that do and do not exert direct health

379 co-benefits. Eleven mitigation actions were judged to exert positive direct health co-benefits (see

- Table 2); 53 were judged not to exert direct health co-benefits. To assess whether receiving
- information on direct health co-benefits affected the popularity of these 11 actions with a known
- health effect, we first compared all households' overall willingness to implement these 11
- actions with households` overall willingness to implement the 53 actions without known health
- 384 effect. Generally, households were more willing to implement those actions that do not exert
- 385 direct health co-benefits (M=3.04, SD=.59) than those actions that do exert direct health co-
- benefits (M=2.63, SD=.69), t(302)=11.4, p<.001. In a next step, we assessed whether this
- 387 difference was smaller for the group receiving health information. Contrary to our expectation,

this difference was not significantly reduced for households receiving information on health cobenefits (M=.44, SD=.64) compared to not receiving this information (M=.38, SD=.60),

390 t(300)=.84, p=.40.

391

392 Exploratory analysis: Impact on carbon footprint. Table 4 reports the results of the 393 simulation exercise in which householders were asked to select the mitigation actions they would 394 implement if required to aim for a substantial reduction in emissions by 2030. As noted above, 395 households given health information had somewhat lower baseline carbon footprints. This was 396 true overall and for the groups of actions relating to housing interventions, food/dietary change, 397 and mobility/transport individually. However, the percentage reductions in carbon footprint 398 achieved by the selections made by participants were greater among those given health 399 information. There was evidence of a greater percentage reduction among those given health 400 information with respect to food/dietary emissions (difference in percentage reduction -4.45%, 401 95% CI -8.26, -0.64, fully adjusted analysis) and for all mitigation actions (difference in 402 percentage reduction -2.70%, 95% CI -5.34, -0.04, fully adjusted). These results indicate that 403 providing information on direct health co-benefits does not just alter the preference to choose 404 health-relevant mitigation actions but also the overall total of emissions (summed across all 405 actions, not just those with health effects).

	Tonnes CO ₂ eq emissions/household:			:	Percent change (95% CI): simulation/baseline*100%	
	Baseline		Simulation: asked to aim for 50% reduction by 2030			
	Without health information	With health information	Without health information	With health information	Without health information	With health informatior
Housing	3,641	3,204	2,813	2,295	-22.0% (-18.4, -25.7)	-24.2% (-21.2, -27.2
Food	4,855	4,740	3,365	3,086	-31.1% (-28.4, -33.7)	-34.9% (-32.1 <i>,</i> -37.6
Mobility/transport	7,007	5,762	5,481	4,527	-23.5 (-20.4, -26.7)	-24.0% (-20.7, -27.3
ALL ⁺	18,200	16,092	14,061	12,148	-23.0% (-21.2, -24.8)	-24.9% (-22.9, -26.8

Table 4. Change in calculated carbon footprint in tonnes CO₂-eq/household/year with simulated 50% target reduction: results with

* -- Initial carbon footprint, household type, tenure, household income, whether vegetarian household, age of house, urban-rura
 + -- ALL also includes emissions relating to consumer goods which are not separately shown in the table as choices in consumer go n health.

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408 Discussion

409 This is the first empirical study to investigate the provision of information on strictly 410 unconditional, individual health impacts (direct health co-benefits) of household climate change 411 mitigation actions on householders' preferences to adopt such actions. Our results suggest that 412 being presented with evidence on direct health co-benefits *does* have an appreciable influence on 413 stated preferences to adopt mitigation actions. Specifically, households receiving information on 414 health impacts reported a greater mean willingness to adopt actions from the sectors housing 415 and—particularly—, food, and were even more likely to report the highest level of willingness 416 ('very willing') to adopt one or more such actions overall. Moreover, the simulated carbon 417 footprint reductions under the requirement to reduce carbon emissions were appreciably greater 418 overall and in relation to food/dietary change actions as a group.

419 European households thus seem to be more willing to implement a given mitigation 420 action when given additional information on health co-benefits that arise irrespective of whether 421 others join in, or not. This is a crucial difference compared to previous research (Bernauer & 422 McGrath, 2016; Myers et al., 2012) where health co-benefits were typically contingent on 423 others' behavior. The present results therefore suggest that direct health benefits for the person 424 performing the mitigation action can be a convincing factor when deciding on whether to 425 perform the mitigation action. Please note that these results hold for a *given* mitigation action 426 where additional information on health co-benefits is provided.

The present results, however, also show that mitigation actions that do exert health effects (e.g., shifting from car to public transport) were generally favored less by European households compared to actions that do not exert health effects (e.g., buying more efficient electrical appliances). Also, this intrinsic difference between actions that do versus do not provide health co-benefits was considerably stronger than the difference between providing versus not providing information on health co-benefits for only those actions that do provide health cobenefits.

Taken together, European households' willingness to implement mitigation actions varied significantly depending on the type and sector of the action. However, households' willingness to implement mitigation actions for which direct health co-benefits exist could be increased by making these benefits explicit.

Households did not decrease their simulated carbon footprint in the categories housing
and mobility. While households' emissions reduction in the sector food is in line with
participants' stated preferences to adopt these actions in their household, the lack of emission
reduction in the sector housing is somewhat surprising. A potential explanation for this finding is
that the average reduction *potential* of the health-related housing options available to the

households was too low to yield an effect. In fact, differential effects of intent-oriented actions as
opposed to impact-oriented actions are in line with previous research. In a sample of Dutch
households, participants who indicated they behaved more pro-environmentally did not
necessarily consume less energy, and actual household energy use was not reflected by proenvironmental intentions, but rather by household characteristics such as income and household

448 size (Gatersleben, Steg & Vleg, 2002).

449 Informing about health co-benefits did not increase preferences to adopt in the sector 450 mobility. This finding is in line with previous research demonstrating that behavioral changes in 451 the mobility sector seem particularly difficult to achieve (Goessling, 2017), or particularly 452 dependent on environmental-friendly attitudes (Bopp, Kaczynski & Wittman, 2011)..Another 453 explanation is that the positive health effects of physical activity are more salient among 454 households, than those of a vegetarian diet or well-insulated houses. Thus, providing additional 455 information about the positive health effect of being more physically active might not make a 456 difference.

The effect of the health information is fairly modest (just a few percent greater overall reduction in carbon emissions among those given health information, for example a 4.5 % age point difference for dietary actions), but given that providing semi-quantitative information on direct health co-benefits is an inexpensive and easy-to-implement strategy, these benefits come at low costs.

However, future studies need to demonstrate if and under what circumstances a larger
effect of informing about health co-benefits can be achieved. Potential approaches may entail (a)
changing the format or (b) the context in which the health information is presented.

465 Concerning presentation format, the health information consisted of only an indication of the 466 strength and direction of the health effect for each particular action in form of small plus or 467 minus symbols. This may easily be substituted by adding more salient pictographic information 468 about the specifics of the health benefit, (e.g., a heart symbolizing heart diseases) to potentially 469 strengthen its impact.

470 Concerning presentation context, while providing only rudimentary information on health 471 itself, we presented participants with a rather great amount of information overall, entailing not 472 only health effects of mitigation options, but also associated costs and carbon reduction, all of 473 which were competing for participants' attentional resources. While it is likely that the 474 conditions in our study are more realistic than the artificial setting in many laboratory studies 475 which test the effect of only one piece of information at a time, presenting heath information as 476 the only source of information might yield stronger effects.

477 Future research could also estimate the impact of providing information on direct health co-478 benefits versus public health co-benefits on citizens' willingness to implement mitigation 479 actions. This could be done by providing one group of households with information on direct 480 health co-benefits, and a second group with information on public health co-benefits of the same 481 mitigation actions. Moreover, it might me worthwhile to include actions of personal preferences 482 or beliefs regarding health. It could be the case, for example, that the present results were driven 483 mainly by participants who have comparatively high preferences for healthy life choices, 484 particularly since a positive relationship between health behaviors and climate mitigation 485 behavior has been demonstrated (Geiger, Otto & Schrader, 2018). Such research could further 486 elucidate the motivational factors that drive citizens' willingness to implement mitigation 487 actions.

488 Unlike typical framing studies, the present study did not employ subtle changes in 489 message wording to test the health argument's effectiveness. Rather, we simply provided 490 participants with health outcomes as an additional piece of information. Thus, we refrained from 491 using persuasive message wordings which might have questionable long-term effects (de Vreese, 492 2004; Druckman, & Nelson, 2003) and have been considered manipulative and undemocratic 493 (Fischhoff, 2013). The present results are therefore unlikely to depend on subtle differences in 494 message wording - something which might be hard to accommodate for in real-life 495 communication campaigns (cf. Lecheler, & de Vreese, 2013).

496 On a more general level, what might be the relevance of our results for climate policy 497 strategies and actions? Relying on the insights from our study, we would argue for the following 498 policy recommendations: (a) Whenever relevant, direct health co-benefits should be included in 499 public communication supporting the introduction of new climate policy measures; and (b) 500 introducing health co-benefits in climate policy discourse can trigger a mechanism known as 501 'policy redressing'. In policy redressing, old programs, for example to mitigate local air 502 pollution, are renewed by linking them up to new climate policy initiatives. A survey of current 503 climate policies in the four countries investigated in the HOPE project showed that health to a 504 very limited degree is linked in any way with policies aimed at reducing households' greenhouse 505 gas emissions (Moberg et al., submitted). The present results suggest there might be a potential 506 for 'redressing' the climate policy discourse by including unconditional, individual health co-507 benefits.

508 Our study provides empirical support for the idea that, in European high-income
509 countries, linking up climate policies with direct health effects can support GHG mitigation
510 efforts at two levels: Firstly, by accruing the individual citizen (actor), this can lead to small, but

- 511 tangible results on households' willingness to adopt suggested climate friendly consumption
- 512 changes. Secondly, our findings support the idea that underlining the potential health co-benefit
- 513 may increase the public acceptance of regulation of private consumption to reduce household
- 514 carbon footprint.
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Declarations of interest

- 520 Declarations of interest: none

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