THE COST-EFFECTIVENESS OF ANTENATAL MALARIA PREVENTION IN SUB-SAHARAN AFRICA

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Abstract. Antimalarial chemoprophylaxis during pregnancy significantly increases the birth weight of babies born to primigravidae, but coverage in sub-Saharan Africa is very limited. This analysis assessed whether increasing coverage is justified on cost-effectiveness grounds. A standardized modeling framework was used to estimate ranges for the cost per discounted year of life lost averted by weekly chloroquine chemoprophylaxis and intermittent sulfadoxine-pyrimethamine (SP) treatment for primigravidae in an operational setting with moderate to high malaria transmission. The SP regimen was found to be more cost-effective than the chloroquine regimen, because of both lower costs and higher compliance. Both regimens appear to be a good value for money in comparison with other methods of malaria control and based on rough cost-effectiveness guidelines for low-income countries, even with high levels of drug resistance. However, extending the SP regimen to all gravidae and increasing the number of doses per pregnancy could make the intervention significantly less cost-effective.

INTRODUCTION

The Global Malaria Control Strategy of the World Health Organization (WHO) advocates that all pregnant women in malaria-endemic areas receive regular chemoprophylaxis against malaria. In practice, the coverage of effective antenatal malaria prevention in sub-Saharan Africa is very limited. A survey of 4 African countries found that only 1–18% of women reported following an antimalarial drug regimen close to the WHO recommendation. Common problems are that provision in government health facilities is limited, antenatal care attendance inadequate, compliance with therapy poor, and resistance to the most widely used drug widespread. This analysis assesses whether, in view of these constraints, a drive to increase coverage in endemic areas of sub-Saharan Africa is justified on cost-effectiveness grounds.

Pregnant women are particularly vulnerable to malaria. Infection may cause harmful effects for the mother, and placental parasitemia retards the growth of the fetus and increases the prevalence of low birth weight (LBW), the proportion of newborns weighing less than 2,500 g. This is of particular concern because LBW is associated with increased neonatal mortality. In areas of high transmission, the effects are most marked in women during their first pregnancy (primigravidae). Antimalarial chemoprophylaxis during pregnancy has been shown to reduce the risk of malaria infection in all pregnant women and to increase significantly the birth weight of babies born to primigravidae. To date, the most commonly used regimen is weekly chloroquine (CQ). Chloroquine resistance has been recorded across Africa, although levels vary considerably: in some areas, particularly in West Africa, the prevalence of treatment failure remains under 25%, but rates of over 50% have been recorded in Kenya, Tanzania, Ethiopia, and Rwanda (WHO/AFRO, unpublished data). Intermittent treatment with sulfadoxine-pyrimethamine (SP) is now considered a potential alternative and has been introduced in Malawi and approved in Kenya. Apart from lower levels of drug resistance, SP has the additional advantage that it is taken in only 2 or 3 doses and is therefore likely to have higher rates of compliance than the weekly CQ regimen.

To decide on the priority to attach to antenatal prevention, policy makers need information on cost-effectiveness and on the way cost-effectiveness changes with increasing drug resistance. Although a few studies are available, they relate to a limited number of settings and intervention specifications, and the results are not comparable with cost-effectiveness estimates for other malaria control interventions because of the partial costing methodology and/or intermediate outcome measures used. A standardized modeling framework was used to provide estimates of the cost-effectiveness of the CQ and SP regimens in an operational setting with moderate to high malaria transmission. The model combined data from the Cochrane meta-analysis of strategies for preventing malaria in pregnancy with information on costs and compliance from a range of published and unpublished sources. To facilitate comparisons with other methods of malaria control and interventions for other health problems, the full incremental costs of the interventions were included, and the benefits were expressed in terms of a generic health outcome measure. The impact on cost-effectiveness of different levels of antimalarial drug resistance and changes in the intervention specification were explored.

METHODS

Interventions. The interventions considered were the provision of CQ chemoprophylaxis or SP intermittent treatment as an addition to standard antenatal care (ANC) services, with antimalarial drugs distributed during 2 ANC visits in the second and third trimesters. The regimens considered were a weekly CQ dose of 300 mg, with tablets prescribed at ANC visits and taken home by women, and 2 doses of SP, 1,500 mg/75 mg, taken during ANC visits. Both interventions were considered initially for primigravidae only.

Modeling framework. This analysis formed part of a larger study that used a standardized analytical framework to calculate comparable cost-effectiveness estimates for a range of interventions to prevent malaria and improve its treatment in sub-Saharan Africa. Effectiveness was cal-
culated in terms of discounted years of life lost (DYLLs) averted, the mortality component of the disability adjusted life year (DALY). The analysis was performed for 3 economic strata, defined on the basis of per capita gross national product (GNP): very low income (under US$315), middle income (between US$315 and $1,000), and higher income (above US$1,000). Discounted years of life lost averted were combined with costs to produce a likely range for the cost per DYLL averted of each intervention. To capture the uncertainty and variability surrounding many of the parameter estimates, probabilistic sensitivity analysis was used, meaning that ranges and probability distributions were attached to model input variables, and the cost-effectiveness outcome was calculated as a probability distribution, rather than a single point estimate.

Calculating effectiveness. Effectiveness estimates were based on the Cochrane meta-analysis of chemoprophylaxis, which found a significant increase in the birth weight of children born to primigravidae (but not multigravidae). The sample sizes of the studies in the meta-analysis were too small to demonstrate a significant impact on neonatal mortality, so in order to calculate DYLLs averted, a model of birth weight and child survival was developed to extrapolate from the increase in birth weight to a reduction in neonatal mortality, and the fall in mortality was converted to DYLLs. Unlike other analyses, which have used a birth weight cutoff of 2.500 g and have ascribed different neonatal mortality rates to children above and below this threshold, this model incorporated the continuum of birth weights and birth weight–specific mortality.

Birth weight distribution. An initial birth weight distribution for children born to primigravidae in the absence of any antimalaria intervention was constructed using empirical estimates. It was assumed that birth weights were normally distributed and, therefore, fully defined by the mean μ and standard deviation σ. It would be more accurate to describe birth weights as a combination of 2 distributions: a predominantly normal distribution, and a second residual distribution for low birth weights with a much lower mean. By assuming that the whole distribution is normal, the prevalence of LBW and therefore effectiveness will be slightly underestimated, making the estimates of cost-effectiveness conservative. The relative frequency f of birth weight x kg was calculated from the area under the normal distribution, assuming that birth weights lie in the range of 1.0 to 4.5 kg and grouping birth weights to the nearest 0.1 kg. A range of estimates for μ and σ for babies born to women who had not received chemoprophylaxis or intermittent treatment was gathered from published data on birth weights in primigravidae in malarious settings in sub-Saharan Africa. 20–23

Birth weight–specific neonatal mortality. Birth weight–specific neonatal mortality m_x kg was defined as the proportion of live newborns of birth weight x kg who die during the neonatal period, the first 28 d of life. Data on continuous birth weight–specific mortality for Africa are very limited, and empirical observations were restricted to 2 studies, conducted in The Gambia and Malawi. A Wilcoxon-Russell model of mortality, was fitted directly to the empirical neonatal mortality observations (see Figure 1a). Empirical m_x kg values were assumed to have binomial errors, and equation 1 was fitted iteratively in Microsoft Excel to maximize the likelihood between the observed and fitted values. 29 The fitted estimates for the parameters in equation 1 were k = 0.0249, a = 6.3749, a’ = −37.2204, and b = 4.3301.

The crude neonatal mortality rate (NNMR) was calculated from the f_x distribution and m_x curve as

\[ \text{NNMR} = \sum_{x=1.0}^{4.5} f_x m_x. \]  (2)

Effect of antimalarial drugs. Antenatal chemoprophylaxis or intermittent treatment was considered to lead to a shift in the birth weight frequency distribution, increasing μ by 1 kg relative to no protection. The proportion of high birth weights (greater than 3.5 kg) was assumed to be the same with or without the intervention. 22,29 It follows that if μ_u and σ_u describe the birth weight distribution in women who did not receive the intervention (referred to as “unprotected”)
and $\mu_x$ and $\sigma_p$, the distribution in those who did (referred to as “protected”), then

$$\frac{3.5 - \mu_x}{\sigma_p} = \frac{3.5 - \mu_p}{\sigma_p}.$$  \(3\)

The average magnitude of $i$ was reported in the Cochrane meta-analysis.\(^9\) In studies including all pregnant women, the overall impact on birth weight was not significant, but for primigravidae alone, the mean birth weight in the protected group was significantly greater than in the unprotected group (difference of 0.101 kg, with a standard deviation of 0.042 kg). It was assumed that the efficacy data from the meta-analysis were based on study cohorts with close to full compliance, in the absence of drug resistance, and therefore represent the maximum effectiveness that could be achieved. Although none of the studies included in the meta-analysis used the SP regimen, it was assumed that the same effect on birth weight would be obtained, as this regimen is highly effective in decreasing the prevalence of both placental parasitemia and severe maternal anemia.\(^{31,32}\)

For a given value of $i$, the difference $d$ between NNMRs in protected and unprotected primigravidae was calculated from the birth weight distributions and the $m_t$ schedule (see Figure 1b). The NNMR difference $d$ is the maximum level of effectiveness achieved in a situation of complete parasite sensitivity to the drugs, where all pregnant woman complied fully with the correct regimen and there were no stillbirths. This will not occur in practice for several reasons: resistance to CQ is already high in many areas of Africa, and resistance to SP is starting to grow; attendance at antenatal services is often intermittent; and compliance with drug therapy has been shown to be very low.\(^{12,33,34}\) It was therefore necessary to estimate the effective NNMR difference $D$, which adjusts $d$ by taking into account the level of drug resistance $r$, the probabilities of attending the antenatal clinic in the first or second trimester, $v_1$, returning for a second clinic visit, $v_2$, and compliance with the correct drug regimen, $g$, and the stillbirth rate $s$, as it was assumed that neither intervention affects the primigravidae stillbirth rate.\(^{30}\) It was assumed that the regimens were only effective if women made their first visit before the end of the second trimester and made at least 2 visits per pregnancy. These are conservative assumptions, as it is possible that there would be some beneficial effects if chemoprophylaxis or intermittent treatment were started later in the pregnancy or if fewer doses were received.\(^{32}\)

Full compliance implies that the drugs were procured, the correct dose taken, and the doses taken at the correct time. Patient compliance $g$ was considered to equal 1 minus the probability of underdosing, with the effects of overdosing and incorrect timings ignored, although overdosing could potentially lead to toxic side effects. Estimates for underdosing with CQ prescribed during ANC were taken from 2 studies in Malawi.\(^{11,12}\) No data were available for compliance with SP. It was assumed that because it is taken in a series of single doses rather than on a weekly basis, compliance would be much higher than with CQ, though not necessarily 100%, because, for example, patients may be anxious about side effects. Compliance with SP was therefore set at between 85% and 95% for each dose. In some cases of non-compliance with CQ, only minor underdosing may occur; its impact on effectiveness is not known, but it is very unlikely that all those who underdose receive zero effects. In the absence of drug resistance, it was assumed that underdosing with the 2-dose SP regimen led to zero effectiveness, but the proportion of underdosed CQ cases where the drug was still effective ($z$) was set between 0.1 and 0.3.

The effective reduction in the NNMR $D$ was calculated as

$$D = d \times (1 - s) \times (1 - r) \times (g + [1 - g]z) \times v_1 \times v_2.$$  \(4\)

The parameter values for $s$, $v_1$, $v_2$, $z$, and $g$ were estimated from the published literature and from consultation with researchers and practitioners. $D$ was calculated as a function of resistance $r$, which was allowed to vary between 0 (complete sensitivity) and 1 (complete resistance). Resistance $r$ was defined as RII/RIII resistance (no clearance of parasitemia within 7 d of drug administration). It was assumed that parasitological cure was required to increase birth weight.\(^3\)

The possibility of side effects from the drug regimens was not included, but they are not expected to be important. Chloroquine is known to be safe in pregnancy. Fatal side effects have been observed with SP at a rate of between 0 and 1:25,000 when used as weekly prophylaxis,\(^{35}\) but side effects are likely to be much less frequent with intermittent treatment. Studies of SP use in pregnancy have not found increased risk of spontaneous abortion, congenital defects, jaundice, or kernicterus among neonates.\(^{32,36,37}\)

Translating mortality reduction to DYLLs averted. The effective reduction in the NNMR, $D$, was converted to years of life lost (YLLs) averted per primigravida. Years of life lost were calculated using age-specific life expectations calculated from a West African life table with a life expectancy at birth of 50 yr for very-low-income and middle-income countries and from a General Pattern life table with a life expectancy at birth of 65 yr for higher-income countries.\(^{38}\) Years of life lost were discounted at 3%, and no age weighting was applied. All effectiveness input parameters are listed in Table 1.

Costs. Costs were calculated using the ingredients approach,\(^{39}\) with data obtained through reviews of published and unpublished literature, program budgets, price catalogs, and consultation with researchers and program managers. The incremental cost per woman was based on adding the service to an existing ANC program and included the cost of staff training, the production of health education materials, the drugs, and incremental staff time.

All cost input variables are listed in Table 2. Implementing the intervention would require additional time from ANC and supervisory staff. The cost of this time was included, because staff could be undertaking other useful activities if they were not involved in this intervention (i.e., their time has an opportunity cost). It was estimated that the intervention would add 10 min to the first visit and 5 min to the second visit, and that an additional 15 min of supervisory time would be required per month. Estimates of costs of full-time equivalent staff for Safe Motherhood interventions were extrapolated to provide salary estimates for the 3 economic strata.

A single price was used for other inputs, because they are
traded internationally. A full course of SP consists of 2 doses of 1.500 mg/75 mg, and a full course of CQ consists of 16 doses of 300 mg, based on an average first recruitment time of 22 wk and an average gestational age at delivery of 38 wk. Transport, insurance, and delivery were assumed to add an additional 25% to the drug price, and 25% of drugs were assumed to be wasted.

Annualized costs were also included for health education materials and training, which were annualized using a discount rate of 3%, and all costs were converted to 1995 US$. Using the US$ period average market exchange rate in the study year and the U.S. Consumer Price Index, and all costs were converted to 1995 US$, materials and training, which were annualized using a discount rate of 3%, and all costs were converted to 1995 US$, using the US$ period average market exchange rate in the study year and the U.S. Consumer Price Index.

The incremental cost varies with $v_2$ (probability of returning for second clinic visit), as the drug and staff costs of the additional visit would not be incurred if the woman failed to return. It was assumed that costs were independent of compliance, as the drug cost would be incurred whether or not the tablets were taken as prescribed, and independent of $v_1$ (attendance before the end of the second trimester), as drugs would be prescribed even when the first visit took place during the third trimester (although it was assumed that this would result in zero effectiveness).

### Cost-effectiveness

The incremental cost-effectiveness ratio (CER) was calculated as the cost per primigravida divided by the number of DYLLs averted per primigravida. The CER outcomes were calculated as a function of parasitological resistance $r$ between 0 and 0.99.

#### Sensitivity analysis

Sensitivity analysis was conducted on the values assigned to model input values and the intervention specifications.

#### Input variables

To allow for the high degree of uncertainty and variability surrounding many of the parameters, probability distributions based on available empirical data were assigned to the cost and effectiveness model inputs. Truncated normal distributions were assigned to the cost and effectiveness model inputs. For other variables, if only range information was known, a uniform distribution was assigned; if available data showed a peaked distribution, a triangular distribution was used. For each intervention scenario, the models were iterated using Monte Carlo simulations with Palisade @RISK software, and the CER was expressed as a probability distribution.

#### Table 1

Effectiveness input variables

<table>
<thead>
<tr>
<th>Input variable</th>
<th>Probability distribution</th>
<th>Distribution parameters</th>
<th>Source</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean birth weight in unprotected primigravida, $\mu_u$ (kg)</td>
<td>Truncated normal</td>
<td>Mean = 2.788, Standard deviation = 0.082, Min = 2.690, Max = 2.875</td>
<td>Greenwood and others, 1989; Greenwood and others, 1989; Fleming and others, 1986</td>
</tr>
<tr>
<td>Standard deviation of birth weight, $\sigma_u$ (kg)</td>
<td>Truncated normal</td>
<td>Mean = 0.476, Standard deviation = 0.098, Min = 0.36, Max = 0.57</td>
<td>Greenwood and others, 1989; Fleming and others, 1986</td>
</tr>
<tr>
<td>Increase in birth weight, $i$ (kg)</td>
<td>Truncated normal</td>
<td>Mean = 0.101, Standard deviation = 0.042, Min = 0.003, Max = 0.198</td>
<td>Gülmezoglu and Garner, 1998</td>
</tr>
<tr>
<td>Stillbirth rate in primigravida, $s$</td>
<td>Uniform</td>
<td>Min = 0.062, Max = 0.116</td>
<td>Greenwood and others, 1994</td>
</tr>
<tr>
<td>Probability of initial clinic visit in first or second trimester, $v_1$</td>
<td>Triangular</td>
<td>Min = 0.540, Max = 0.936, BE = 0.868</td>
<td>Stewart and others, 1997</td>
</tr>
<tr>
<td>Probability of returning for second clinic visit, $v_2$</td>
<td>Triangular</td>
<td>Min = 0.835, Max = 0.989, BE = 0.937</td>
<td>Stewart and others, 1997</td>
</tr>
<tr>
<td>Probability of returning for third clinic visit, $v_3$ (relevant where HIV prevalence is high)</td>
<td>Triangular</td>
<td>Min = 0.481, Max = 0.942, BE = 0.796</td>
<td>Stewart and others, 1997</td>
</tr>
<tr>
<td>Probability of compliance to prescribed drug, $g$</td>
<td>Uniform</td>
<td>Min = 0.25, Max = 0.57</td>
<td>Heymann and others, 1990; Helitzer-Allen and others, 1993</td>
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<tr>
<td>Chloroquine (average for whole course)</td>
<td>Uniform</td>
<td>Min = 0.85, Max = 0.95</td>
<td>Greenwood and others, 1995; Fleming and others, 1986</td>
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<tr>
<td>SP (per dose)</td>
<td>Uniform</td>
<td>Min = 0.72, Max = 0.90</td>
<td>Greenwood and others, 1995; Fleming and others, 1986</td>
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<tr>
<td>SP (for 2 doses)</td>
<td>Uniform</td>
<td>Min = 0.61, Max = 0.86</td>
<td>Greenwood and others, 1995; Fleming and others, 1986</td>
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<tr>
<td>SP (for 3 doses)</td>
<td>Uniform</td>
<td>Min = 0.1, Max = 0.3</td>
<td>Greenwood and others, 1995; Fleming and others, 1986</td>
</tr>
</tbody>
</table>

Note: $\mu_u$, $\sigma_u$, $v_1$, $v_2$, $v_3$, $s$, $g$ are parameters used in the model. $\sigma_u$ is the standard deviation of birth weight, $v_1$ is the probability of initial clinic visit in first or second trimester, $v_2$ is the probability of returning for second clinic visit, $v_3$ is the probability of returning for third clinic visit, $s$ is the stillbirth rate in primigravida, $g$ is the probability of compliance to prescribed drug.
TABLE 2
Cost input variables (all costs in 1995 US$)

<table>
<thead>
<tr>
<th>Drug Cost</th>
<th>Probability distribution</th>
<th>Minimum</th>
<th>Maximum</th>
<th>Best estimate</th>
<th>Source</th>
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</thead>
<tbody>
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<td>Chloroquine</td>
<td>Triangular</td>
<td>0.008</td>
<td>0.013</td>
<td>0.01</td>
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<td>Tablets per dose</td>
<td>2</td>
<td></td>
<td></td>
<td>WHO, Model Prescribing Information: Drugs Used in Parasitic Diseases, 1995</td>
</tr>
<tr>
<td></td>
<td>Doses per pregnancy</td>
<td></td>
<td></td>
<td></td>
<td>WHO, Model Prescribing Information: Drugs Used in Parasitic Diseases, 1995</td>
</tr>
<tr>
<td>Sulfadoxine-pyrimethamine</td>
<td>Triangular</td>
<td>0.032</td>
<td>0.040</td>
<td>0.036</td>
<td>Management Sciences for Health, International Drug Price Indicator Guide, 1996</td>
</tr>
<tr>
<td></td>
<td>Tablets per dose</td>
<td>3</td>
<td></td>
<td></td>
<td>WHO, Model Prescribing Information: Drugs Used in Parasitic Diseases, 1995</td>
</tr>
<tr>
<td></td>
<td>Doses per pregnancy</td>
<td>2</td>
<td></td>
<td>25%</td>
<td>Foster, 199144</td>
</tr>
<tr>
<td></td>
<td>Delivery cost as percentage of warehouse cost</td>
<td></td>
<td></td>
<td></td>
<td>25% Foster, 199144</td>
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<tr>
<td>Salaries per annum (US$)</td>
<td>Health center staff</td>
<td></td>
<td></td>
<td></td>
<td>Extrapolation from Tinker &amp; Koblinsky 1992, plus and minus 15%</td>
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<td></td>
<td>Very-low-income countries</td>
<td>Uniform</td>
<td>1,934</td>
<td>2,617</td>
<td>Extrapolation from Tinker &amp; Koblinsky 1992, plus and minus 15%</td>
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<td></td>
<td>Middle-income countries</td>
<td>Uniform</td>
<td>2,544</td>
<td>3,442</td>
<td>Extrapolation from Tinker &amp; Koblinsky 1992, plus and minus 15%</td>
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<td></td>
<td>Higher-income countries</td>
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<td>7,312</td>
<td>9,893</td>
<td>Extrapolation from Tinker &amp; Koblinsky 1992, plus and minus 15%</td>
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<td></td>
<td>Supervisory staff</td>
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<td>Extrapolation from Tinker &amp; Koblinsky 1992, plus and minus 15%</td>
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<td>Very-low-income countries</td>
<td>Uniform</td>
<td>3,137</td>
<td>4,244</td>
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<td>4,127</td>
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<td>Hours worked per day</td>
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<td>Health center incremental staff time</td>
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<td>Per initial visit (min)</td>
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<td></td>
<td></td>
<td>Per initial visit (min)</td>
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<tr>
<td></td>
<td>Per subsequent visit (min)</td>
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<td></td>
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<td>Supervisory incremental staff time per health center per month (min)</td>
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<td>Health Education</td>
<td>Uniform</td>
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<td>14.81</td>
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<td></td>
<td>Cost per book/poster (US$)</td>
<td>Uniform</td>
<td>4.11</td>
<td>5.55</td>
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<td></td>
<td>Number of flipcharts per clinic</td>
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<td>1</td>
<td></td>
<td>Picard and others, 199355</td>
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<tr>
<td></td>
<td>Number of books/posters per clinic</td>
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<td></td>
<td>Picard and others, 199355</td>
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<tr>
<td></td>
<td>Useful life of health education materials (yr)</td>
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<td>4</td>
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<td>Picard and others, 199355</td>
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<tr>
<td>Training</td>
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<td></td>
<td>Picard and others, 199355</td>
</tr>
<tr>
<td></td>
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<td>Picard and others, 199355</td>
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<td>Useful life of training (yr)</td>
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<td>Picard and others, 199355</td>
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<td>209</td>
<td>1,056</td>
<td>Ogunbekun and others, 199646 Hanson and Chindele, 199224 Hanson and Nkunzimana, 199229</td>
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<td></td>
<td>Primigravidae as a proportion of antenatal attendees</td>
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<td>0.25</td>
<td>Approximation based on data from Stewart and others, 1997, World Bank 1997, Steketee and others, 1996, Mnyika and others, 199570</td>
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<td></td>
<td>Number of visits per pregnancy</td>
<td>Triangular</td>
<td>0.7</td>
<td>5.6</td>
<td>Stewart and others, 199747</td>
</tr>
</tbody>
</table>

Variables had reached convergence (defined as a percentage change in the output mean, standard deviation, and percentiles of less than 1.5% after each 100 iterations). Summary statistics presented for the model outputs were the mean and range within which 90% of the cost-effectiveness estimates fell, termed the “cost-effectiveness range.” The evaluation of CERs rests on a comparison with the cost-effectiveness of alternative uses of resources, which will depend on the
specific context. However, based on rough guidelines for countries with per capita GNP below US$765, the intervention was classified as an "attractive" use of resources if the CER range fell entirely below US$150.46

Intervention specifications. For SP intermittent treatment, the robustness of the cost-effectiveness conclusions was tested by considering 2 further treatment scenarios. First, the 2-dose SP regimen was extended to include all gravidae, to capture probable actual practice. The Cochrane review did not report a significant impact on birth weight for multigravidae, so the increase in birth weight \( i \) was set at zero for multigravidae by multiplying the effective reduction in NNMR, \( D \), by the proportion of antenatal attendees who are primigravidae. The cost per woman was reduced to reflect the allocation of the fixed costs of supervision, training, and health education over more women per facility (total costs would of course rise). Second, a 3-dose SP regimen was considered also to be given to all gravidae, as a possible response to increasing human immunodeficiency virus (HIV) prevalence. With this specification, costs were raised to include the additional dose of SP and an extra clinic visit. To capture lower patient compliance with the 3-dose regimen, \( g \) was reduced, and the effective reduction in NNMR \( D \) was multiplied by a parameter \( v_s \), the probability of returning for a third clinic visit. The other effectiveness assumptions and the probability distribution of unprotected birth weights were not altered from the 2-dose all-gravidae scenario.

Total cost. To assess affordability, the total cost of full coverage was estimated as a percentage of the existing public-sector health budget for a typical very-low-income country (based on Tanzanian population and health-sector budget data46,47).

RESULTS

At complete drug sensitivity \( (r = 0) \), the DYLLs averted per primigravida with the CQ regimen were on average 0.09 for very-low-income and middle-income countries and 0.10 for higher-income countries. With the 2-dose SP regimen for primigravidae, 0.14 DYLLs were averted per primigravida for very-low-income and middle-income countries and 0.16 for higher-income countries. More DYLLs were averted per woman in higher-income countries because a life table with a higher life expectancy at birth was used, meaning that more YLLs were saved per death averted. Sulfadoxine-pyrimethamine was more effective than CQ even when there was no resistance to either drug, because compliance was assumed to be lower with CQ than with SP.

The mean incremental costs per pregnancy for these CQ and SP regimens are shown in Table 3. The incremental cost with SP was $1.13 in very-low-income countries, $1.25 in middle-income countries, and $2.14 in higher-income countries, the increase in costs in higher-income countries being explained by higher salaries. In very-low-income and middle-income countries, drugs and staff each accounted for approximately one-third of the incremental cost, with health education and training making up the remaining third. In higher-income countries, salary costs made up about two-thirds of the incremental cost, with drugs falling to only 15%. All costs were higher with the CQ regimen because of the higher drug cost per pregnancy. The mean incremental cost with CQ was $1.30 in very-low-income countries, $1.42 in middle-income countries, and $2.31 in higher-income countries. Although the average drug cost per dose was higher with SP than with CQ, the CQ regimen demanded a third clinic visit. The other effectiveness assumptions and the probability distribution of unprotected birth weights were not altered from the 2-dose all-gravidae scenario.

The cost-effectiveness ranges for SP and CQ at complete drug sensitivity \( (r = 0) \) for the 3 economic strata are shown in Figure 2. In very-low-income countries, the range was $4–$27 for SP and $7–$49 for CQ. CQ was slightly less cost-effective than SP, because of both higher costs per primigravida and lower DYLLs averted.

The cost-effectiveness ranges for middle-income countries were little different from those for very-low-income countries. However, the ranges were significantly higher in countries in the higher-income stratum, as the impact of increased costs due to higher salaries outweighed the increased DYLLs averted due to the higher-life expectancy. For all income levels, the cost-effectiveness ranges for both SP and CQ fell clearly below $150 per DYLL averted, meaning that with no resistance, either regimen would be considered an "attractive" option for all economic strata.

The above results were all based on full drug sensitivity.

### Table 3

Expected incremental cost per primigravidae (1995 US$)*

<table>
<thead>
<tr>
<th>Drug regimen</th>
<th>Very low income</th>
<th>Middle income</th>
<th>Higher income</th>
</tr>
</thead>
<tbody>
<tr>
<td>Chloroquine</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Drugs</td>
<td>$0.50 (38%)</td>
<td>$0.50 (35%)</td>
<td>$0.50 (21%)</td>
</tr>
<tr>
<td>Staff</td>
<td>$0.36 (28%)</td>
<td>$0.48 (34%)</td>
<td>$1.38 (59%)</td>
</tr>
<tr>
<td>Health education and training</td>
<td>$0.44 (34%)</td>
<td>$0.44 (31%)</td>
<td>$0.44 (19%)</td>
</tr>
<tr>
<td>Mean incremental cost</td>
<td>$1.30 (100%)</td>
<td>$1.42 (100%)</td>
<td>$2.31 (100%)</td>
</tr>
<tr>
<td>Sulfadoxine-pyrimethamine (2 doses, primigravidae only)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Drugs</td>
<td>$0.32 (29%)</td>
<td>$0.32 (26%)</td>
<td>$0.32 (15%)</td>
</tr>
<tr>
<td>Staff</td>
<td>$0.36 (32%)</td>
<td>$0.48 (38%)</td>
<td>$1.38 (64%)</td>
</tr>
<tr>
<td>Health education and training</td>
<td>$0.44 (39%)</td>
<td>$0.44 (36%)</td>
<td>$0.44 (21%)</td>
</tr>
<tr>
<td>Mean incremental cost</td>
<td>$1.13 (100%)</td>
<td>$1.25 (100%)</td>
<td>$2.14 (100%)</td>
</tr>
</tbody>
</table>

* Percentage of mean incremental cost in parentheses.
Cost-effectiveness implications of the model variations are shown in Figure 4, which compares the relationship between the upper limit of the cost-effectiveness range and the level of RII/RIII resistance for very-low-income countries for 3 specifications of the SP regimen: 2 doses per pregnancy for primigravidae only, 2 doses for all gravidae, and 3 doses for all gravidae (assuming zero benefits for multigravidae). If a 2-dose regimen were provided to all gravidae, the cost-effectiveness range at zero resistance would be $12–$70, and the level of RII/RIII resistance up to which the cost-effectiveness range remained lower than $150 would fall to 54%. Providing 3 doses to all gravidae further increased the cost-effectiveness range at zero resistance to $22–$129, reducing the $150 threshold to 14% resistance.

The expected total cost to government of full coverage with each regimen is shown in Table 4. The CQ regimen in a typical very-low-income country represented 0.19% of the existing government health care budget, the 2-dose SP regimen for primigravidae 0.17%, the 2-dose regimen for all gravidae 0.97%, and the 3-dose regimen for all gravidae 1.19%.

DISCUSSION

The model provides a framework to investigate the operational cost-effectiveness of alternative malaria prevention strategies in pregnancy and allows the impact of variations in drug resistance, intervention specification, and socioeco-
nomic setting to be explored. The results demonstrate the relative cost-effectiveness of SP intermittent treatment compared with CQ prophylaxis, even with zero resistance to both drugs. The relative attractiveness of the SP regimen is driven by the assumed higher compliance and lower costs per woman. Even with SP RII/RIII resistance levels of over 40%, the SP regimen would on average still be more cost-effective than the CQ regimen with zero CQ resistance. Because current levels of resistance to CQ in sub-Saharan Africa are much higher than those for SP, in practice the difference in cost-effectiveness between the 2 drug regimens would be accentuated. However, in addition to the costs and effects included in the model, the choice between the regimens will be affected by other concerns, such as the attitude of policy makers to the risk of SP side effects and the potential impact of widespread SP use on the growth rate of SP resistance. The importance of compliance in influencing the effectiveness of a regimen highlights the need to investigate interventions to improve adherence to CQ chemoprophylaxis through, for example, the use of coated tablets, prepackaging of drugs, or intensified health education, and to evaluate the effectiveness of intermittent treatment regimens with CQ. This analysis formed part of a larger study to investigate the cost-effectiveness of a range of interventions to prevent malaria and to improve its treatment in sub-Saharan Africa. The use of a consistent modeling framework allows the cost-effectiveness of the different interventions to be compared. For example, in a very-low-income country, the cost-effectiveness range for insecticide treated nets (ITNs) for children under 5 yr was $19–85 per DALY averted, and $4–10 where there is already a high degree of net ownership and only insecticide treatment is provided. For residual spraying, the cost-effectiveness range was $16–29 per DALY averted with 1 spraying round per year, and $32–58 with 2 rounds. Even though the analysis presented here considers only the mortality component of the DALY, making the effectiveness estimates conservative, the results are of comparable cost-effectiveness with these other malaria prevention methods. Because of the considerable overlap in the cost-effectiveness ranges, no single intervention emerges as clearly the most cost-effective. However, up to SP RII/RIII resistance levels of 65%, the 2-dose SP regimen for primigravidae would be at least as cost-effective as ITNs for children under 5 yr (including the provision of nets). Chloroquine prophylaxis would be at least as cost-effective as ITNs up to 40% CQ resistance.

Rough cost-effectiveness guidelines for low-income countries classified interventions as an “attractive” use of resources if the cost per DALY averted was below $150. The results show that in the absence of drug resistance, the interventions are clearly a good use of resources. For both CQ and the 2-dose SP regimen for primigravidae, the cost-effectiveness ranges fell below the $150 threshold for all economic strata. Even when resistance is allowed for, both drug regimens remained cost-effective up to high levels of parasitological resistance. The interventions were not only highly cost-effective but also relatively affordable to developing countries, absorbing well under 1.5% of the current health-sector budget.

The interventions appeared highly cost-effective despite the relatively narrow definition of health benefits included. The model incorporated only reduced mortality in the neonatal period, excluding increased survival for children older than 28 d and potential benefits from reductions in morbidity and mortality for mothers. Chemoprophylaxis or intermittent treatment in primigravidae is associated with a significant reduction in the number of malaria episodes treated for the mother relative to unprotected primigravidae. Although this morbidity reduction in mothers would be unlikely to have a significant effect on the DALYs averted (because of the low disability weighting and short duration of a malaria episode), it could lead to significant cost savings for households and providers because of the reduction in treatment seeking. More significantly, the impact on maternal anemia was not incorporated. Malaria infection is strongly associated with moderate and severe anemia in primigravidae, which in turn is associated with increased maternal morbidity and mortality. In fact, the earliest trials of antenatal prophylaxis were prompted primarily by the need to reduce maternal mortality due to severe anemia. Any impact on maternal mortality could have spin-off effects on the health and well-being of the whole household. In Kenya, intermittent SP treatment reduced the prevalence of severe anemia (hemoglobin less than 8 g/dL) among primigravidae to 14.5%, compared with 23.7% in the control group. The impact of severe anemia on maternal mortality is not known, but assuming, for example, that severe anemia doubled the risk of maternal mortality, and that the baseline maternal mortality rate was 450 per 100,000, SP intermittent treatment would avert 0.0002 deaths per primigravida, after adjusting for noncompliance. Converting to DYLLs using an average age at death of 19 yr and a life expectancy of 52 yr, this would be equivalent to 0.005 DYLLs averted per primigravida, increasing the total DYLLs averted per primigravida by the intervention by around 5%.

A number of caveats should be borne in mind in interpreting these promising cost-effectiveness results. First, the modeled impact on child survival has not been empirically demonstrated; second, varying the specification of the SP regimen reduced its estimated cost-effectiveness; and third, the analysis was based on the assumption that ANC services were already in place. These caveats are explored below.

Although the impact of prophylaxis and intermittent treatment on birth weight in primigravidae has been clearly

<table>
<thead>
<tr>
<th>Drug</th>
<th>Regimen</th>
<th>Government cost per annum</th>
<th>Cost as percentage of government health budget</th>
</tr>
</thead>
<tbody>
<tr>
<td>CQ</td>
<td>Weekly doses</td>
<td>$179,349</td>
<td>0.19%</td>
</tr>
<tr>
<td>SP</td>
<td>2-dose regimen, primigravidae only</td>
<td>$155,896</td>
<td>0.17%</td>
</tr>
<tr>
<td></td>
<td>2-dose regimen, all pregnant women</td>
<td>$913,535</td>
<td>0.97%</td>
</tr>
<tr>
<td></td>
<td>3-dose regimen, all pregnant women</td>
<td>$1,119,080</td>
<td>1.19%</td>
</tr>
</tbody>
</table>

* Based on the following estimates for Tanzania: population of 29.2 million, 47% crude birth rate of 42.6 per 1,000; 17% of all births born to primigravidae; 93.2% of primigravidae and 91.8% of all pregnant women receive antenatal care; 70% of population at high risk of malaria; government health budget per annum of $94 million (including donor contributions); no cost recovery.
shown, the sample sizes of the studies included in the meta-analysis were too small to demonstrate a significant impact on neonatal mortality even among primigravidae. It was therefore necessary to model the impact on mortality based on empirical evidence of birth weight distributions and birth weight–specific NNMRs. However, the available data on birth weights and mortality are limited and are unlikely to be representative of the whole of sub-Saharan Africa. One might expect birth weights to be positively correlated with economic development, and birth weight–specific mortality rates negatively correlated. Because the data on birth weights in unprotected primigravidae and birth weight–specific mortality all came from studies conducted in middle-income and very-low-income countries, it is possible that the model overestimates effectiveness in higher-income countries.

In modeling the reduction in the NNMR following an increase in birth weight, no distinction was made between pre-term delivery and intrauterine growth retardation as causes of low birth weight. The NNMR at a given birth weight is higher for a preterm baby than for a full-term baby with retarded growth. Although there is evidence that malaria affects both gestational age and weight for age, no studies have demonstrated that prophylaxis or intermittent treatment reduces the prevalence of preterm births. It is possible that the intervention could have a differential impact, with, for example, a bigger impact on the birth weight for gestational age than on the gestational age at delivery, and may therefore not have the impact on neonatal mortality predicted by the model. The effectiveness of the interventions may therefore be overestimated.

Although the Cochrane review did not report a significant effect of prophylaxis on LBW in multigravidae, in practice it may be inappropriate or impractical to restrict the intervention to primigravidae. For example, it might be logistically complex to provide different services to different groups and politically problematic to exclude some women from what is perceived to be a valuable service. Altering the specification of the SP regimen to include all gravidae, rather than primigravidae only, had a significant impact on the cost-effectiveness results. The SP RII/RIII resistance level up to which the cost-effectiveness range remained below $150 was reduced from 82% to 54%. However, these estimations assumed zero effects for multigravidae further reduced the resistance threshold up to which the cost-effectiveness range remained below $150 to 14%. Again, it is possible that this is a conservative estimate, because, first, there may be positive benefits to multigravidae, and second, HIV infection, by diminishing the capacity of pregnant women to control malaria infection, may increase the prevalence of LBW and potentially the effectiveness of the intervention. On the other hand, it has been argued that monthly SP doses might be required in populations with high rates of seropositivity, increasing the costs of the intervention still further.

The cost analysis was based on the assumptions that ANC services were already in place and that the new intervention could be added to this existing infrastructure. Although in most of sub-Saharan Africa the percentage of primigravidae receiving ANC is high (median of 89% for 17 countries), there are some exceptions. For instance, in Burkina Faso only 61% of primigravidae attended ANC, and in Niger only 32%. The incremental costs to both the facility and to women would be much greater than estimated here if it were necessary to set up an entirely new ANC service, and these costs would depend on many factors, including local unit costs, the package of ANC offered, and the existing infrastructure of health facilities. This variation is demonstrated by ANC costing studies, which provide cost estimates per visit ranging from $0.05–$0.15 in Tanzanian government dispensaries, to $0.75 in the Gambia, and to $8.41 and $12.53 in South Africa (1995 US$).

It is possible that if wide coverage of ANC services were not attainable, community-based health workers such as traditional birth attendants or village health workers (VHWs) could deliver the intervention. Traditional birth attendants were found to be an effective and appropriate channel for providing antimalarial and iron prophylaxis to pregnant women in the Gambia, even though they received no additional payment. However, a well-supported VHW program in Kenya was able to achieve only 29% coverage of primigravidae, leading the authors to conclude that the program was not effective in providing prophylaxis and that asking VHWs to distribute the drugs may have overloaded them.

In summary, assuming the hypothesized link between an increase in birth weight and reduction in NNMR holds, CQ prophylaxis and SP intermittent treatment for primi-
gravidae are cost-effective interventions compared to other methods of malaria control and to interventions for other health problems. This remains the case up to high levels of parasitological resistance to either drug. In addition to being cost-effective, the interventions are also relatively affordable for sub-Saharan African governments. The analysis lends weight to the argument that there should be a drive to extend coverage of preventive strategies in pregnancy to all women in malarious areas of Africa. It also shows that modifying the intervention in response to increasing HIV prevalence could substantially reduce its cost-effectiveness, although the impact of HIV on the burden of disease, the effectiveness of the interventions, and the costs of any intervention modifications required needs further exploration. Finally, because resistance to SP is expected to grow rapidly as it is more widely used in sub-Saharan Africa, research is urgently needed to identify potential replacement drugs to enable the provision of effective antenatal malaria prevention in the future.

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