

1 **Quantifying the public’s view on social value judgments in**  
2 **vaccine decision-making: a discrete choice experiment**

3 **Abstract**

4 Vaccination programs generate direct protection, herd protection and, occasionally,  
5 side effects, distributed over different age groups. This study elicits the general public’s  
6 view on how to balance these outcomes in funding decisions for vaccines. We  
7 performed an optimal design discrete choice experiment with partial profiles in a  
8 representative sample (N=1499) of the public in the United Kingdom in November  
9 2016. Using a panel mixed logit model, we quantified, for four different types of  
10 infectious disease, the importance of a person’s age during disease, how disease was  
11 prevented—via direct vaccine protection or herd protection—and whether the vaccine  
12 induced side effects. Our study shows clear patterns in how the public values  
13 vaccination programs. These diverge from the assumptions made in public health and  
14 cost-effectiveness models that inform decision-making. We found that side effects and  
15 infections in newborns and children were of primary importance to the perceived value  
16 of a vaccination program. Averting side effects was, in any age group, weighted three  
17 times as important as preventing an identical natural infection in a child whereas the  
18 latter was weighted six times as important as preventing the same infection in elderly  
19 aged 65-75 years. These findings were independent of the length or severity of the  
20 disease, and were robust across respondents’ backgrounds. We summarize these  
21 patterns in a set of preference weights that can be incorporated into future models.  
22 Although the normative significance of these weights remains a matter open for  
23 debate, our study can, hopefully, contribute to the evaluation of vaccination programs  
24 beyond cost-effectiveness.

25

26

27 **Keywords**

28 United Kingdom; age; side effects, herd immunity, cost-effectiveness analysis,  
29 decision making; priority-setting, equity

## 30 **1. Introduction**

31 Economic evaluation methods such as cost-effectiveness analysis (CEA) are common  
32 components in public funding decisions for vaccines (Drummond, Sculpher, Torrance,  
33 O'Brien, & Stoddard, 2005; Walker, Hutubessy, & Beutels, 2010). They feature in the  
34 standard evidence considered by e.g. the Advisory Committee on Immunization  
35 Practices in the US, the Joint Committee on Vaccination and Immunization in England,  
36 the World Health Organization and non-governmental organizations such as the Bill &  
37 Melinda Gates Foundation (Ricciardi et al., 2015). At the same time, it is widely  
38 acknowledged that these evaluation frameworks have important shortcomings and  
39 that they alone offer insufficient basis for making fair and efficient vaccine funding  
40 decisions (Cookson, Drummond, & Weatherly, 2009; Dukhanin et al., 2018). There is  
41 a growing literature about the limits of CEA in assessing the value of vaccination  
42 (Barnighausen, Bloom, Cafiero-Fonseca, & O'Brien, 2014; Bloom, 2011; Bloom, Fan,  
43 & Sevilla, 2018; Luyten & Beutels, 2016).

44 One important criticism is that CEA is limited in how it values the consequences of  
45 vaccination. Summary outcome measures [such as e.g. infections prevented or  
46 Quality-Adjusted Life Years (QALYs) gained] neglect the particular social context in  
47 which these outcomes occur. Nonetheless, such contextual features are important  
48 aspects to consider when evaluating a vaccination strategy. Vaccination induces  
49 disease protection in those who become vaccinated, but it also creates *herd* protection  
50 (or indirect effects in third parties because of reduced pathogen transmission (Fine,  
51 Eames, & Heymann, 2011)) and, occasionally, adverse clinical *side* effects. There are  
52 qualitative differences between these direct, herd and side effects. Creating herd  
53 protection can be of particular ethical value (e.g. to protect vulnerable groups who  
54 otherwise cannot protect themselves) and there is a profound psychological impact of

55 vaccine-induced side effects. Moreover, the *distribution* of these three different effect  
56 types over different age groups is important. Side effects can be concentrated in one  
57 age group despite indirect protection from reduced transmission benefitting either the  
58 wider population, or in some cases a different age group entirely (Anderson & May,  
59 1991). Examples include protecting the elderly through childhood influenza  
60 vaccination or future generations through a *polio* eradication program. Such broader,  
61 distributive aspects of vaccination are important but they remain neglected in standard  
62 cost-effectiveness or public health impact models.

63 Several notable examples illustrate that this broader social context of health outcomes  
64 needs to be considered in vaccine decision-making (Schwartz & Caplan, 2017). For  
65 instance, vaccines against rotavirus (Rotashield®) and pertussis (whole cell pertussis  
66 vaccine) were withdrawn from many countries because of a perceived risk of side  
67 effects, even though from a medical perspective the benefit from vaccination largely  
68 outweighed any potential risk (Blume & Zanders, 2006; Granstrom, 2011; Lynch et al.,  
69 2006). Also, despite persuasive economic and public health benefits of childhood  
70 influenza vaccination, few countries have actually implemented such a preventive  
71 strategy, due in large part to concerns about the social acceptability and equity of  
72 targeting vaccination at children to protect the wider population (McGuire, Drummond,  
73 & Keeping, 2016). And, in many countries introduction of an effective varicella  
74 vaccination program has been delayed because of concerns about the possible  
75 ‘exogenous boosting effect’ and its social repercussions, i.e. that reduced chickenpox  
76 transmission among children (due to varicella vaccination) might temporarily increase  
77 shingles incidence among older generations (Luyten, Ogunjimi, & Beutels, 2014).

78 Misjudging ethical norms and social sensitivities in vaccination policy by over-relying  
79 on CEA can have important implications. It may affect the perceived equity of a

80 program, its support by the public and its long-term sustainability (Charo, 2007;  
81 Feudtner & Marcuse, 2001; Salmon et al., 2006; Yaqub, Castle-Clarke, Sevdalis, &  
82 Chataway, 2014) (Hornsey, Harris, & Fielding, 2018; Tomeny, Vargo, & El-Toukhy,  
83 2017). It can invoke public backlash to the vaccine, leading to reduced uptake,  
84 increased vaccine hesitancy and reduced overall effectiveness of the program (Bauch  
85 & Earn, 2004; Bhattacharyya, Bauch, & Breban, 2015; Ndeffo Mbah et al., 2012).  
86 Therefore, an empirical evidence-base is needed about the public's view on the key  
87 value judgments that need to be made in vaccine funding decisions (Bombard,  
88 Abelson, Simeonov, & Gauvin, 2011; Field & Caplan, 2012; Luyten, Dorgali, Hens, &  
89 Beutels, 2013; Makarovs & Achterberg, 2017; Poland & Marcuse, 2011). Such  
90 evidence can complement formalized appraisals like CEA, stimulate deliberation and  
91 discussion on how to prioritize vaccines within a budget constraint and, moreover, it  
92 can be explored whether such evidence can become quantitatively integrated into  
93 formal decision frameworks in some sort of 'extended' or 'weighted' CEA (Cookson et  
94 al., 2009; Fleurbaey, Luchini, Muller, & Schokkaert, 2013).

95 The objective of this study is to address this challenge by analyzing how the population  
96 in the United Kingdom prioritizes vaccination programs and to investigate whether its  
97 values diverge from the assumptions that are implicitly underlying CEA. We use a  
98 discrete choice experiment (DCE) among a representative sample of the population in  
99 the United Kingdom (UK) to investigate, for four different types of infectious diseases,  
100 the role played by different age groups in a program's overall evaluation and the extent  
101 to which it matters whether these age groups are affected by either direct, herd or side  
102 effects. We summarize these findings into a set of social preference weights for health  
103 outcomes (e.g. QALYs) that could be incorporated into economic evaluation or public  
104 health impact models.

105

## 106 **2. Methods**

107 DCEs are a widely used survey method to quantify individuals' preferences (Louviere,  
108 Hensher, & Swait, 2000; Ryan, Gerard, & M, 2008) (for a general review of  
109 applications, see (de Bekker-Grob, Ryan, & Gerard, 2012)). Participants are presented  
110 with a series of choices, usually between two goods described by the same attributes  
111 but differing in their attribute levels. By observing respondents' preferred choices,  
112 researchers can infer how the value of the competing options is determined by the  
113 attributes of the product. In our case, we observe how people prioritize between  
114 vaccination programs based on the number of direct, herd and side effects generated  
115 by the program, and their distribution over different age groups. This allows us to  
116 estimate a utility function that describes how the public values vaccination programs,  
117 taking into account the different types of vaccine effect and their distribution.

118

### 119 **2.1 Choice context**

120 For all of their choices, respondents were randomly assigned one of four disease  
121 scenarios (see **Appendix A**). [insert link to appendix] These were introduced before  
122 the start of the DCE. After five choice sets this disease was presented again to the  
123 respondent as a reminder. The four disease profiles were described as (1) severe—  
124 lasting nine days, (2) mild—lasting nine days, (3) severe—lasting 160 days, and (4)  
125 mild—lasting 160 days. Influenza and pertussis were used as proxies for an acute  
126 severe and a longer lasting milder disease, respectively (van Hoek et al., 2014; van  
127 Hoek, Underwood, Jit, Miller, & Edmunds, 2011). To avoid participants' preconceived  
128 ideas, the diseases were unnamed and only described to participants by means of

129 severity using the generic descriptors of the dimensions of a standard instrument to  
130 measure health-related quality of life, the EuroQoL EQ-5D-3L, based on average  
131 reported values for both influenza and pertussis (van Hoek et al., 2014; van Hoek et  
132 al., 2011). To exclude considerations about age differences in remaining life  
133 expectancy, we explicitly told the participants that the diseases were not fatal.

134 Before every choice set we told respondents the following: *“the government has to*  
135 *choose between two vaccination programs that will each be used in 100 000 people.*  
136 *Considering your conviction about vaccination policy, which program do you think the*  
137 *government should choose? Both options are equally costly, and identical in every*  
138 *way except for the following 5 differences.”*

139

## 140 **2.2 Attributes and levels of vaccination programs**

141 To develop the final attributes and levels of the vaccine programs included in the DCE,  
142 we followed a three stage iterative process. We performed a literature search of other  
143 vaccine-related DCEs to assess the choice context and which attributes were typically  
144 considered. These attributes were disease incidence, case fatality risk, economic  
145 impact, duration of illness and duration of vaccine protection, severity of illness and  
146 severity of side effects, and various personal characteristics including age, gender and  
147 willingness/ability to get vaccinated. (de Bekker-Grob et al., 2010; Hofman et al., 2014;  
148 Lambooij et al., 2015; Sadique, Devlin, Edmunds, & Parkin, 2013; Veldwijk, Lambooij,  
149 Bruijning-Verhagen, Smit, & de Wit, 2014) From this list, we took the attributes that  
150 were, in combination with the four disease profiles, best suited to answer our research  
151 question. We presented several attribute combinations to a convenience sample of lay  
152 persons, colleagues and collaborators at the market research company in a pilot

153 questionnaire, which we revised in response to received comments. We re-iterated  
154 this process until we found the right form for the DCE from which, with a relatively  
155 simple set of in total five core attributes (**Table 1**), we could robustly calculate  
156 preference weights.

157 The first two attributes described the age group targeted for vaccination and  
158 magnitude of the direct effects among those vaccinated. The third attribute described  
159 the number of side effects occurring among those vaccinated. The side effects of  
160 vaccination were presented in the DCE as identical to an episode of the disease that  
161 the vaccine usually prevents, in order to enable a direct comparison between the three  
162 effect types. Not doing so would have meant using a second health profile within one  
163 choice option (one for the disease and one for the side effects) and this would also  
164 have made the experiment substantially more difficult for the participants. The fourth  
165 and fifth attribute described the magnitude of the herd effects and the age group that  
166 received them. We decided to focus only on the morbidity aspects of illness because  
167 including mortality would require additional attributes for infected people in order to  
168 account for their differing life expectancy.

169 For direct and herd protection we used 1000, 3000 or 5000 disease episodes  
170 prevented per 100,000 people vaccinated (an attack rate of 1-5% for a vaccine with a  
171 100% efficacy), and for side effects 100, 300 or 500 disease episodes per 100,000  
172 people vaccinated (an attack rate of 0.1-0.5%). For direct protection and side effects,  
173 we considered the following three age groups: children aged between 3 months and  
174 3 years of age, adults aged between 30 and 50 years, and elderly aged between 65  
175 and 75 years. The age groups for herd protection represented groups that, in the case  
176 of the first two, are often difficult to vaccinate for immunological reasons: young

177 children under 3 months, elderly above 80 years and unvaccinated adults between 30  
178 and 50 years.

179

180 (insert **Table 1**)

181

182 We depicted both the age group and quantity of cases avoided or caused by  
183 vaccination using simple graphics (Ancker, Senathirajah, Kukafka, & Starren, 2006)  
184 (**Figure 1**). To explicitly investigate the assumption whether individuals ultimately look  
185 at the total impact of the program and to reduce the chance that respondents would  
186 adhere to a simple counting heuristic without reflection, we presented the net number  
187 of disease cases averted for each strategy separately (the sum of direct and herd  
188 effects minus side effects).

189

190 (insert **Figure 1**)

191

### 192 **2.3 Experimental design of the choice sets**

193 The design of a DCE refers to the number and composition of choice sets presented  
194 to each participant (Reed Johnson et al., 2013). A set of 45 choice sets was selected  
195 out of the 58,806 possible choice sets (see **Appendix B** for more info on the selection  
196 process [insert link to appendix]) and distributed over three survey versions, so to limit  
197 the number of choice sets to be completed per respondent to 15. Therefore, each of  
198 the four disease profiles was represented in three different surveys (see **Figure 2**).

199

200 **(Insert Figure 2)**

201

202 The choice alternatives (i.e. profiles) themselves were '*partial* profiles' (Kessels,  
203 Jones, & Goos, 2015). We varied and highlighted the levels of two to four of the five  
204 attributes in the choice sets and kept the remaining attribute(s) constant so that  
205 respondents did not have to simultaneously trade-off all five dimensions per choice  
206 (see **Appendix B [insert link to appendix]**). Limiting the cognitive burden for  
207 respondents in a DCE increases the validity and reliability of their answers (Dellaert,  
208 Donkers, & van Soest, 2012). The design we generated was 'D-optimal' in a Bayesian  
209 framework fitting with a multinomial logit (MNL) model for the attributes' main effects  
210 and six interactions between the two age attributes (direct and herd effects) and the  
211 three magnitude attributes we deemed to be important *a priori*. We chose a Bayesian  
212 framework to integrate prior information on the respondents' likely preferences  
213 (Kessels, Jones, Goos, & Vandebroek, 2011) (see **Appendix C [insert link to**  
214 **appendix]**). The Bayesian D-optimal design then results in the smallest possible  
215 standard errors for the utility estimates at the given sample size.

216

## 217 **2.4 Sample**

218 After the design, we tested our survey among a pilot sample of the online panel (N=69)  
219 to confirm that respondents could fully understand and complete the survey. Based on  
220 the feedback from this pilot sample we judged that the experiment was understandable  
221 and that no further changes were needed.

222 From a consumer panel of 1 million UK members, 9613 random panelists were  
223 approached to participate in “a scientific study on resource allocation in healthcare”.  
224 Of these people, 4144 (43%) responded to the invitation. We recruited 1950 of them  
225 to fulfill predetermined quotas to provide a representative sample of the UK population  
226 in terms of gender, socio-economic strata (indicated by the occupation of the head of  
227 the household), age groups (20-29, 30-39, 40-49, 50-59, 60+ years), and urban vs.  
228 rural background.

229 The DCE was conducted in November 2016. An email containing a link to the survey  
230 website was sent to participants and by clicking on the link respondents consented to  
231 participate, although they were free to stop or close the survey at any point. All  
232 respondents received a nominal incentive for study completion (£0.50 per 12-minute  
233 questionnaire). Before completing the DCE, respondents were asked to administer a  
234 survey tool to measure vaccine hesitancy (Larson et al., 2015), and were asked social-  
235 demographic questions and whether they have or had children. After the DCE, we  
236 asked about their experience with severe diseases, their interpretation of the validity  
237 of the answers they provided and the overall difficulty of the DCE survey.

238 We obtained informed consent from all respondents and ethical approval of the study  
239 from the Ethics Committee of the London School of Hygiene & Tropical Medicine (Ref  
240 10335). We conducted the research in accordance with the Code of Conduct of the  
241 Market Research Society, which ensured that information is collected for research  
242 purposes only, is kept confidential, and respondent anonymity is guaranteed.

243

## 244 **2.5 Data analysis**

245 To quantify the weight of the five attributes and their levels in the utility attributed to a  
246 vaccination strategy, a panel mixed logit model (fitted by the Hierarchical Bayes  
247 method (Train, 2009)) was used (see **Table 3**). The model involved seven main  
248 effects: four related to the two three-level categorical attributes describing the utility  
249 impact of a change in the targeted age group in direct and herd effects, and three  
250 related to the continuous attributes describing the impact of a change in the absolute  
251 number of disease cases via direct effects, side effects and herd effects. Besides  
252 these seven main effects the model also includes attribute interaction effects,  
253 indicating the additional change in utility because of a particular combination of  
254 attribute levels. We computed the overall significance of the attributes using likelihood  
255 ratio (LR) tests and measured the relative importance of the attributes by the logworth  
256 statistic (i.e.  $-\log_{10}$  (p-value of the LR-test)). The coefficients of the logit model were  
257 obtained by estimating the *a priori* model, i.e. the model with the utility function that  
258 seemed most appropriate when planning the DCE, and subsequently dropping the  
259 non-significant model terms until we obtained a *final* model in which all effects had  
260 significant explanatory value at the 5% level. Models were fitted using the JMP 13 Pro  
261 Choice platform (based on 10,000 iterations, with the last 5000 used for estimation)  
262 assuming normally distributed parameters with no correlation between the attributes.  
263 Combining the main and interaction effects, this model allows calculating the additional  
264 utility of a vaccination program generated per additional health effect, i.e. per type of  
265 effect per age group (see the nine variations in **Table 3**). The 95% confidence intervals  
266 for the equity weights were estimated using the Delta method (Bliemer & Rose, 2013).

267

268 We investigated heterogeneity in respondents' preferences in two ways. First, by  
269 exploring the influence of the observed respondent characteristics on the average

270 preferences and, second, by studying the unobserved preference heterogeneity by  
271 means of a hierarchical cluster analysis on the subject-specific estimates resulting  
272 from the Hierarchical Bayes approach. We favoured this two-stage modelling method  
273 as it performs equally well as one-stage modelling methods such as latent class  
274 modelling (Crabbe, Jones, & Vandebroek, 2013; Kessels, Jones, & Goos) while  
275 enabling us to parsimoniously derive the preference weights and their 95% confidence  
276 intervals.

277

## 278 **3. Results**

279

### 280 **3.1 Response**

281 A total of 1546 respondents out of 1950 (79%) who were sent the questionnaire  
282 completed it, of which 47 (3%) indicated that the questions were too difficult or their  
283 answers invalid, leaving 1499 questionnaires for analysis. Our final sample was  
284 sufficiently representative of the UK population in terms of gender, family size, socio-  
285 economic status and education level (**Table 2**).

286

287 (insert **Table 2**)

288

### 289 **3.2 Main effects and calculated weights**

290 Across all questionnaires, respondents made a total of 22,485 choices between  
291 vaccination programs. There was no significant effect observed of which of the three

292 survey versions a participant received. Respondents did not systematically choose the  
293 program with the highest overall public health impact, i.e. the total of all prevented  
294 cases including direct, herd and side effects. In fact, only 99 respondents (6.6%)  
295 consistently opted for the most effective program in all of their choice sets. However,  
296 about half the respondents (738/1499) chose the most effective alternative in at least  
297 70% of their choices, indicating that the total effect on the disease burden is important,  
298 but not the only factor in prioritizing vaccination programs.

299 **Table 3** presents an overview of the incremental utility of the main effects and  
300 interactions. The vaccination program that was least preferred (i.e. yielding minimum  
301 utility) was one that targeted the elderly (65-75y), generated the lowest number of  
302 prevented cases, the highest number of side effects, and the lowest number of cases  
303 prevented via herd protection in unvaccinated adults. The most preferred program (i.e.  
304 yielding maximum utility) was one that targeted children, generated the highest  
305 number of prevented cases, the lowest number of side effects, and the highest number  
306 of cases prevented via herd protection in newborns.

307

308 (insert **Table 3**)

309

310 Using the same logit model, we then calculated preference weights for each effect  
311 type per age group. These weights act as a multiplicative factor to transform identical  
312 clinical symptoms into health effects with equal value in the public's view. We  
313 compared the additional utility of a vaccination program that is generated through  
314 preventing one specific disease case relative to the utility gained through directly  
315 preventing a single disease case via vaccinating a child (**Figure 3**). These preference

316 weights reveal important patterns. First, preventing side effects of vaccination was  
317 highly preferable to preventing natural infections, even though the symptoms were  
318 equal in length and severity. The mean weight for side effects across all ages was -  
319 2.93, meaning that avoiding one vaccine-induced infection was weighted equally to  
320 avoiding around three natural infections among children. This finding was consistent  
321 whether side effects occurred in children (-2.95 (95% CI: -3.21; -2.69)), adults (-3.16  
322 (95% CI: -3.51; -2.81)) or the elderly (-2.68 (95% CI: -2.98; -2.37)). Second,  
323 respondents preferred vaccination programs that prevented disease among newborns  
324 and children compared with those for adults and the elderly, even though the  
325 prevented disease burden was similar. One episode prevented in a newborn via herd  
326 protection was considered about twice as valuable as directly protecting an adult via  
327 vaccination. Third, the extent to which respondents preferred protecting adults and the  
328 elderly depends on the type of benefit conferred by the program. Direct effects were  
329 the preferred mode of protection for adults whereas herd effects were preferred for the  
330 elderly. Reducing disease burden by directly vaccinating adults (aged 30-50 years)  
331 was weighted equally to reducing disease burden in the elderly (aged 80+ years) via  
332 herd effects [0.75 (0.64; 0.85) compared to 0.67 (0.58; 0.76), respectively]. In contrast,  
333 reducing disease burden in adults (aged 30-50 years) by herd effects counted equally  
334 to reducing disease burden in elderly (aged 65-75 years) directly via vaccination (0.12  
335 (0.03; 0.20) compared to 0.16 (0.06; 0.25), respectively).

336

337 (insert **Figure 3**)

338

339 From these results, we also calculated the number of infections needed to avert in  
340 order to obtain equal utility as that from protecting 100 children directly via vaccination  
341 (**Table 4**). Avoiding 100 infections in children via vaccination was considered  
342 equivalent to protecting 632 elderly (65-75 years) or 134 adults. In turn, these  
343 outcomes were equivalent to protecting 71 newborns, 865 adults or 150 elderly (>80y)  
344 via herd protection. Similarly, a vaccination strategy reduces its utility by causing side  
345 effects. Avoiding 34 side effects in children generates the same utility as preventing  
346 100 natural infections among the same age group.

347

348 (insert **Table 4**)

349

350 **Figure 4** illustrates the significant interaction in our model between the age of the  
351 vaccinated group and the age of the herd protection recipients (see **Table 3**). This  
352 interaction must be understood as the additional utility that is given to (or taken away  
353 from) a vaccination program depending on the particular combination of age groups  
354 that are involved, regardless of the magnitude of direct, herd or side effects that are  
355 being generated. It presents the attractiveness of particular intergenerational  
356 vaccination strategies. Whereas a CEA perspective would consider all possible age  
357 combinations equally attractive (as long as they lead to the same number of infections  
358 prevented), our sample had clear intergenerational preferences over vaccination  
359 strategies. Any age group was deemed acceptable to vaccinate when there were herd  
360 protection benefits for newborns. To generate herd protection for adults, children were  
361 the most attractive age group. To generate it to protect the elderly >80, adults were  
362 deemed most appropriate. The least attractive intergenerational combination was

363 vaccinating elderly 65-75 years while generating herd protection in adults 30-50 years.  
364 The most attractive age combination was vaccinating children while generating herd  
365 protection in newborns.

366

367 (insert **Figure 4**)

368

### 369 **3.3 Preferences across disease types and respondents**

370 As shown in **Appendix D** ([insert link to appendix], our results remained robust  
371 across all four different disease types: the equity weights were statistically equivalent,  
372 regardless of whether the condition was mild vs. severe or acute vs. chronic (indicated  
373 by a non-significant interaction effect in our model between the attributes and the  
374 disease type). Also, the appendix [insert link to appendix] illustrates that our findings  
375 also remained robust across most respondent characteristics: gender, age,  
376 occupation, level of education, urban-rural, socio-economic background, experience  
377 with severe illness or parental status. Although individuals with a low degree of vaccine  
378 hesitancy (indicated by high values on the 'vaccine hesitancy scale' (VHS) (Larson et  
379 al., 2015)) attributed less importance to side effects ( $p < 0.0001$ ), this effect was  
380 relatively small (a 10 unit increase in the VHS score (on a scale from 10 to 50) led to  
381 a 10% decrease in absolute magnitude of the utility for side effects ( $\sim 0.03$ )).

382 The hierarchical cluster analysis of the individual preferences (see methods) revealed  
383 two distinct groups of respondents: one group ( $N=564$ , *Cluster 1*) who attached almost  
384 no importance to the number of side effects (with a mean weight of  $-0.91$  for side  
385 effects) and a larger group ( $N=935$ , *Cluster 2*) who valued this attribute fairly highly  
386 (with a mean weight of  $-4.40$ ) (**Table 3**). This clustering explains the relatively high

387 variation across respondents for the weight estimate for side effects (the standard  
388 deviation to mean absolute value ratio of 0.043 for side effects is almost twice the ratio  
389 for direct and herd effects). We used a logistic regression to determine predictors of  
390 cluster membership. Cluster 1, who attached almost no importance to the number of  
391 side effects, was characterized by high values on the VHS, indicating little hesitancy  
392 ( $p < 0.0001$ ). On the other hand, cluster 2, who valued side effects more highly, was  
393 characterized by higher degrees of hesitancy on the VHS. However, the predictive  
394 power of this association for membership of the group was small (McFadden's pseudo  
395  $R^2 = 0.6\%$ ), implying that there is much unexplained heterogeneity in the importance  
396 placed on side effects.

397

398

## 399 **4. Discussion**

400 In this study, we used a discrete choice experiment to analyse and quantify how the  
401 public values the outcomes of vaccination programs. We observed several general  
402 preference patterns, which were robust across different lengths and severities of  
403 disease and respondent characteristics (socio-economic background, age, education  
404 and parenthood). We observed that most respondents did not make choices purely  
405 based on how to minimize the number of infections. In particular, individuals, on  
406 average, weighted one averted instance of a side effect equal to about three similarly  
407 severe natural infections in children and weighted one averted health outcome in  
408 children up to six times more than preventing similarly severe health outcomes in the  
409 elderly. Interestingly, our study has disentangled this latter phenomenon from the type  
410 of effect as we observed a different weight given to protecting older people depending  
411 on whether the benefits were directly vs. indirectly received. Our results support a duty  
412 of care principle to provide herd protection for the elderly and an aversion to protecting  
413 adults who are better able to protect themselves. The weight given to side effects when  
414 evaluating a vaccination program was divisive, splitting our sample into two clusters.

415 Our study, as far as we are aware, is the first of its kind to quantify the important social  
416 value judgements that need to be made in vaccine funding decisions. Although this  
417 limits comparability, our findings are in line with what can be learned from other study  
418 domains. The finding that individuals weighted one averted instance of a side effect  
419 equal to about three similarly severe natural infections in children can be explained  
420 with general theory on decision-making. For instance, well-documented psychological  
421 phenomena such as ‘loss aversion’ (Kahneman & Tversky, 1979) (overvaluing risks  
422 and losses over opportunities and gains), the ‘act-omission bias’ (Spranca, Minsk, &  
423 Baron, 1991) [judging the effects of an act (becoming vaccinated) differently from

424 identical effects resulting from an omission (becoming infected)], or ‘hyperbolic  
425 discounting’ (Frederick, Loewenstein, & O’Donoghue, 2002) [overvaluing the present  
426 (in which side effects occur) over the future (in which disease prevention will occur)]  
427 suggest that people put an extraordinary weight on side effects when evaluating a  
428 vaccination strategy. Moreover, also empirical studies that have investigated people’s  
429 (stated) choices about whether or not they would personally become vaccinated with  
430 a particular vaccine (e.g. (Sadique et al., 2013; Seanehia et al., 2017)) generated  
431 findings that highlight the extraordinary weight of side effects. The preference given to  
432 health benefits in younger people (newborns and children), up to six-fold, is also in line  
433 with related studies on ‘ageism’ in other contexts of healthcare priority-setting  
434 (reviewed in (Gu, Lancsar, Ghijben, Butler, & Donaldson, 2015) and discussed  
435 elsewhere, e.g. (Bognar, 2015; Tsuchiya, 2000)).

436 It is important to study which aspects of health policy choices matter most to the public.  
437 This is especially true in vaccination where public trust, goodwill and participation are  
438 sensitive and key to success (Cooper, Larson, & Katz, 2008). There is a growing  
439 concern that public and political trust in scientific evidence is eroding, particularly in  
440 the context of vaccination (Karafillakis et al., 2016; Larson, Cooper, Eskola, Katz, &  
441 Ratzan, 2011; Leask, Willaby, & Kaufman, 2014). By being aware of the sensitivities  
442 around vaccination, decision makers can understand and address some of the root  
443 causes of vaccine hesitancy, adapt to concerns of the population and improve  
444 responses in communication strategies.(Diekema & American Academy of Pediatrics  
445 Committee on, 2005) Our findings provide empirical evidence on how to set vaccine  
446 priorities in line with public preferences. There is an important debate over the extent  
447 to which the public’s opinion should drive resource allocation in healthcare (see e.g.  
448 (Hausman, 2004, 2015)). But, many believe that the values of the public, who pays for

449 healthcare, should at least somehow be acknowledged in the decision-making  
450 process. In the context of vaccination, where public support and participation is key to  
451 success, this concern becomes particularly crucial. Therefore, our results can be  
452 useful additions to vaccine appraisals. They can provide guidance in specific  
453 epidemiological cases where CEA does not provide the answers needed. For  
454 instance, our results would suggest that, despite their attractiveness in terms of cost-  
455 effectiveness, the public may not support a childhood influenza vaccination program  
456 that mainly benefits adults or elderly (Baguelin et al., 2013), because preventing side  
457 effects in vaccinated children is preferred over preventing disease burden among  
458 adults and elderly. Furthermore, our study suggests that a childhood varicella-zoster  
459 vaccination program, in the case that it protects children against varicella disease at  
460 the expense of increased zoster in the elderly (the 'exogenous boosting hypothesis'),  
461 might be justifiable. In contrast, previous analyses where QALY losses for children are  
462 weighted equally to those for the elderly find that the increased burden in the elderly  
463 offsets the QALY gains in children and determine the program not cost-effective  
464 (Brisson, Edmunds, & Gay, 2003; Luyten et al., 2014).

465 Our results can also be directly incorporated into economic evaluations as sensitivity  
466 analyses to better align the underlying assumptions of CEA with the values of the  
467 population. Our estimated preference weights can be used in decision-analytic models  
468 as a parameter to weight QALYs or infections according to their 'social value'. This  
469 would re-adjust the (equal) weight that QALYs receive in CEA according to how  
470 important people think that the age of the QALY-recipient is and whether the benefit  
471 was generated through direct protection, herd immunity or (avoiding) side effects.  
472 There is an increased interest in such 'extended', 'distributive' or 'equity-weighted'  
473 economic evaluation (see e.g. (Asaria, Griffin, & Cookson, 2016; Bleichrodt, 1997;

474 Cookson et al., 2009; Dolan, 1998; Fleurbaey et al., 2013; Nord, Pinto, Richardson,  
475 Menzel, & Ubel, 1999; Round & Paulden, 2017; Samson et al., 2017)), but, to our  
476 knowledge, such studies do not exist for the evaluation of vaccines. Our estimates are  
477 developed particularly for this context, and provide an opportunity to do so.

478 There are several limitations. We did not include any mortality effects, nor did we  
479 include a difference in severity between the three vaccine effects, even though this  
480 would be more realistic (as side effects of vaccines are usually milder than the disease  
481 being prevented). We chose not to include these aspects because we wanted to avoid  
482 increasing the complexity of the survey and reducing the validity of the respondents'  
483 answers by adding a second disease profile. Also, keeping the disease outcome  
484 constant over age groups and effects enabled trade-offs that were wholly reflective of  
485 the preference between age groups and effects instead of also reflecting additional  
486 considerations about disease severity. We also chose to present the number of side  
487 effects rather than its complement the number of vaccinated people *without* side  
488 effects. This framing may have played a role in the observed weight for side effects.  
489 The alternative framing would probably have drawn less attention to side effects and  
490 might have generated smaller weights. We however wanted people to make explicit  
491 trade-offs between side effects with protective benefits and chose for the more direct  
492 framing. Using the alternative is a suggestion for further research. Also, we used  
493 generic disease profiles based on a description in EQ-5D terms to minimize  
494 respondents making personal associations to the disease and vaccine when we would  
495 have named the diseases (e.g. 'flu' or 'whooping cough'), but this may also have  
496 increased the level of abstraction and reduced the level of personal involvement. A  
497 suggestion for further research is to repeat our study with named diseases and to test  
498 whether our finding that the disease profile did not matter to people's preferences is

499 confirmed. Another limitation is that, while our sample was broadly representative of  
500 the UK population, it was recruited from an online panel where membership may be  
501 associated with unobserved characteristics (e.g. interest in technology).

502 In conclusion, our study demonstrates clear and robust preference patterns in how  
503 people value the impact of vaccination programs. A large majority of respondents had  
504 a strong preference to minimize side effects and to prevent disease among newborns  
505 and children. Our observations provide quantitative evidence about public preferences  
506 around important and sensitive but neglected trade-offs in vaccine policy decision-  
507 making, and can hopefully inspire further research and discussion.

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698 **Table 1. Attributes and levels used in the DCE**

Attribute	Level
Age of vaccinated group (N=100 000)	Children (3 months - 3 years)
	Adults (30-50 years)
	Elderly (65-75 years)
Disease episodes prevented in vaccinated group	1000 cases
	3000 cases
	5000 cases
Number of vaccine-induced side-effects	100 cases
	300 cases
	500 cases
Disease episodes prevented via herd protection	1000 cases
	3000 cases
	5000 cases
Age of people receiving herd protection	Newborns (<3 months)
	Adults (30-50 years)
	Elderly (>80 years)

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702 **Table 2: Respondent characteristics.**

	Sample	UK population*
Total recruited	1546	
Excluded for analysis	47	
Included in the analysis	1499 (100%)	
<i>Gender</i>		
Male	703 (47%)	49%
Female	796 (53%)	51%
<i>Age (years)</i>		
20-29	296 (20%)	13%
30-39	285 (19%)	13%
40-49	288 (19%)	14%
50-59	308 (21%)	13%
60 and over	322 (21%)	23%
<i>Living in a city with more than 10,000 inhabitants</i>	1011 (67%)	83%
<i>Social grades based on the profession of the highest paid household member</i>		
A (upper middle class)	85 (6%)	4%
B (middle class)	297 (20%)	23%
C1 (lower middle class)	385 (26%)	27%
C2 (skilled working class)	330 (22%)	21%

D (working class)	72 (5%)	16%
E (non-working)	330 (22%)	9%
<i>Education level</i>		
No qualifications	48 (3%)	15%
Secondary education	322 (21%)	14.2%
Post-secondary education	288 (19%)	14.5%
Vocational qualification	254 (17%)	20.3%
Undergraduate degree, Post-graduate degree & Doctorate	427 (39%)	30%
Not sure	2 (0.1%)	/
<i>Having children</i>		
No children	585 (39%)	42%
Children aged 0-4 years	168 (11%)	42%**
Children aged 5-20 years	358 (24%)	/
Children aged over 20 years	388 (26%)	15%
<i>Exposure to poor health</i>		
Participant affected by poor health	407 (27%)	
Close friends or family of the participant affected by poor health	470 (31%)	
Neither participant nor close friends nor family affected by poor health	622 (41%)	

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704 \*UK population data 2016: Office for National Statistics <https://www.gov.uk/government/publications>

705 \*\*Percentage of UK families living with dependent children (<18 years old)

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721 **Table 3. Attributes that affected respondent choices, based on panel mixed logit model estimates (means and standard**  
 722 **deviations) with p-values from likelihood ratio (LR) tests for significant attribute effects.**

<b>Model term</b>	<b>Posterior mean</b>	<b>Posterior std dev</b>	<b>Subject std dev</b>	<b>P-value</b>	
Cases prevented in unvaccinated by herd effects (per 1000 cases)	0.715	0.018	0.101	<0.0001	
Cases prevented in vaccinated by direct effects (per 1000 cases)	0.619	0.018	0.100	<0.0001	
Cases of side effects in vaccinated (per 100 cases)	-0.285	0.012	0.110	<0.0001	
Age of unvaccinated	[Newborns <3m]	0.614	0.048	0.090	<0.0001
	[Adults 30-50y]	-0.597	0.043	0.105	
	[Elderly >80y]	-0.017	NA	NA	
Age of unvaccinated*Cases prevented in vaccinated by direct effects	[Newborns <3m]	-0.043	0.009	0.054	<0.0001
	[Adults 30-50y]	0.071	0.009	0.041	
	[Elderly >80y]	-0.028	NA	NA	
Age of vaccinated	[Children 3m-3y]	0.305	0.040	0.063	<0.0001
	[Adults 30-50y]	0.142	0.048	0.062	

	<i>[Elderly 65-75y]</i>	-0.446	NA	NA	
Age of unvaccinated*Age of vaccinated	[Newborns <3m]* [Children 3m-3y]	-0.131	0.036	0.053	<0.0001
	[Newborns <3m]* [Adults 30-50y]	-0.210	0.041	0.065	
	<i>[Newborns &lt;3m]* [Elderly 65-75y]</i>	0.341	NA	NA	
	[Adults 30-50y]* [Children 3m-3y]	0.250	0.052	0.044	
	[Adults 30-50y]* [Adults 30-50y]	-0.079	0.049	0.045	
	<i>[Adults 30-50y]* [Elderly 65-75y]</i>	-0.171	NA	NA	
	<i>[Elderly &gt;80y]* [Children 3m-3y]</i>	-0.119	NA	NA	
	<i>[Elderly &gt;80y]* [Adults 30-50y]</i>	0.289	NA	NA	
	<i>[Elderly &gt;80y]* [Elderly 65-75y]</i>	-0.170	NA	NA	
Age of vaccinated*Cases of side effects in vaccinated	[Children 3m-3y]	-0.032	0.008	0.040	<0.0001
	[Adults 30-50y]	-0.037	0.009	0.044	
	<i>[Elderly 65-75y]</i>	0.069	NA	NA	
	[Newborns <3m]	0.052	0.009	0.048	<0.0001

Age of unvaccinated*Cases prevented in unvaccinated by herd effects	[Adults 30-50y]  <i>[Elderly &gt;80y]</i>	-0.005  <i>-0.047</i>	0.008  NA	0.043  NA	
Age of vaccinated*Cases prevented in vaccinated by direct effects	[Children 3m-3y] [Adults 30-50y] <i>[Elderly 65-75y]</i>	0.051 -0.032 <i>-0.019</i>	0.010 0.009 NA	0.044 0.037 NA	<0.0001

723 Note: Mean estimates corresponding to the last level of an attribute, either as a main effect or involved in an interaction, are italicized and calculated as minus  
724 the sum of the estimates for the other levels of that attribute; NA means 'not assigned'.

725 **Table 4. Number of infections to prevent to gain equal utility, with 95%**  
 726 **confidence intervals.**

<b>Age group of vaccine effect</b>	<b>Direct effects</b>	<b>Herd effects</b>	<b>Side effects</b>
Newborns (<3 months)	NA	<b>71</b> [66; 76]	NA
Children (3 months – 3 years)	<b>100</b> [index]	NA	<b>-34</b> [-37; -31]  Cluster 1: -221 [-340; -102] Cluster 2: -21 [-23; -20]
Adults (30–50 years)	<b>134</b> [115; 153]	<b>865</b> [242; 1487]	<b>-32</b> [-35; -28]  Cluster 1: -72 [-93; -51] Cluster 2: -23 [-25; -20]
Elderly (65–75 years)	<b>632</b> [255; 1010]	NA	<b>-37</b> [-42; -33]  Cluster 1: -113 [-163; -64] Cluster 2: -25 [-27; -22]
Elderly (>80 years)	NA	<b>150</b> [130; 169]	NA

727 Note: Cluster 1 and 2 have 564 and 935 respondents, respectively; NA refers to combinations of  
 728 attribute levels not included in the choice profiles.

729

730 **Figure 1. Example of a choice set.**

731

732 **Figure 2. Schematic representation of the different arms of the questionnaire.**

733 **For each disease stratum, there was also an equal sampling over the socio-**  
734 **economic groups (25% A+B; 25% C1; 25% C2; 25% E+D).**

735

736 **Figure 3. Utility weights representing public preferences for identical health**  
737 **outcomes with different attributes, with 95% confidence intervals.**

738

739 **Figure 4. Intergenerational preferences: interaction effects between the age**  
740 **group vaccinated and the age group receiving herd protection effects.**

741 **Marginal utility values consist of main effects of the attributes involved and**  
742 **their interaction effect.**

743

744