Psychometric evaluation of the Major Depression Inventory among young people living in Coastal Kenya [version 1; referees: 2 approved]

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Abstract

Background: The lack of reliable, valid and adequately standardized measures of mental illnesses in sub-Saharan Africa is a key challenge for epidemiological studies on mental health. We evaluated the psychometric properties and feasibility of using a computerized version of the Major Depression Inventory (MDI) in an epidemiological study in rural Kenya.

Methods: We surveyed 1496 participants aged 13-24 years in Kilifi County, on the Kenyan coast. The MDI was administered using a computer-assisted system, available in three languages. Internal consistency was evaluated using both Cronbach’s alpha and the Omega Coefficient. Confirmatory factor analysis was performed to evaluate the factorial structure of the MDI.

Results: Internal consistency using both Cronbach’s Alpha (α= 0.83) and the Omega Coefficient (0.82; 95% confidence interval 0.81 - 0.83) was above acceptable thresholds. Confirmatory factor analysis indicated a good fit of the data to a unidimensional model of MDI ($\chi^2 (33, N = 1409) = 178.52, p < 0.001$, TLI = 0.947, CFI = 0.961, and Root Mean Square Error of Approximation, RMSEA = 0.056), and this was confirmed using Item Response Models (Loevinger’s H coefficient 0.38) that proved the MDI was a unidimensional scale. Equivalence evaluation indicated invariance across sex and age groups. In our population, 3.6% of the youth presented with scores suggesting major depression using the ICD-10 scoring algorithm, and 8.7% presented with total scores indicating presence of depression (mild, moderate or severe). Females and older youth were at the highest risk of depression.

Conclusions: The MDI has good psychometric properties. Given its brevity, relative ease of usage and ability to identify at-risk youth, it may be useful for epidemiological studies of depression in Africa. Studies to establish clinical thresholds for depression are recommended. The high prevalence of depressive symptoms suggests that depression may be an important public health problem in this population group.
Keywords
Major Depression Inventory, youth, Kenya, psychometric evaluation

This article is included in the KEMRI | Wellcome Trust gateway.

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Competing interests: No competing interests were disclosed.

How to cite this article: Otiende M, Abubakar A, Mochamah G et al. Psychometric evaluation of the Major Depression Inventory among young people living in Coastal Kenya [version 1; referees: 2 approved] Wellcome Open Research 2017, 2:113 (doi: 10.12688/wellcomeopenres.12620.1)

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Grant information: This work was supported by the Wellcome Trust [084535, 091758]; and the Hewlett Foundation through a grant awarded to INDEPTH [IND/0107].

The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

First published: 29 Nov 2017, 2:113 (doi: 10.12688/wellcomeopenres.12620.1)
**Introduction**

Young people aged 10–24 years constitute about a third of the global population, and behaviour and mental health at this age have substantial health effects later in life (Gore et al., 2011). For instance, half of all adult mental health disorders first appear during the adolescent years (Gore et al., 2011), though many are undiagnosed and untreated prior to adulthood. Mental health disorders are common and are a leading cause of disability in young people globally (Gore et al., 2011; World Health Organization, 2014). In Africa Disability-Adjusted-Life-Years (DALYs) lost due to depressive disorders is higher than in other regions and neuropsychiatric disorders are estimated to contribute 5% of the total DALYs lost (Vos et al., 2016; World Health Organization, 2008). Studies in Kenya have suggested that roughly one-quarter of out-patients attending health facilities have mental disorders (Kiima et al., 2004). Despite this, prevention, screening, and treatment of mental health problems and neuropsychiatric disorders have been widely overlooked in public health programmes in low and middle-income countries (LMICs).

The lack of adequately standardized and psychometrically robust measures for evaluating mental health problems (Mutumba et al., 2014; Patel et al., 2007; Saxena et al., 2007; Tennison et al., 2016) has contributed to the shortage in high quality epidemiological studies to estimate the burden of mental health problems in Africa. Only recently have there been efforts, to develop, adapt and evaluate measures of mental health (Betancourt et al., 2009; Mels et al., 2010) for use in Africa. However, there are still very few reports of the psychometric properties of the widely used measures of mental health (Mutumba et al., 2014).

Several screening tools for depression have been used in previous studies in Kenya, including the Patient Health Questionnaire (PHQ-9), the Center for Epidemiologic Studies-Depression (20-item CES-D) questionnaire, and Beck’s Depression Inventory (21-item BDI) (Kingori et al., 2015; Seth et al., 2014). Although an extensive study in Kilifi found that the BDI was a good measure of depressive symptoms, it is lengthy and therefore not suitable for use in large epidemiological studies (Abubakar et al., 2016). Similarly, the CES-D has 20 items, thus likely to present the same problems as the BDI. PHQ-9 is constructed to screen for major depression, but does not measure depression severity (Bech et al., 2015). We chose to evaluate the psychometric properties of the Major Depression Inventory (MDI) (Bech et al., 2001) when applied to young people in a rural African context, because of its brevity and validity for assessing common depressive disorders (Konstantinidis et al., 2011). Compared to other scales that measure depressive disorders, the MDI was shown to be superior to the Zung-Self-Rating Depression Scale (Zung-SDS) and BDI (Bech et al., 2015; Konstantinidis et al., 2011).

The selection of measures for adaptation and evaluation for use in a new context requires careful consideration of a number of factors. First, the validity of the measure in the original context and in other cultural contexts where it was validated should be considered. The MDI is a measure which has accrued a significant amount of evidence on its construct validity, and has sound psychometric properties both in its original sample (Olsen et al., 2003) and across different cultural contexts such as Greece (Fountoulakis et al., 2003), Brazil (Parcias et al., 2011) and Egypt (Fawzi et al., 2012). Second, it is important to consider the feasibility of using the measures in the new context. The brevity and relative ease of use of the MDI makes it particularly suitable for screening depressive symptoms in large epidemiological studies compared to other scales or questionnaires. Lastly, the cultural appropriateness and relevance of the items for the new context is critical for the validity of measures being moved from one cultural context to the other. Thus, content validity of the items should be ascertained to ensure that items are suitable for use in the new context. In Kilifi an earlier qualitative study of symptoms closely associated with depression identified an overlap between the themes, symptoms and contents of the MDI with community perceptions of depressive symptoms (Abubakar et al., 2016). For instance, some of the key symptoms identified by the members of the community included feelings of sadness, loss of weight, and lack of sleep among others. The information collected from our qualitative work provides a good indication that the items covered by MDI were relevant to rural Kenya. These considerations favored the selection of MDI as the most suitable measure of depressive symptoms to evaluate for use in this context.

In addition, social desirability bias is a common problem inherent in face-to-face interviews that involve sensitive questions. Social desirability is the inclination to report behaviors that are viewed in a favorable way by society (Beauchalt et al., 2013). In order to mitigate against this form of bias and other concerns of privacy when responding to sensitive questions, appropriate interviewing/data collection mechanisms, such as the Audio Computer-Assisted Self-Interviewing (ACASI) system need to be utilized (West & Peyicheva, 2014). In this study, interviews were conducted using the ACASI in three different languages.

When scales and measures are used in another context, it is important to evaluate their psychometric properties, as it cannot be assumed that psychometric properties from the original contexts are retained. An adequate psychometric evaluation process needs to check for reliability, validity and feasibility of the scale.

Consequently, we performed an evaluation of MDI and evaluated the feasibility of using this tool in rural Coastal Kenya. Specifically, we aimed to evaluate:

- a) the internal consistency of the MDI;
- b) if the MDI would present with the single factorial structure as envisioned by the authors;
- c) whether there is equivalence of these structures across sex and age.

**Ethical statement**

The Kenya Medical Research Institute National Scientific and Ethical Committee approved this study (approval number: SSC...
The major depression inventory (MDI) is a self-reported measure of depressive symptoms with items examining the moods one experienced in the previous 2 weeks. The measure covers DSM-IV and has 12 items, though only 10 items are scored, since items 8 (a and b) and items 10 (a and b) are collated into two items. The questions are scored on a 6 point Likert scale with ‘0’ being at no time and at ‘5’ all the time. The individual items measure how much of the times the symptoms have been present during the past 14 days. According to Olsen (2003) the MDI can be used both as:

- A diagnostic instrument with the algorithms leading to the DSM-IV or ICD-10 “major” or “moderate to severe depression”
- A depression rating scale to measure depression severity in which the total score of the 10 items is a sufficient statistic.

In this study, we assessed depressive symptoms using both methods. We categorized the severity of depression following the MDI scores: 20–24 for mild depression; 25–29 for moderate depression; and 30 or higher for severe depression.

The MDI, which originally is developed and evaluated in a Danish population, has been found to have high sensitivity (between 0.86 and 0.92) and specificity (between 0.82 and 0.86) compared to Schedules for Clinical Assessment in Neuropsychiatry (SCAN) as the index of validity (Bech et al., 2001). Likewise, in the original population its internal and external validity was evaluated using both classical methods (i.e. Cronbach’s Alpha) and modern methods (i.e. Mokken scale analysis) and the MDI was found to be a unidimensional scale (Olsen et al., 2003). Consequently, in this study, we use both the classical and modern methods to evaluate the validity of the MDI in our context.

Translation of questionnaire items
All the items (including the MDI items) were translated to Swahili and Giryama. Forward translation of the MDI to English was done by other independent translators. Through a harmonization meeting, it was ensured that all the translated items of the MDI were equivalent to the original English version in terms of meaning and context.

Measure of behavioral patterns among the youth
Of interest to this study were measures of patterns of alcohol use and experiences of bullying which were part of behavioral data collected during the survey. Drinkers were defined as those who consumed any amount of alcohol. Bullying was classified into three categories: those who were never bullied; those who were less bullied (2 or 3 times); and those who were often bullied (4 to 7 times). The period of assessment was within 30 days before the day of the interview.

Data collection procedures
The interviews were done using an Audio Computer-Assisted Self–Interview software (ACASI) to ensure respondent privacy.
and improve the accuracy of the information collected. A previous study demonstrated that participants found the ACASI modality to be acceptable, private and user friendly (Beauclair et al., 2013). All questions were recorded and programmed into the ACASI software in three languages (English, Swahili, and Giryama). Data was collected during household visits where study participants were instructed in detail on how to use the ACASI prior to being given a laptop and a pair of headphones. Participants were left to complete the interview in private.

Prior to the start of the survey, a pilot was done to assess the functionality of the ACASI software. This pilot was covered by ethical approval that was granted for the study. During the pilot, respondents also filled out physical paper forms (alongside ACASI), and these responses were later compared to the responses from ACASI. For each question, response from the paper questionnaire and ACASI version were compared and were considered similar if they were exactly the same.

Statistical analyses
Internal consistency. We computed the Cronbach’s alpha and Omega coefficient with 95% confidence interval to assess the internal consistency of the scale. The omega coefficients were computed using the MBESS package in R (Kelley, 2007). Owing to criticisms, surrounding the use of Cronbach’s alpha in the assessment of reliability of psychological scales (Black et al., 2014; Dunn et al., 2014; Iacobucci & Duhachek, 2003), we also computed the Omega coefficient. Conventionally, both coefficients are acceptable when the value is above 0.70 (Cicchetti, 1994).

Factorial structure. The factorial structure of the MDI was assessed using both classical methods (Exploratory Factor Analysis) and modern methods (Item Response theory). We performed factor analysis using iterated principal factors. A scree-plot and eigenvalues of all possible factors were evaluated to determine the number of factors to retain. Any factor with an eigenvalue greater than 1 was retained. Mokken scale analysis, a class of statistical models called item response models (IRM), was also used to assess whether the MDI was a unidimensional scale. A Loevinger’s H coefficient between 0.30 and 0.39 was considered an acceptable indication of a unidimensional scale (Mokken, 1971; Olsen et al., 2003).

To assess the equivalence of psychological measures across age and sex, we used a multi-group confirmatory factor analyses (MGCFA). A unidimensional model including all 10 items was estimated. We assessed the goodness of fit for each model using Chi-Square, the Tucker-Lewis Index (TLI), and the Comparative Fit Index (CFI). The general guideline for TLI and CFI is that values greater than 0.95 are considered to reflect an excellent fit, while those between 0.95 and 0.90 are considered indicative of an acceptable fit. We report the Root Mean Square of Approximation (RMSEA), which is sensitive to model misspecification: values of less than 0.06 are considered, indicative of a good fit while those between 0.06 and 0.08 are considered indicative of an acceptable model.

In a multi-group analysis, the change in CFI is an important indicator for assessing the suitability of hierarchically nested models. Thus, a CFI change of less than 0.01 is taken to be supportive of the more restrictive model (Milfont & Fischer, 2010; van de Schoot et al., 2012). It is recommended that three levels of statistical equivalence are evaluated (van de Schoot et al., 2012) i.e. configural equivalence, metric equivalence and scalar invariance. Configural equivalence is achieved when items in the measuring instrument show the same pattern of factor loadings within each group. Metric equivalence (second level) indicates whether or not respondents from different groups respond to the questions in a similar manner. It requires that the factor loadings linking items and a construct are equal, and is an indicator of similarity of measurement unit (the scale metric). Scalar invariance (third level), which requires equality in both factor loading and intercepts across groups. Mean score comparisons are only permissible when one achieves scalar (full or partial) invariance and metric (full or partial) invariance then. Only then is it permissible to compare the relationship between variables across groups (Milfont & Fischer, 2010).

All analysis except for Omega coefficients were done using STATA version 13.1 (StataCorp, College Station, Tex).

Results
Complete data for this analysis was obtained from 1496 participants, whose characteristics are shown in Table 1. Only 2% could not work with ACASI and therefore were assisted in completing the interviews. All of the respondents had some form of education; 62% had reached or were still in primary school, while 3% had reached or were still in tertiary institutions. Overall, 11% admitted to have taken alcohol within 30 days prior to the interview and 25% reported being bullied at least once within 30 days preceding the interview.

Reliability
Cronbach’s alpha for MDI was 0.83, which is above the acceptable threshold score of 0.70. We also obtained an omega coefficient of 0.82 (95% CI 0.81-0.83) indicating that regardless of the method we used the internal consistency was acceptable for our sample. (Table 2)

Factorial structure
To evaluate the factorial structure of the MDI, we only utilized data of those who completed the interview in Swahili (N=1386), since there were insufficient number of participants that completed the interview in English (n=57) and Giryama (n=53).

The results of the exploratory factor analysis showed a single factor that had an eigenvalue of 3.44 with the other factors having eigenvalues of less than 0.3. All the 10 items appeared to load well onto the single factor that had the highest eigenvalue.
### Table 1. Socio-demographic characteristics of respondents.

<table>
<thead>
<tr>
<th></th>
<th>Total Sample (N=1496)</th>
<th>13–14 yrs (n=351)</th>
<th>15–19 yrs (n=689)</th>
<th>20–24 yrs (n=456)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Sex</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>773</td>
<td>187 (24.2)</td>
<td>359 (46.4)</td>
<td>227 (29.4)</td>
</tr>
<tr>
<td>Female</td>
<td>723</td>
<td>164 (22.7)</td>
<td>330 (45.6)</td>
<td>220 (31.7)</td>
</tr>
<tr>
<td><strong>Education</strong>*</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Primary</td>
<td>921</td>
<td>320 (34.7)</td>
<td>506 (54.9)</td>
<td>95 (10.3)</td>
</tr>
<tr>
<td>Secondary</td>
<td>128</td>
<td>3 (2.3)</td>
<td>61 (47.7)</td>
<td>64 (50.0)</td>
</tr>
<tr>
<td>Tertiary</td>
<td>40</td>
<td>0 (0)</td>
<td>10 (25.0)</td>
<td>30 (75.0)</td>
</tr>
<tr>
<td>Other/None</td>
<td>310</td>
<td>26 (8.4)</td>
<td>96 (31.0)</td>
<td>188 (60.7)</td>
</tr>
<tr>
<td>Missing</td>
<td>97</td>
<td>2 (2.1)</td>
<td>16 (16.5)</td>
<td>79 (81.4)</td>
</tr>
<tr>
<td><strong>Drinking habits</strong>*</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Non-drinkers</td>
<td>1233</td>
<td>327 (26.4)</td>
<td>578 (46.9)</td>
<td>328 (26.6)</td>
</tr>
<tr>
<td>Drinkers</td>
<td>169</td>
<td>9 (5.3)</td>
<td>62 (36.7)</td>
<td>98 (58.0)</td>
</tr>
<tr>
<td>Missing</td>
<td>94</td>
<td>15 (16.0)</td>
<td>49 (52.1)</td>
<td>30 (31.9)</td>
</tr>
<tr>
<td><strong>History of bullying</strong>*</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Never bullied</td>
<td>1118</td>
<td>257 (23.0)</td>
<td>509 (45.5)</td>
<td>352 (31.5)</td>
</tr>
<tr>
<td>Less bullied</td>
<td>313</td>
<td>79 (25.2)</td>
<td>151 (48.2)</td>
<td>83 (26.5)</td>
</tr>
<tr>
<td>Often bullied</td>
<td>59</td>
<td>15 (25.4)</td>
<td>28 (47.5)</td>
<td>16 (27.1)</td>
</tr>
<tr>
<td>Missing</td>
<td>6</td>
<td>0 (0)</td>
<td>1 (16.7)</td>
<td>5 (83.3)</td>
</tr>
</tbody>
</table>

### Table 2. Measure of internal consistency using Cronbach’s alpha and Omega coefficient.

<table>
<thead>
<tr>
<th></th>
<th>Mean MDI</th>
<th>Cronbach’s Alpha</th>
<th>Omega (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Overall</strong></td>
<td>7.25</td>
<td>0.83</td>
<td>0.82 (0.81, 0.83)</td>
</tr>
<tr>
<td><strong>Sex</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>6.77</td>
<td>0.81</td>
<td>0.80 (0.79, 0.82)</td>
</tr>
<tr>
<td>Female</td>
<td>7.79</td>
<td>0.85</td>
<td>0.83 (0.82, 0.85)</td>
</tr>
<tr>
<td><strong>Age</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>13–14</td>
<td>5.43</td>
<td>0.84</td>
<td>0.83 (0.81, 0.86)</td>
</tr>
<tr>
<td>15–19</td>
<td>7.51</td>
<td>0.82</td>
<td>0.81 (0.79, 0.83)</td>
</tr>
<tr>
<td>20–24</td>
<td>8.42</td>
<td>0.84</td>
<td>0.82 (0.80, 0.85)</td>
</tr>
<tr>
<td><strong>Language</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Swahili</td>
<td>7.22</td>
<td>0.84</td>
<td>0.82 (0.81, 0.84)</td>
</tr>
<tr>
<td>English</td>
<td>6.59</td>
<td>0.75</td>
<td>0.76 (0.62, 0.89)</td>
</tr>
<tr>
<td>Giriama</td>
<td>8.13</td>
<td>0.75</td>
<td>0.72 (0.61, 0.83)</td>
</tr>
</tbody>
</table>
All the items with exception of item 7 (difficulty in concentrating) had Loewinger’s coefficient greater than 0.3. Table 3 shows the factor loadings and Loewinger’s coefficient of the 10 items of the MDI. Overall Loewinger’s coefficient for the 10 items was 0.38 indicating that the MDI is a unidimensional scale.

Using confirmatory factor analytic approaches, we tested a single factor model as originally conceptualized by the developers of the MDI in the original population. The data indicated that the model did have a good fit to the data at the configural level with the fit indices being within acceptable limits, χ^2 (35, N = 1386) = 271.023, p < 0.001, TLI = 0.916, CFI = 0.935, and RMSEA = 0.070 (fit indices prior to adding the correlated error). However on inspection of the modification indices, we found that for several items there was a large residual error, suggesting correlation between items. These items were correlated since theoretically, they may have a shared variance beyond that in the other items. The correlated items were 3 and 5 (lack of energy and feelings of guilt) and 2 and 8 (low interest in daily living activities and being restless or subdued). After modifying the model to improve its fit (i.e. by accounting for the correlation between the items 3 and 5), the model had a better fit to the data at the configural level with the fit indices in the modified model changing to: (χ^2 (33, N = 1386) = 173.221 p < 0.001, TLI = 0.947, CFI = 0.961, and RMSEA = 0.055).

**Invariance analysis**

Using MGCFA, we investigated the equivalence of the identified model across age and sex. Results indicated that the scale achieved scalar invariance across sex. The fit indices in this case were: χ^2 (84, N = 1386) = 295.799, p < 0.001, RMSEA = 0.060, TLI = 0.937, CFI = 0.941, and ΔCFI= 0.004 (see Table 4).

### Table 3. Factor loadings from exploratory factor analysis (EFA) of the 10-item Major Depression Inventory.

<table>
<thead>
<tr>
<th>Item</th>
<th>Factor 1</th>
<th>Loewinger’s coefficient of homogeneity</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>0.59</td>
<td>0.37</td>
</tr>
<tr>
<td>2</td>
<td>0.57</td>
<td>0.37</td>
</tr>
<tr>
<td>3</td>
<td>0.65</td>
<td>0.41</td>
</tr>
<tr>
<td>4</td>
<td>0.62</td>
<td>0.38</td>
</tr>
<tr>
<td>5</td>
<td>0.65</td>
<td>0.40</td>
</tr>
<tr>
<td>6</td>
<td>0.53</td>
<td>0.37</td>
</tr>
<tr>
<td>7</td>
<td>0.45</td>
<td>0.29</td>
</tr>
<tr>
<td>8</td>
<td>0.67</td>
<td>0.43</td>
</tr>
<tr>
<td>9</td>
<td>0.58</td>
<td>0.37</td>
</tr>
<tr>
<td>10</td>
<td>0.53</td>
<td>0.37</td>
</tr>
<tr>
<td><strong>All items</strong></td>
<td></td>
<td>Eigen value = 0.53 Hs=0.38</td>
</tr>
</tbody>
</table>

*Hs – Loewinger’s scalability coefficient for MDI

### Table 4. Measurement Invariance.

<table>
<thead>
<tr>
<th></th>
<th>χ^2 (df)</th>
<th>Δχ^2 (Δdf)</th>
<th>RMSEA</th>
<th>TLI</th>
<th>CFI</th>
<th>ΔCFI</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Sex</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Configural</td>
<td>247.06 (66)</td>
<td>0.063</td>
<td>0.932</td>
<td>0.950</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Metric</td>
<td>272.65 (75)</td>
<td>25.58 (9)</td>
<td>0.062</td>
<td>0.934</td>
<td>0.945</td>
<td>0.005</td>
</tr>
<tr>
<td>Scalar</td>
<td>295.80 (84)</td>
<td>23.15 (9)</td>
<td>0.060</td>
<td>0.937</td>
<td>0.941</td>
<td>0.004</td>
</tr>
<tr>
<td><strong>Age</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Configural</td>
<td>311.55 (99)</td>
<td>0.068</td>
<td>0.921</td>
<td>0.942</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Metric</td>
<td>367.54 (117)</td>
<td>55.99 (18)</td>
<td>0.068</td>
<td>0.921</td>
<td>0.931</td>
<td>0.011</td>
</tr>
<tr>
<td>Scalar</td>
<td>416.25 (135)</td>
<td>48.70 (18)</td>
<td>0.067</td>
<td>0.923</td>
<td>0.923</td>
<td>0.008</td>
</tr>
</tbody>
</table>
Additionally, factor loadings for each item by sex were all significant (Figure 1). Similarly, we investigated the equivalence of the identified model across age groups as previously defined. Scalar invariance was once again achieved and the fit indices were: $\chi^2 (135, N = 1,386) = 416.25, p < 0.001$, TLI = 0.923, RMSEA = 0.067, CFI = 0.923, and $\Delta$CFI = 0.008.

Prevalence of major depressive symptoms

We observed that 3.6% [95%CI 2.7, 4.7] of our youth presented with DSM-IV major depressive symptoms. Female and older adolescents had higher prevalence of DSM-IV major depressive symptoms. No statistically significant difference in the prevalence of DSM-IV major depressive symptoms was observed across sex and age sub-groups.

We carried out additional analysis of the summed MDI scores and based on the conventional threshold score (i.e. higher than 20 indicating the presence of depression), we found that that 8.7% [95% CI 7.3, 10.2] of the youth had depressive symptoms. In this regard, the females had significantly higher prevalence (P-Value 0.0053) of 10.8% [95% CI 8.6, 13.3] compared to that of males at 6.7% [95%CI 5.1, 8.7]. The highest prevalence of depressive symptoms was observed among females aged 20–24 years (see Table 5).

![Factorial structure of the MDI items by sex. Italicized numbers in parentheses represents results from females.](image)

Figure 1. Factorial structure of the MDI items by sex. Italicized numbers in parentheses represents results from females.
Moderate (25–29) Bontempo, 1971). This possibly implies that Fawzi Mokken, 1971 (30 or more) Foxcroft, 2004). We have Mild Severe, 2004 DSM-IV % [CI] and the Arabic version which had an only unidimensional, but also that sub-groups i.e. males versus females and different age groups responded in a similar manner to the questions. The results add to the extant literature by showing that this measure, which is already widely used; can be applied across sub-samples in a statistically equivalent manner.

**Discussion**

We set out to evaluate the psychometric properties of the MDI. Our results indicate that the MDI in rural Kenya presents with acceptable internal consistency. Our results are supported by other studies that have evaluated the MDI in other geographical settings for example the Greek version (alpha of 0.86) (Fountoulakis et al., 2003) and the Arabic version which had an alpha of (0.91) (Fawzi et al., 2012). This possibly implies that the internal consistency of MDI is fairly stable across regions. Moreover, our high internal consistence obtained using the Cronbach’s alpha is further endorsed by the high Omega coefficient that we obtained.

Our findings from the exploratory factor analysis also indicate that MDI was measuring a single construct in our population. This conclusion was supported by the acceptable range of the Loewinger’s H coefficients (Mokken, 1971; Olsen et al., 2003) that we obtained when we utilized other modern psychometric methods besides exploratory factor analysis.

We carried out a series of multi-group confirmatory factor analysis to evaluate the factorial structure and invariance of this structure across sex and age groups. Evaluating the invariance of a scale across different sub-samples is very important as studies indicate that when scales that are non-invariant are used to compare results there is a risk of reaching invalid conclusions (Bontempo et al., 2007; Chen, 2008). Our results indicate that MDI was not only unidimensional, but also that sub-groups i.e. males versus females and different age groups responded in a similar manner to the questions. The results add to the extant literature by showing that this measure, which is already widely used; can be applied across sub-samples in a statistically equivalent manner.

An important challenge in carrying out assessments in Africa is the multilingual nature of the population (Foxcroft, 2004). In the Kenyan context for instance, an average youth may be able to speak three languages with different levels of fluency. Therefore, the question that arises is: which is the best language to use in the administration of the question? In our case we had three options, Swahili, Giryama and English, with participants choosing at the touch of a button, what language they would like to use at any point of the test. We had planned to evaluate the language equivalence of the measures to ensure that the use of different languages does not threaten validity, but most of our participants preferred to use the Swahili version making it difficult to evaluate the other language versions. Future work, will aim to target the evaluation of MDI in this other languages to provide us with information that allows us to utilize it with confidence.

We found the prevalence of DSM-IV depression and ICD-10 depression to be 3.6% and 8.7% respectively, which was consistent with the study conducted on the Danish general population which had a prevalence for DSM-IV depression and ICD-10 depression of 3.5% and 8.8% respectively for individuals aged 20–34 years. The main difference between the two countries is that Denmark is a high resource country compared to Kenya. This consistency gives further evidence of the robustness of the construct of depression in different settings (Olsen et al., 2004). We have shown that the MDI measures a single construct from evaluating

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**Table 5. Mean scores and prevalence of major depressive symptoms.** *: P<0.05, significant difference between two groups.

<table>
<thead>
<tr>
<th></th>
<th>Mean MDI Score (CI)</th>
<th>MDI Score ≥ 20 (20–24)</th>
<th>Mild (25–29)</th>
<th>Moderate (30 or more)</th>
<th>Severe Symptoms of Major Depression</th>
<th>DSM-IV % [CI]</th>
</tr>
</thead>
<tbody>
<tr>
<td>All (n=1496)</td>
<td>7.3 [6.9, 7.6]</td>
<td>8.7% [7.3, 10.2]</td>
<td>4.3% [3.3, 5.4]</td>
<td>2.5% [1.7, 3.4]</td>
<td>1.9% [1.3, 2.8]</td>
<td>3.6% [2.7, 4.7]</td>
</tr>
<tr>
<td>Females (n=723)</td>
<td>7.8* [7.2, 8.3]</td>
<td>10.8%* [8.6, 13.3]</td>
<td>5.0% [3.5, 6.8]</td>
<td>2.9% [1.8, 4.4]</td>
<td>2.9%* [1.8, 4.4]</td>
<td>4.4% [3.0, 6.2]</td>
</tr>
<tr>
<td>Males (n=773)</td>
<td>6.8* [6.3, 7.2]</td>
<td>6.7%* [5.1, 8.7]</td>
<td>3.6% [2.4, 5.2]</td>
<td>2.1% [1.2, 3.3]</td>
<td>1.0%* [0.4, 2.0]</td>
<td>2.8% [1.8, 4.3]</td>
</tr>
<tr>
<td>Age 13–14 (n=351)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Both (351)</td>
<td>5.4 [4.8, 6.1]</td>
<td>6.0% [3.7, 9.0]</td>
<td>3.4% [1.8, 6.0]</td>
<td>0.8% [0.2, 2.5]</td>
<td>1.7% [0.6, 3.7]</td>
<td>2.6% [1.2, 4.8]</td>
</tr>
<tr>
<td>Females (164)</td>
<td>5.9 [4.9, 7.2]</td>
<td>9.1%* [4.7, 13.6]</td>
<td>5.5%* [2.5, 10.2]</td>
<td>1.2% [0.1, 4.3]</td>
<td>2.4% [0.7, 6.1]</td>
<td>4.3% [1.7, 8.6]</td>
</tr>
<tr>
<td>Males (187)</td>
<td>5.0 [4.2, 5.8]</td>
<td>3.2%* [0.7, 5.7]</td>
<td>1.6%* [0.3, 4.6]</td>
<td>0.5% [0.3, 3.0]</td>
<td>1.1% [0.1, 3.8]</td>
<td>1.1% [0.1, 3.8]</td>
</tr>
<tr>
<td>Age 15–19 (n=689)</td>
<td></td>
<td></td>
<td></td>
<td></td>
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</tr>
<tr>
<td>Both (689)</td>
<td>7.5 [7.0, 8.0]</td>
<td>8.9% [6.8, 11.2]</td>
<td>4.9% [3.4, 6.8]</td>
<td>2.2% [1.2, 3.6]</td>
<td>1.7% [0.9, 3.0]</td>
<td>3.0% [1.9, 4.6]</td>
</tr>
<tr>
<td>Females (330)</td>
<td>7.8 [7.1, 8.6]</td>
<td>10.3% [7.0, 13.6]</td>
<td>4.8% [2.8, 7.8]</td>
<td>2.4% [1.1, 4.7]</td>
<td>3.0%* [1.5, 5.5]</td>
<td>3.3% [1.7, 5.9]</td>
</tr>
<tr>
<td>Males (359)</td>
<td>7.2 [6.6, 7.9]</td>
<td>7.5% [4.8, 10.2]</td>
<td>5.0% [3.0, 7.8]</td>
<td>1.9% [0.8, 4.0]</td>
<td>0.6%* [0.1, 2.0]</td>
<td>2.8% [1.3, 5.1]</td>
</tr>
<tr>
<td>Age 20–24 (n=456)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Both (456)</td>
<td>8.4 [7.8, 9.1]</td>
<td>10.5% [7.9, 13.7]</td>
<td>3.9% [2.4, 6.2]</td>
<td>4.2% [2.5, 6.4]</td>
<td>2.4% [1.2, 4.3]</td>
<td>5.3% [3.4, 7.7]</td>
</tr>
<tr>
<td>Females (229)</td>
<td>9.2* [8.2, 10.2]</td>
<td>12.7% [8.4, 17.0]</td>
<td>4.8% [2.4, 8.4]</td>
<td>4.8% [2.4, 8.4]</td>
<td>3.1% [1.2, 6.2]</td>
<td>6.1% [3.4, 10.0]</td>
</tr>
<tr>
<td>Males (227)</td>
<td>7.7* [6.9, 8.6]</td>
<td>8.4% [4.8, 12.0]</td>
<td>3.1% [1.2, 6.3]</td>
<td>3.5% [1.5, 6.8]</td>
<td>1.8%* [0.5, 4.5]</td>
<td>4.4% [2.1, 8.0]</td>
</tr>
</tbody>
</table>
its factorial structure and shown that the MDI is a unidimensional scale, using Item Response Models. These findings are consistent to the ones found in the original population from which the MDI was initially used.

As reported elsewhere, females in our sample reported higher prevalence of depressive symptoms and this seems to increase with age. In support of our findings, a recent study involving more than 10 European countries (Boyd et al., 2015) reported that “it has also been shown that females tend to report more depressive symptoms than males” (pg. 251). This is consistent with what has been reported in other parts of the world where females are usually at a higher risk of experiencing an episode of depression compared to males (Mazur-Mosiewicz et al., 2015; Pesola et al., 2015). This difference can also be due to gender-biased items in the MDI. For instance, some items (fear, loneliness, crying, weight loss, fatigue, suicide ideation, appetite loss) in the Beck Depression Inventory are known to have a greater value for girls (Salle et al., 2011). Some of these items are also in the MDI.

An additional methodological issue we wanted to evaluate was the acceptability and feasibility of using a computer-assisted assessment approaches. Due to low literacy levels, and a general lack of reading habits among members of our target population, using written questionnaires is not always possible in this context. We have previously used oral interviews when carrying out mental health, and other psychological assessment (Abubakar et al., 2016; Kariuki et al., 2012). However, oral interviews may lead to response bias since interviewees may not always give the most truthful answer when dealing with highly personal and sensitive issues. It was therefore encouraging to observe that only a few (2%) of the participants needed assistance through the test administration procedures, which is an indication, that computer-assisted administration is an acceptable approach in our context. Given the many advantages such as increased anonymity and feelings of privacy especially when dealing with sensitive topics, it is encouraging to know that this is a method with potential for further development. After the pilot, before the main survey, on comparing responses from the paper questionnaires versus responses from ACASI, we found that the responses were 100% similar showing just how feasible ACASI is as a mode of interviewing persons aged 13–24 years in our setting.

Most mental disorders are often misdiagnosed because general health workers lack the diagnostic skills and training required to manage such cases (Kiima et al., 2004). Having a screening tool such as the MDI will reduce the cases of misdiagnosis and given that it is self-administered, and easy to understand it can easily be deployed even in health centers where trained psychiatrists are not available.

Since the last prevalence study on depressive disorders in Kenya was done in the 90s, it is important to report prevalence studies of depressive disorders even if they are from specific subsets of the population in order to increase the evidence base. We think that the MDI can be used in other rural parts of Kenya similar to the population in Kilifi to screen for depressive symptoms.

By focusing efforts towards screening for depressive symptoms in adolescents and youth and providing timely interventions, many of the cases of depressive disorder that are diagnosed in older ages can be reduced.

Limitations
Our study has several strengths and adds salient information on the applicability of MDI in our context. Validity is a complex construct with various facets to it. We have been able to confirm structural validity, although we did not carry out clinical validation to evaluate the most suitable thresholds in this rural Kenyan population. This is essential to ensure that the current recommended approaches for estimating prevalence are valid for these settings. Future studies should focus on this, so as to further accumulate the evidence on the validity of the MDI. Future research should also examine test-re-test reliability and language invariance of the measure, particularly as the measure does not seem to have undergone cognitive pre-testing to ensure that the questions are understood as intended.

Conclusions
Our results indicate that MDI has acceptable psychometric properties and is a suitable instrument for screening depressive symptoms in young people living in rural Kenya.

Data availability
Data supporting the findings of this study are available from the KEMRI Institutional Data Access/Ethics committee, for researchers who meet the criteria for access to confidential data. Details of the criteria can be found in the KEMRI-Wellcome data sharing guidelines. Access to data is provided via the KEMRI-Wellcome Data Governance Committee: Data_Governance_Committee@kemri-wellcome.org; Tel, +254708587210; Contact person, Marianne Munene (Secretary; Tel, +254709983436).

Competing interests
No competing interests were disclosed.

Grant information
This work was supported by the Wellcome Trust [084535, 091758]; and the Hewlett Foundation through a grant awarded to INDEPTH [IND/0107].

The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Acknowledgements
We thank our colleagues who dedicated their time towards the design of the study, collection of data, analysis and commenting on the manuscript in its early stages. This paper is published with the permission of the director of Kenya Medical Research Institute (KEMRI). Hewlett Foundation through INDEPTH Network funded the Study.
References


Published Abstract | Publisher Full Text | Free Full Text

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Reference Source

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Open Peer Review

Current Referee Status: ✓ ✓

Version 1

Glyn Lewis
Division of Psychiatry, University College London, London, UK

This is a valuable study of the psychometric properties of the Major Depression Inventory (MDI) in Kenya. The authors rightly pose the issue about the use of self-administered scales in sub-Saharan Africa. However, they could also cite the work of Vikram Patel in Zimbabwe who developed the Shona Symptom Questionnaire within the local population there.

The paper is very well written and easy to follow. It is clearly laid out and logical in its flow. The analysis is very thorough and seemed an appropriate approach towards studying the psychometric properties of the scale.

The main limitations of the article are primarily the absence of any comparison with other potential instruments. Furthermore, there is no comparison with a more detailed assessment to provide a more thorough "gold standard" measure. The arguments that given the prevalence is similar in Kenya to Denmark is not convincing. One of the main purposes of these scales is to compare the prevalence in populations and I would have thought these could well be different in Kenya and Denmark.

The further limitation is why they chose the MDI. They cite papers suggesting it performed better than some other competitors, and it has the attraction of brevity and I presume is not copyrighted. However, there are so many of these scales it seems rather arbitrary that they have chosen the MDI. They state that the PHQ does not measure severity - but I don't think many people would agree with that. The PHQ is used as an outcome in the IAPT service in the UK and also in a number of clinical trials. I am not sure what they can do about this and they have tried to justify their choice in the intro.

Is the work clearly and accurately presented and does it cite the current literature? Yes

Is the study design appropriate and is the work technically sound? Yes

Are sufficient details of methods and analysis provided to allow replication by others? Yes

If applicable, is the statistical analysis and its interpretation appropriate? Yes
Are all the source data underlying the results available to ensure full reproducibility?
Yes

Are the conclusions drawn adequately supported by the results?
Yes

**Competing Interests:** No competing interests were disclosed.

**Referee Expertise:** psychiatric epidemiology

I have read this submission. I believe that I have an appropriate level of expertise to confirm that it is of an acceptable scientific standard.
subsequent results on categorical differences.

7. Here are some typos throughout the manuscript that should be corrected in the course of final copy editing.

8. Last not least, it would be of great importance to know where the Swahili and Giryama versions of the MDI can be obtained in order to promote their dissemination.

Is the work clearly and accurately presented and does it cite the current literature?
Yes

Is the study design appropriate and is the work technically sound?
Yes

Are sufficient details of methods and analysis provided to allow replication by others?
Yes

If applicable, is the statistical analysis and its interpretation appropriate?
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Yes

Are the conclusions drawn adequately supported by the results?
Yes

**Competing Interests:** No competing interests were disclosed.

We have read this submission. We believe that we have an appropriate level of expertise to confirm that it is of an acceptable scientific standard.