

Active trachoma and community use of sanitation, Ethiopia

William E Oswald,^a Aisha EP Stewart,^b Michael R Kramer,^c Tekola Endeshaw,^d Mulat Zerihun,^d Berhanu Melak,^d Eshetu Sata,^d Demelash Gessese,^d Tesfaye Teferi,^d Zerihun Tadesse,^d Birhan Guadie,^e Jonathan D King,^b Paul M Emerson,^b Elizabeth K Callahan,^b Dana Flanders,^c Christine L Moe^c & Thomas F Clasen^c

Objective To investigate, in Amhara, Ethiopia, the association between prevalence of active trachoma among children aged 1–9 years and community sanitation usage.

Methods Between 2011 and 2014, prevalence of trachoma and household pit latrine usage were measured in five population-based cross-sectional surveys. Data on observed indicators of latrine use were aggregated into a measure of community sanitation usage calculated as the proportion of households with a latrine in use. All household members were examined for clinical signs, i.e. trachomatous inflammation, follicular and/or intense, indicative of active trachoma. Multilevel logistic regression was used to estimate prevalence odds ratios (OR) and 95% confidence intervals (CI), adjusting for community, household and individual factors, and to evaluate modification by household latrine use and water access.

Findings In surveyed areas, prevalence of active trachoma among children was estimated to be 29% (95% CI: 28–30) and mean community sanitation usage was 47% (95% CI: 45–48). Despite significant modification ($p < 0.0001$), no pattern in stratified ORs was detected. Summarizing across strata, community sanitation usage values of 60 to $< 80\%$ and $\geq 80\%$ were associated with lower prevalence odds of active trachoma, compared with community sanitation usage of $< 20\%$ (OR: 0.76; 95% CI: 0.57–1.03 and OR: 0.67; 95% CI: 0.48–0.95, respectively).

Conclusion In Amhara, Ethiopia, a negative correlation was observed between community sanitation usage and prevalence of active trachoma among children, highlighting the need for continued efforts to encourage higher levels of sanitation usage and to support sustained use throughout the community, not simply at the household level.

Abstracts in [عربي](#), [中文](#), [Français](#), [Русский](#) and [Español](#) at the end of each article.

Introduction

It has been estimated that, as a result of trachoma, approximately 1.2 million people are blind and a further 1.7 million have low vision.¹ Globally, trachoma remains the leading infectious cause of blindness. In 2009, an estimated 40.6 million people had active trachoma and 8.2 million had trichiasis – i.e. the blinding stage of the disease.² About 77% of those living in trachoma-endemic areas of the world are to be found in 29 of the countries in the World Health Organization's (WHO's) African Region, and Ethiopia is the country most affected by trachoma worldwide.³ Trachoma is caused by ocular infection with a bacterium: *Chlamydia trachomatis*. Inflammation attributable to repeat infections during childhood constitutes the disease's active stage. This inflammation may then lead to scarring of the conjunctiva and trichiasis. In trichiasis, the eyelashes rub and damage the cornea, causing pain and, eventually, blindness. Trachoma is predominantly found in resource-poor, rural communities in low-income countries.^{4,5} By afflicting some of the most deprived people in the world, it leads to disability, dependency and further poverty.⁶

A pilot programme for trachoma control, begun in four districts of the Amhara region in north-west Ethiopia, was scaled-up so that the programme covered the whole of Amhara by 2007.^{7,8} This ongoing programme, based on WHO's SAFE strategy, was set in Amhara because this region has the highest burden of active trachoma within Ethiopia.⁹ The SAFE strategy is a comprehensive WHO strategy – based on the available

relevant biological and epidemiological evidence – to treat and prevent trachoma. It combines four measures: (i) surgery for the correction of trichiasis; (ii) antibiotics, given in mass drug administrations, to reduce the infection reservoir in the community; (iii) facial cleanliness, to reduce transmission; and (iv) environmental improvements – e.g. control of flies through sanitation and improved access to water for hygiene – for further reductions in the potential for transmission.^{10,11}

The presence of latrines or other facilities for the disposal of human faeces is understood to have an indirect beneficial effect on the risk of trachoma because it reduces the access of *Musca sorbens* flies – a probable vector of *C. trachomatis* – to potential breeding sites.^{12–14} As flies can easily move throughout an area, however, a few scattered latrines may have little impact on trachoma in that community – or even in households with latrines.⁶ Effective fly control through sanitation requires not only high levels of latrine access but also consistent latrine use throughout the community.¹⁵ Programmes to improve sanitation and/or control trachoma typically measure changes in sanitation at regional or national level. If they do investigate latrine coverage at community level at all, they tend to record latrine access rather than use. The role of latrine use – at community level – in trachoma control requires elucidation.¹⁶

We believed that, rather than latrine access at household level, the proportion of households in a community with latrines in use would be a stronger indicator of the effectiveness of fly control and levels of exposure to *C. trachomatis*.¹⁵ We hy-

^a London School of Hygiene and Tropical Medicine, Keppel Street, London, WC1E 7HT, England.

^b The Carter Center, Atlanta, United States of America (USA).

^c Rollins School of Public Health, Emory University, Atlanta, USA.

^d The Carter Center, Addis Ababa, Ethiopia.

^e Amhara Regional Health Bureau, Bahir Dar, Ethiopia.

Correspondence to William E Oswald (email: william.oswald@lshtm.ac.uk).

(Submitted: 13 May 2016 – Revised version received: 15 December 2016 – Accepted: 15 December 2016 – Published online: 26 January 2017)

pothesized that higher community sanitation usage would be associated with a lower prevalence of active trachoma.

Methods

Study overview and subjects

Between 2011 and 2014, the Amhara Regional Health Bureau conducted trachoma-impact surveys in various areas of Amhara. These surveys were designed to provide population-based estimates of trachoma prevalence, quantify uptake of trachoma control efforts and estimate the proportions of households with water and sanitation access.¹⁷ A district – known locally as a *woreda* – only became eligible for surveying when at least five annual rounds of mass administrations of azithromycin had occurred. Each survey was conducted at least six months after the last such antibiotic administration in the target district.

For the present study, we combined data collected in trachoma-impact surveys that were conducted in the South Gondar zone in June–August 2011,^{18,19} in the North Gondar and West Gojjam zones in May–June 2012, in eastern Amhara between December 2012 and January 2013, in western Amhara in June–July 2013, and in eastern Amhara in January–February 2014.

All five surveys used multistage cluster random sampling to estimate the district-level prevalence of trachomatous inflammation–follicular. Villages – known locally as *gott* – represented the smallest administrative unit with available population data and were the primary sampling units. Within each target district, villages were selected, from a geographically-ordered listing, using probability-proportional-to-size sampling. Within each selected village, smaller administrative units of approximately 40 households – i.e. household clusters that were locally called development teams – were used as segments for a modified segment design.^{20,21} Such clusters were listed and given identification numbers upon arrival in a study village with the assistance of a designated village representative, who then drew a number from a hat to select the cluster to be surveyed. The entire village was surveyed if it consisted of 40 or fewer households.

Community information was collected in interviews with village leaders. In each selected cluster, all residents who

gave verbal consent were examined for clinical signs of trachoma, according to WHO guidelines.¹⁷ Each eye was examined separately, by a trained trachoma grader using a 2.5x binocular loupe, for the presence or absence of all five clinical signs of the simplified trachoma grading system.²² Heads of household were interviewed for demographic and socioeconomic information as well as knowledge and practices regarding trachoma, hygiene, sanitation and water. Visual inspections were made of household latrines and hand-washing stations. Responses were recorded electronically using tablet computers with Swift Insights software (The Carter Center, Atlanta, USA). Questionnaires at community, household and individual levels were linked.²³

Measures

The exposure variable – community sanitation usage – was calculated as the proportion of households within the cluster with a latrine with evidence of use. A latrine was considered to be in use if there was a defined path to it and faeces were observed in the pit.²⁴

For trachoma, the outcome variable was a dichotomous measure, based on WHO's simplified grading scale, for the absence/presence of active trachoma – i.e. absence/presence of trachomatous inflammation – follicular and/or intense.¹⁷

Analyses

We used multilevel logistic regression to estimate the association between the proportion of households in each cluster with a latrine in use and active trachoma among children aged 1–9 years – accounting for dependence of observations nested within households and clusters. Multilevel analysis can assess the influence of area-level effects – e.g. community sanitation usage – on an individual outcome in addition to between-group and within-group variability.²⁵ These variance-based measures provide a useful complement to standard measures of association for the analysis of contextual effects.²⁶

Accounting for study design and unequal selection probabilities, means and proportions were estimated, with 95% confidence intervals (CI), across categories of community sanitation usage. Generalized linear mixed models were fitted, specifying random intercepts for cluster and households nested

within clusters. Models were estimated using adaptive quadrature with eight integration points, and robust standard errors (SE) were requested to account for clustering within districts. Sampling weights, based on the inverse selection probability for cluster, household and individual, were incorporated. Individual and household weights were scaled to sum to the household and cluster sample size, respectively.²⁷ After scaling weights, to restrict analysis to the subpopulation of interest, participants who were not aged 1–9 years were assigned an individual weight of 0.0001.²⁸ Our treatment of the exposure variable as a categorical measure – rather than a linear or quadratic measure – was based on the results of a preliminary assessment that considered both fit and interpretability. An empty model was fitted to measure between-cluster and between-household variance.²⁹ Potential confounders were identified, from the community-level, household and individual covariates recorded in the surveys (Table 1), based on the results of a literature review, an evaluation of directed acyclic graphs,^{31,32} univariable analyses and initial unweighted modelling. We used a sequential modelling approach to explore confounding – indicated by change in exposure estimates – and changes in residual variance. All models controlled for survey round, to account for year and possible between-survey variations in the method. Most surveys were conducted during the rainy season, so an assessment of seasonality was not possible. We calculated intraclass correlation coefficients by converting individual-level and area-level components of variance to the same scale – using the latent variable method – and median odds ratios (ORs) were calculated as measures of residual variance on the OR scale.^{29,33} To evaluate multiplicative modification of the effect of community sanitation usage on active trachoma, by household latrine use and water access, we used likelihood ratio tests, after controlling for covariates included in the fully-adjusted model. We calculated summary measures of association with CI from the weighted means of stratum-specific log ORs.^{34,35} Models included all observations with information available for included covariates. Initial unweighted modelling was conducted in SAS 9.4 (SAS Institute, Cary, United States of America). All described analyses were conducted

Table 1. Active trachoma prevalence and community, household and individual characteristics of children aged 1–9 years with eye examination results, categorized according to community sanitation usage, Amhara, Ethiopia, 2011–2014

Characteristics	n	Value (95% CI) for clusters where percentage of households with latrines with evidence of use was: ^a					All
		0 to < 20	20 to < 40	40 to < 60	60 to < 80	≥ 80	
% of children							
With active trachoma	62 869	0.31 (0.29–0.33)	0.33 (0.30–0.36)	0.30 (0.28–0.33)	0.26 (0.23–0.28)	0.23 (0.21–0.25)	0.29 (0.28–0.30)
With trachomatous inflammation							
Follicular	62 869	0.28 (0.26–0.31)	0.30 (0.27–0.34)	0.28 (0.25–0.31)	0.23 (0.21–0.25)	0.21 (0.19–0.23)	0.26 (0.25–0.27)
Intense	62 869	0.06 (0.05–0.07)	0.06 (0.05–0.07)	0.07 (0.05–0.08)	0.05 (0.04–0.06)	0.04 (0.04–0.05)	0.06 (0.05–0.06)
Community							
No. of children's communities	1 510	412	279	260	302	257	1 510
Median no. of MDA received ^b	1 510	2.37 (2.28–2.47)	2.62 (2.53–2.71)	2.72 (2.62–2.83)	2.72 (2.63–2.81)	2.91 (2.80–3.02)	2.65 (2.61–2.69)
People per square kilometer ^c	1 508	303 (206–399)	391 (280–501)	665 (348–982)	1 022 (701–1 343)	1 696 (1 244–2 149)	786 (670–901)
Proportion of communities with health facility ^d	1 414	0.15 (0.11–0.20)	0.28 (0.22–0.35)	0.28 (0.22–0.35)	0.26 (0.21–0.32)	0.33 (0.27–0.39)	0.25 (0.23–0.28)
No. of wealth indicators per household ^e	1 510	0.74 (0.68–0.79)	0.95 (0.87–1.03)	1.19 (1.09–1.28)	1.38 (1.27–1.49)	1.72 (1.58–1.87)	1.17 (1.13–1.21)
Household							
No. of children's households	35 977	10 279	6 764	6 070	7 043	5 821	35 977
% of households in which highest education of any adult was: ^f	35 595	–	–	–	–	–	–
None	–	0.56 (0.54–0.59)	0.53 (0.49–0.56)	0.49 (0.45–0.52)	0.43 (0.40–0.46)	0.43 (0.39–0.47)	0.49 (0.48–0.51)
Religious	–	0.02 (0.02–0.03)	0.03 (0.02–0.04)	0.03 (0.02–0.03)	0.03 (0.02–0.04)	0.02 (0.02–0.03)	0.03 (0.02–0.03)
Primary school	–	0.21 (0.19–0.23)	0.16 (0.14–0.18)	0.19 (0.16–0.21)	0.18 (0.16–0.20)	0.18 (0.16–0.21)	0.18 (0.18–0.19)
Junior secondary	–	0.13 (0.11–0.14)	0.17 (0.15–0.19)	0.17 (0.15–0.19)	0.19 (0.17–0.20)	0.17 (0.15–0.19)	0.16 (0.16–0.17)
Senior secondary	–	0.04 (0.03–0.04)	0.06 (0.05–0.08)	0.08 (0.07–0.09)	0.11 (0.10–0.12)	0.12 (0.10–0.14)	0.08 (0.07–0.08)
College or university	–	0.00 (0.00–0.01)	0.01 (0.01–0.01)	0.01 (0.01–0.02)	0.03 (0.02–0.04)	0.04 (0.03–0.05)	0.02 (0.02–0.02)
Non-formal	–	0.04 (0.03–0.05)	0.04 (0.03–0.05)	0.04 (0.03–0.05)	0.04 (0.03–0.05)	0.03 (0.02–0.04)	0.04 (0.03–0.04)

(continues. . .)

(...continued)

Characteristics	n	Value (95% CI) for clusters where percentage of households with latrines with evidence of use was: ^a					All
		0 to < 20	20 to < 40	40 to < 60	60 to < 80	≥ 80	
% of households owning mobile phone	35584	0.11 (0.09–0.13)	0.18 (0.15–0.21)	0.23 (0.20–0.26)	0.25 (0.21–0.28)	0.32 (0.29–0.36)	0.21 (0.19–0.22)
% of households owning radio	35599	0.11 (0.10–0.13)	0.13 (0.11–0.15)	0.18 (0.16–0.20)	0.22 (0.19–0.24)	0.24 (0.21–0.27)	0.17 (0.16–0.18)
% of households owning television	35589	0.00 (0.00–0.01)	0.01 (0.01–0.02)	0.02 (0.01–0.03)	0.05 (0.03–0.08)	0.12 (0.09–0.14)	0.04 (0.03–0.04)
% of households with bathing water within a journey of < 30 minutes	37502	0.61 (0.57–0.66)	0.69 (0.64–0.73)	0.68 (0.62–0.73)	0.70 (0.65–0.74)	0.74 (0.69–0.78)	0.68 (0.66–0.70)
% of households with a latrine	35863	0.12 (0.11–0.14)	0.38 (0.36–0.39)	0.58 (0.56–0.59)	0.77 (0.76–0.78)	0.93 (0.92–0.94)	0.52 (0.50–0.53)
% of households with a latrine with evidence of use	35762	0.08 (0.07–0.09)	0.32 (0.31–0.33)	0.53 (0.52–0.54)	0.74 (0.73–0.75)	0.92 (0.91–0.93)	0.48 (0.47–0.50)
% of households with any trachoma prevention knowledge ^b	35565	0.52 (0.48–0.55)	0.58 (0.53–0.62)	0.60 (0.55–0.64)	0.70 (0.66–0.72)	0.70 (0.66–0.73)	0.61 (0.60–0.62)
% of households with electricity	35584	0.02 (0.01–0.04)	0.06 (0.03–0.09)	0.08 (0.05–0.12)	0.18 (0.14–0.23)	0.26 (0.21–0.31)	0.11 (0.10–0.13)
% of households with iron roof	35782	0.51 (0.48–0.55)	0.58 (0.53–0.62)	0.67 (0.62–0.71)	0.73 (0.70–0.77)	0.75 (0.71–0.79)	0.64 (0.62–0.65)
Individual							
No. of children	62869	19484	11755	10162	11832	9636	62869
Age in years	62869	4.99 (4.94–5.04)	5.10 (5.04–5.17)	5.15 (5.09–5.22)	5.18 (5.12–5.25)	5.23 (5.15–5.30)	5.12 (5.09–5.14)
% being boys	62806	0.48 (0.47–0.49)	0.48 (0.47–0.49)	0.49 (0.47–0.50)	0.49 (0.48–0.50)	0.48 (0.47–0.49)	0.48 (0.48–0.49)
Reporting receipt of:	62158						
0 MDA		0.26 (0.23–0.29)	0.22 (0.19–0.24)	0.19 (0.16–0.21)	0.18 (0.15–0.21)	0.15 (0.13–0.17)	0.21 (0.19–0.22)
1 MDA		0.15 (0.14–0.16)	0.15 (0.14–0.17)	0.15 (0.14–0.16)	0.16 (0.14–0.17)	0.14 (0.13–0.15)	0.15 (0.14–0.16)
2 MDA		0.32 (0.30–0.34)	0.33 (0.31–0.35)	0.31 (0.29–0.34)	0.32 (0.30–0.34)	0.29 (0.27–0.32)	0.32 (0.31–0.33)
3 MDA		0.22 (0.20–0.24)	0.24 (0.22–0.27)	0.26 (0.24–0.29)	0.26 (0.24–0.29)	0.30 (0.27–0.32)	0.25 (0.24–0.26)
4 MDA		0.04 (0.03–0.05)	0.05 (0.04–0.07)	0.07 (0.06–0.09)	0.07 (0.05–0.08)	0.10 (0.08–0.12)	0.06 (0.06–0.07)
5 MDA		0.01 (0.01–0.02)	0.01 (0.00–0.01)	0.01 (0.01–0.02)	0.02 (0.01–0.03)	0.03 (0.02–0.04)	0.01 (0.01–0.02)
Reporting receipt of at least one MDA	62767	0.74 (0.72–0.77)	0.79 (0.76–0.81)	0.82 (0.79–0.84)	0.82 (0.79–0.85)	0.85 (0.83–0.88)	0.80 (0.78–0.81)

(continues...)

(...continued)

Characteristics	n	Value (95% CI) for clusters where percentage of households with latrines with evidence of use was: ^a				
		0 to <20	20 to <40	40 to <60	60 to <80	≥ 80
Reporting school attendance	61 185	0.17 (0.16–0.19)	0.21 (0.20–0.23)	0.24 (0.22–0.26)	0.25 (0.24–0.27)	0.29 (0.28–0.31)
Without ocular or nasal discharge	58 450	0.72 (0.69–0.75)	0.75 (0.72–0.78)	0.79 (0.76–0.82)	0.77 (0.74–0.80)	0.82 (0.80–0.85)

CI: confidence interval; MDA: mass drug administration.

^a All values followed by confidence intervals are means.

^b Reported by, by all residents.

^c Values extracted, using geographical coordinates in ArcMap 10.1 (ESRI, Redlands, USA), from raster surface with 2011 population density, generated using the Oak Ridge National Laboratory's LandScan.³⁰

^d A health centre, health post or hospital.

^e The household wealth indicators were ownership of mobile phone, ownership of radio, ownership of television, an iron roof and access to electricity.

^f The highest level of education reported to have been completed either by respondents in the 2011 survey or by any household member in the subsequent surveys.

^g That is, the unprompted reporting of one or more forms of trachoma prevention.

using Stata version 13.1 (StataCorp. LP, College Station, USA).

Ethical approval

The protocols for the surveys that were our data sources were approved by Emory University's Institutional Review Board and the Amhara Regional Health Bureau. Our secondary analysis was exempt from additional review.

Results

Of 56 425 households surveyed in 1510 clusters throughout Amhara region, 56 169 (> 99%) were linked to eye examination and census information, for 233 363 individuals. Of 68 961 children aged 1–9 years in the linked data set, 62 869 (91%), in 35 977 households, had eye examination results. Of 6092 children aged 1–9 years in the linked data set that lacked eye examination results, 4864 (80%) were reported to be out, travelling or at school during the survey, 734 (12%) refused the examination and 494 (8%) did not have a reason provided. Community sanitation usage was calculated using data on 56 050 (> 99%) of the surveyed households.

Table 1 summarizes overall, and by category of community sanitation usage, the community, household and individual characteristics of children aged 1–9 years with eye examination results. In general, compared with other children, those in communities with relatively low sanitation usage had indicators of poorer hygiene, more impoverished and rural living conditions and less education and health care. In terms of mean household counts of wealth indicators, communities with the lowest category of sanitation usage appeared much poorer than communities with the highest category of such usage (Table 1).

Levels of exposure to mass administrations of antibiotic for trachoma control and levels of prevention knowledge mirrored patterns of community sanitation usage. For example, children in communities with the lowest category of sanitation usage were less likely to have ever received antibiotics during mass administrations than children in communities with the highest category of such usage (Table 1). The median number of times that all community residents had reportedly received such antibiotics was lower in communities with the lowest category of sanitation use-

age than in communities with the highest category of usage. Some trachoma prevention knowledge was reported by more than half of households (Table 1) but such knowledge was less frequently reported in communities with relatively low sanitation usages. Where latrines were present, latrine usage was high. Among households, an estimated 52% (95% CI: 50–53) owned a pit latrine, the primary form of sanitation recorded. Of these pit latrines, 93% (95% CI: 92–93) were classified as in use based on observation. The mean reported age of latrines was 2.59 years (95% CI: 2.49–2.70), based on data from three of the surveys that recorded such data.

Mean community sanitation usage was 47% (95% CI: 45–48) over all 1510 clusters but ranged from 0% in 106 clusters to 100% in 25 clusters. The overall prevalence of active trachoma in children aged 1–9 years was estimated to be 29% (95% CI: 28–30) (Table 1).

We fitted several models, described as models 1–6, to examine the association between community sanitation usage and active trachoma – sequentially controlling for selected community, household and individual factors and adjusting for survey (Table 2). In terms of the available covariates, populations for all six models were similar. In model 1, which only adjusted for survey round, community sanitation usages of 60 to <80% and ≥80% were associated with lower prevalence odds of active trachoma compared with usage of <20%. Adding adjustments for child's age and sex (model 2) and then household water access and latrine use (model 3) did not meaningfully change the estimated ORs. After inclusion of household wealth indicators and education (model 4), an aggregated community wealth measure (model 5) and a community measure for the median number of times that a mass antibiotic administration had been received (model 6), the pattern remained the same but was attenuated towards null (Table 2).

In the empty model, which excluded any covariates, residual variances between households and between clusters were estimated to be 0.68 (SE: 0.08) and 2.22 (SE: 0.20), respectively. Intraclass correlation coefficients for the same household and for different households in the same cluster were calculated as 0.47 and 0.36, respectively. In this study, median ORs indicate the extent, always greater than or equal to 1,

Table 2. Results from six models of the association between active trachoma in children aged 1–9 years and community sanitation usage, Amhara, Ethiopia, 2011–2014

Characteristics	Model ^a					
	1	2	3	4	5	6
No. of individuals	62 869	62 806	62 037	61 351	61 351	61 351
No. of households	35 977	35 963	35 477	35 061	35 061	35 061
No. of clusters of households	1 510	1 510	1 510	1 510	1 510	1 510
aOR (95% CI) with community sanitation usage of: ^b						
≥ 80%	0.51 (0.35–0.73)	0.50 (0.33–0.75)	0.54 (0.36–0.81)	0.62 (0.42–0.92)	0.74 (0.49–1.11)	0.68 (0.45–1.04)
60 to < 80%	0.67 (0.47–0.97)	0.68 (0.45–1.03)	0.73 (0.49–1.10)	0.78 (0.52–1.17)	0.88 (0.58–1.32)	0.84 (0.55–1.26)
40 to < 60%	0.91 (0.66–1.26)	0.95 (0.66–1.35)	1.00 (0.70–1.42)	1.02 (0.72–1.46)	1.11 (0.78–1.58)	1.06 (0.74–1.52)
20 to < 40%	1.03 (0.76–1.39)	1.07 (0.77–1.49)	1.10 (0.79–1.53)	1.11 (0.80–1.54)	1.16 (0.83–1.60)	1.12 (0.81–1.55)
< 20%	Ref.	Ref.	Ref.	Ref.	Ref.	Ref.
Random effects						
Intraclass correlation coefficient						
Same cluster	0.34	0.37	0.37	0.36	0.36	0.35
Same household	0.45	0.52	0.52	0.51	0.51	0.51
Median OR						
Cluster	4.80	6.07	6.02	5.86	5.83	5.82
Households within cluster	2.19	2.66	2.65	2.63	2.63	2.63
Variance between clusters (SE)	2.03 (0.18)	2.52 (0.23)	2.50 (0.22)	2.41 (0.22)	2.39 (0.22)	2.38 (0.22)
Variance between households (SE)	0.68 (0.08)	1.05 (0.10)	1.04 (0.10)	1.03 (0.10)	1.03 (0.10)	1.03 (0.10)

aOR: adjusted odds ratio; CI: confidence interval; OR: odds ratio; Ref.: reference group; SE: standard error.

^a Results were weighted to account for the unequal probabilities of selection and down-weight survey participants excluded from the assessment of trachoma prevalence because they were aged < 1 year or > 9 years. Model 1 was adjusted for survey round. Model 2 was like model 1 but also adjusted for sex and age of the child— with age centred at 5 years. Model 3 was like model 2 but also adjusted for household access to bathing water within a journey of < 30 minutes and household latrine with evidence of use. Model 4 was like model 3 but also adjusted for household

access to electricity, household education, household ownership of mobile phone, radio and/or television and iron roof. Model 5 was like model 4 but also adjusted for community mean household wealth. Model 6 was like model 5 but also adjusted for the median number of times surveyed residents in a community had reportedly received a mass drug administration against trachoma.

^b Usage was measured as the proportion of households with a latrine with evidence of use.

Table 3. Modification by household latrine use and water access of the association between active trachoma in children aged 1–9 years and community sanitation usage, Amhara, Ethiopia, 2011–2014

Community sanitation usage (% of households with latrines in apparent use)	Household owns latrine with evidence of use				Household does not own latrine with evidence of use ^a				n	sOR (95% CI) ^b
	Water access < 30 minutes		Water access ≥ 30 minutes		Water access < 30 minutes		Water access ≥ 30 minutes			
	n	OR (95% CI)	n	OR (95% CI)	n	OR (95% CI)	n	OR (95% CI)		
≥ 80	6 383	0.62 (0.39–0.99)	2 332	0.51 (0.32–0.83)	518	0.68 (0.41–1.13)	182	0.93 (0.52–1.67)	9 415	0.67 (0.48–0.95)
60 to < 80	5 870	0.83 (0.54–1.29)	2 690	0.54 (0.32–0.90)	2 006	0.87 (0.55–1.37)	982	0.68 (0.42–1.09)	11 548	0.76 (0.57–1.03)
40 to < 60	3 636	0.87 (0.56–1.36)	1 741	0.76 (0.46–1.24)	3 094	1.10 (0.75–1.62)	1 471	1.35 (0.92–1.96)	9 942	1.01 (0.74–1.37)
20 to < 40	2 587	1.14 (0.75–1.72)	1 237	0.72 (0.46–1.13)	5 282	1.07 (0.74–1.56)	2 392	1.22 (0.87–1.69)	11 498	1.06 (0.78–1.44)
< 20	1 090	Ref.	492	Ref.	10 679	Ref.	6 687	Ref.	18 948	Ref.

CI: confidence interval; OR: odds ratio; Ref.: reference group; sOR: summary odds ratio.

^a Because the household has no latrine or has a latrine considered to be unused.

^b Summary OR weighted by total population within strata of household water and sanitation access, adjusted for age, sex, household education, household wealth items, community mean sum of household wealth indicators, median number of times surveyed residents in a community had reportedly received a mass drug administration against trachoma and survey round.

to which the individual probability of active trachoma was determined by cluster and household levels. In a comparison of children of different households from different communities, the median OR was calculated to be 5.07. In a comparison of children of different households from the same community, the corresponding ratio was 2.19. Variance did not meaningfully change across additional models (Table 2). Residual heterogeneity between clusters was of greater relevance than community sanitation usage. In model 6, median ORs indicated that residual heterogeneity between children of different communities reflected, on average, a 5.82-fold increase in the individual odds of active trachoma.

The magnitude of association between community sanitation usage and active trachoma was found to vary significantly by household latrine use and water access ($P < 0.0001$). As no clear pattern in stratified OR estimates and CIs was discerned, summary estimates, weighted by population in the strata of household latrine use and water access, are reported by category of community sanitation usage in Table 3.

Discussion

Our study shows that increasing the proportion of households in a community with latrines in use may be protective against active trachoma among children aged 1–9 years, independent of whether a child's household had a latrine in use or better access to water and controlling for potential confounders. There was no clear evidence of multiplicative modification of the effect of community sanitation usage on active trachoma by household latrine use and water access. Multilevel analysis, which allowed estimation of residual variation between communities and households, indicated the importance of additional contextual factors – beyond community sanitation usage and other measures that we included in our models – that may have more influence on an individual's propensity for active trachoma.

Studies in the Gambia identified the fly, *M. sorbens* – that breeds in openly-deposited faeces of humans and other mammals but not in pit latrines – as an insect vector of trachoma, clarifying the relationship between faeces in the environment and the disease.^{12–14} Subsequently, in a randomized controlled trial, fly catches from children's eyes

and mean active trachoma prevalence were reduced through latrine provision – but not by a statistically significant amount.³⁶ In another randomized controlled study – designed to measure the effect of latrine promotion on re-emergence of trachoma after a mass administration of antibiotic – there was no evidence of a significant relationship between increased latrine provision and prevalence of active trachoma or *C. trachomatis* infection in children because there was no rapid re-emergence of infection in either study arm.³⁷ In a later cohort analysis of the communities that had received a single mass azithromycin distribution and promotion of latrine usage, it was found that, for each 10% increase in the proportion of household latrines with evidence of use 12 months after baseline, there was a 2.0% decrease (95% CI: 0.2–3.9) in the community prevalence of ocular *C. trachomatis* infection over the subsequent year.³⁸ However, no corresponding decrease in the prevalence of active trachoma was observed,³⁸ perhaps because the follow-up period was too short to allow the beneficial impact of a cleaner living environment on the occurrence of active trachoma to become apparent.³⁹ Based on available data, most latrines observed in the surveys we used had been in place for more than 12 months.

In the control of neglected tropical diseases, the relative importance of hygiene, sanitation and water components, and of household sanitation compared with community sanitation, remains to be established.^{16,40} Improvements in access to water supplies could lead to increased facial cleansing, one of the four components of the SAFE strategy, by increasing the quantity of water available at household level. However, we did not identify any pattern of difference

in the association between community sanitation usage and active trachoma by household water access and latrine usage, despite statistically significant interaction.

Few studies have examined the relationship between community sanitation usage and health outcomes.^{16,41–44} The conclusions from this study are strengthened by its size, population-based estimates and consideration of latrine use rather than latrine ownership. Our study had limitations. The cross-sectional surveys prevent causal conclusions, and residual confounding remains possible. We could not control for hygiene practices because few relevant measures were collected in the surveys. Also, our indicator of latrine use did not measure usage by all household members or the disposal of children's faeces. Therefore, we had to assume that the proportion of households with a latrine with evidence of use reflected the actual proportion of the community population that consistently deposited their faeces in a latrine.²⁴

Although we observed variation in the prevalence of active trachoma at district-level, our estimate of the overall prevalence of active trachoma among children aged 1–9 years in Amhara, i.e. 29%, indicates the need for continued local implementation of the SAFE strategy. At the time of our study, despite improvements over recent years across Ethiopia,⁴⁵ household latrine usage in Amhara remained below 50%. Trachoma control efforts should continue to emphasize environmental improvements. The association of community sanitation usage with trachoma highlights the need for interventions – particularly ones targeting the environmental component of the SAFE strategy – to create communities free from open

defecation.⁴⁶ By modelling the association between community sanitation usage and *C. trachomatis* infection, it might be possible to clarify the role of sanitation in preventing transmission of the causative agent of trachoma. Future research should focus on both increasing the adoption of latrines – to reach protective levels of community sanitation usage – and improving latrine construction and maintenance – to ensure that any usage improvements are sustained. ■

Acknowledgements

We thank the Amhara National Regional Health Bureau and health offices, The Carter Center support staff, field teams, study supervisors and the residents of selected communities. We also thank Stephanie Ogden and Matthew Freeman, Rob O'Reilly and the Emory Center for Digital Scholarship and David Kleinbaum.

Funding: WEO was supported by the Emory University Laney Graduate School, ARCS Foundation Atlanta and the Global 2000 programme of The Carter Center. This study was supported by the Lions-Carter Center Sight-First Initiative and made possible thanks to the generous support of the American People through the United States Agency for International Development (USAID) and the ENVISION project led by RTI International in partnership with The Carter Center. The contents of this article are the responsibility of the authors and do not necessarily reflect the views of USAID or the United States Government.

Competing interests: None declared.

ملخص

الترخوما النشطة واستخدام المرايق الصحية المجتمعية في إثيوبيا

حاد، أو مؤشرات تدل على الإصابة بحالة من الترخوما النشطة. وتم استخدام التحوف اللوجستي متعدد المراحل لتقدير نسب احتمالات الانتشار (OR) ونسب أرجحية تبلغ 95% (CI)، مع تعديلها حسب العوامل المجتمعية والفردية والأسرية، وتقييم التعديل حسب استخدام العوائل للمرايحض وتوفر الماء لديهم. النتائج في مناطق الاستطلاع، تم تقدير انتشار الترخوما النشطة بين الأطفال بنسبة 29% (95% كمقدار لنسبة الأرجحية: 28 – 30) وتقدير متوسط استخدام المرايق الصحية المجتمعية بنسبة 47% (95% كمقدار لنسبة الأرجحية: 45–48). وبالرغم من التعديل الكبير (بمعدل احتمال > 0.0001)، لم يتم الكشف عن نمط معين

الغرض التحقيق في العلاقة بين انتشار الترخوما النشطة بين الأطفال ممن تتراوح أعمارهم من سنة إلى 9 سنوات واستخدام المرايق الصحية المجتمعية في أمهرة بإثيوبيا. الطريقة تم قياس انتشار الترخوما واستخدام المرايحض الحفري المنزلي في خمسة استطلاعات رأي شملت قطاعات متعددة وقائمة على السكان في الفترة ما بين سنة 2011 و 2014. وتم تجميع البيانات المتعلقة بالمؤشرات الملحوظة لاستخدام المرايحض للخروج بمقياس مدى استخدام المرايق الصحية المجتمعية المحتسب كنسبة العوائل التي تستخدم المرايحض. وقد تم فحص جميع أفراد الأسرة بحثاً عن علامات سريرية، أي التهاب تراخومي، أو مسامي، أو

الاستنتاج في أمهرة بإثيوبيا، لوحظ وجود ارتباط سلبي بين استخدام المرافق الصحية المجتمعية وانتشار التراخوما النشطة بين الأطفال، مع تسليط الضوء على الحاجة إلى مجهودات مستمرة لتشجيع المستويات الأعلى من استخدام المرافق الصحية، ودعم الاستخدام المستمر عبر المجتمع، وليس على مستوى الأسرة فحسب.

في نسب احتمالات الانتشار المطبقة. ومن خلال تلخيص البيانات عبر الشرائح المختلفة، فقد تم ربط نسب استخدام المرافق الصحية المجتمعية التي تتراوح من 60 إلى $>80\%$ و $\geq 80\%$ باحتمالات منخفضة لانتشار التراخوما، مقارنة باستخدام المرافق الصحية المجتمعية بنسبة $>20\%$ (احتمالات الانتشار: 0.76 ؛ 0.95 كمقدار لنسبة الأرجحية: $0.57 - 1.03$ ونسبة انتشار: 0.67 ؛ 0.95 كمقدار لنسبة الأرجحية: $0.48 - 0.95$ ، على التوالي).

摘要

埃塞俄比亚活动性沙眼病和社区卫生设施使用情况

目的 旨在调查埃塞俄比亚阿姆哈拉地区 1 到 9 岁儿童活动性沙眼患病率与社区卫生设施使用情况之间的联系。

方法 2011 年至 2014 年间，五项基于人口的横断面调查对沙眼患病率和家用蹲厕使用情况进行了衡量。我们汇总了观察到的厕所使用指标数据作为社区卫生设施使用情况的量测指标，按照使用厕所家庭的比例进行计算。检查了所有家庭成员的临床症状，例如沙眼性炎症、滤泡性和 / 或严重指示性活动性沙眼。我们采用多级逻辑回归模型估算了患病率比值比 (OR) 和 95% 置信区间 (CI)，调整了社区、家庭和个人因素，以评估通过家用厕所使用和供水情况进行的整改。

结果 在调查地区，儿童活动性沙眼患病率估计值

为 29% (95% CI: 28 - 30)，社区卫生设施使用率的平均值为 47% (95% CI: 45 - 48)。除重大调整项目 ($p < 0.0001$) 以外，未发现分层比值比的模式。汇总横断面数据得出，与社区卫生设施使用率低于 20% 的情况相比，社区卫生设施使用率在 60% 到 80% 之间和 80% 及以上的情况与较低的活动性沙眼患病率有关联 (分别为 OR: 0.76; 95% CI: 0.57 - 1.03 和 OR: 0.67; 95% CI: 0.48 - 0.95)。

结论 我们在埃塞俄比亚阿姆哈拉地区观察到，社区卫生设施使用率与儿童活动性沙眼患病率呈负相关，突出了需要继续努力促进更高水平的卫生设施使用率并支持在整个社会的持续使用，而不仅限于家庭层面上。

Résumé

Trachome actif et utilisation d'installations d'assainissement collectif en Éthiopie

Objectif Étudier dans la région Amhara, en Éthiopie, l'association entre la prévalence du trachome actif chez les enfants âgés de 1 à 9 ans et l'utilisation d'installations d'assainissement collectif.

Méthodes Entre 2011 et 2014, la prévalence du trachome et l'utilisation de latrines à fosse domestiques ont été mesurées dans le cadre de cinq enquêtes transversales menées auprès de la population. Les données sur les indicateurs observés concernant l'utilisation de latrines ont été regroupées dans une mesure de l'utilisation des installations d'assainissement collectif exprimée comme la proportion de ménages utilisant des latrines. Tous les membres des ménages ont fait l'objet d'un examen destiné à détecter les signes cliniques, c'est-à-dire une inflammation trachomateuse folliculaire et/ou intense, révélateurs d'un trachome actif. Une régression logistique à plusieurs niveaux a été utilisée pour estimer le rapport de cote (RC) de prévalence ainsi que les intervalles de confiance (IC) à 95%, en tenant compte de facteurs liés à la communauté, aux ménages et aux individus, et pour évaluer les variations induites par l'utilisation de latrines domestiques et un accès à l'eau.

Résultats Dans les zones étudiées, la prévalence du trachome actif chez les enfants a été estimée à 29% (IC à 95%: 28-30) et l'utilisation moyenne des installations d'assainissement collectif à 47% (IC à 95%: 45-48). Malgré des variations considérables ($p < 0,0001$), aucune tendance liée aux RC stratifiés ne s'est dégagée. Pour résumer en tenant compte des différentes strates, lorsque les valeurs relatives à l'utilisation des installations d'assainissement collectif étaient comprises entre 60 et 80% ou $\geq 80\%$, le rapport de cote de prévalence du trachome actif était plus faible que lorsque les valeurs relatives à l'utilisation des installations d'assainissement collectif étaient $< 20\%$ (RC: 0,76; IC à 95%: 0,57-1,03 et RC: 0,67; IC à 95%: 0,48-0,95, respectivement).

Conclusion Dans la région Amhara de l'Éthiopie, une corrélation négative a été observée entre l'utilisation d'installations d'assainissement collectif et la prévalence du trachome actif chez les enfants, soulignant la nécessité de poursuivre les efforts pour encourager l'utilisation d'installations d'assainissement et promouvoir une utilisation durable dans l'ensemble de la population, et non seulement au niveau des ménages.

Резюме

Активная трахома и использование санитарно-гигиенических удобств общиной, Эфиопия

Цель Изучить в регионе Амхара, Эфиопия, связь между распространенностью активной трахомы среди детей в возрасте 1–9 лет и общинным использованием санитарно-гигиенических удобств.

Методы В период с 2011 по 2014 год в ходе пяти поперечно-секционных обследований популяционного масштаба была изучена распространенность трахомы и использования ямного

туалета домашними хозяйствами. Данные по наблюдаемым показателям использования ямных туалетов были объединены, чтобы получить масштаб общинного использования санитарно-гигиенических удобств, выраженный долей домашних хозяйств, в которых используются ямные туалеты. Все члены семей были обследованы на предмет клинических признаков, т. е. трахоматозного воспаления, фолликулярного и (или)

интенсивного, указывающего на активную трахому. Многоуровневая логистическая регрессия была использована для оценки отношений шансов (ОШ) и 95%-х доверительных интервалов (ДИ) распространенности с учетом общинных, бытовых и индивидуальных факторов, а также для оценки влияния использования ямного туалета и доступа к воде в домашнем хозяйстве.

Результаты В обследованных районах распространенность активной трахомы среди детей оценивалась как 29% (95%-й ДИ: 28–30) и среднее значение использования санитарии общиной составляло 47% (95%-й ДИ: 45–48). Несмотря на значительное влияние ($p < 0,0001$), не было обнаружено никакой закономерности в стратифицированных ОШ. Подводя итог по стратам, можно отметить, что там, где показатели использования

санитарно-гигиенических удобств общиной составляли от 60 до $< 80\%$ и $\geq 80\%$, наблюдалась более низкая вероятность распространенности активной трахомы по сравнению со стратой, где показатель использования санитарно-гигиенических удобств общиной составил менее 20% (ОШ: 0,76; 95%-й ДИ: 0,57–1,03 и ОШ: 0,67; 95%-й ДИ: 0,48–0,95 соответственно).

Вывод В регионе Амхара, Эфиопия, наблюдалась отрицательная корреляция между общинным использованием санитарно-гигиенических удобств и распространенностью активной трахомы среди детей, что указывает на необходимость в непрекращающихся мерах по стимулированию более высокого уровня использования санитарно-гигиенических удобств и поддержке постоянного использования всей общиной, а не только на уровне домашнего хозяйства.

Resumen

Tracoma activo y uso comunitario del saneamiento, Etiopía

Objetivo Investigar, en Amhara, Etiopía, la relación entre la prevalencia del tracoma activo en niños de entre 1 y 9 años y el uso comunitario del saneamiento.

Métodos Entre 2011 y 2014, se midieron la prevalencia del tracoma y el uso doméstico de las letrinas de pozo en cinco encuestas transversales basadas en la población. Los datos sobre los indicadores observados del uso de letrinas se agregaron en una medida del uso del saneamiento comunitario calculado como el porcentaje de hogares con una letrina en uso. Se examinó a todos los miembros del hogar en busca de síntomas clínicos, es decir, inflamación tracomatosa, folicular y/o intensa, indicio de tracoma activo. Se utilizó una regresión logística de varios niveles para estimar la prevalencia de los cocientes de posibilidades (CP) y los intervalos de confianza (IC) del 95%, ajustándose a factores comunitarios, domésticos e individuales, y para evaluar la modificación por uso de letrinas y acceso al agua en los hogares.

Resultados En las zonas encuestadas, se estimó que la prevalencia del tracoma activo en niños era del 29% (IC del 95%: 28–30) y que el uso comunitario del saneamiento era del 47% (IC del 95%: 45–48). A pesar de una importante modificación ($p < 0,0001$), no se detectó ningún patrón en los CP estratificados. Como resumen de los estratos, los valores de uso comunitario del saneamiento del 60 al $< 80\%$ y $\geq 80\%$ se relacionaron con una menor prevalencia de tracoma activo, en comparación con un uso comunitario del saneamiento de $< 20\%$ (CP: 0,76; IC del 95%: 0,57–1,03 y CP: 0,67; IC del 95%: 0,48–0,95, respectivamente).

Conclusión En Amhara, Etiopía, se observó una correlación negativa entre el uso comunitario del saneamiento y la prevalencia del tracoma activo en niños, destacando la necesidad de unos esfuerzos continuos por fomentar unos mayores niveles de uso del saneamiento y para respaldar el uso constante en toda la comunidad, no únicamente en los hogares.

References

- Pascolini D, Mariotti SP. Global estimates of visual impairment: 2010. *Br J Ophthalmol*. 2012 May;96(5):614–8. doi: <http://dx.doi.org/10.1136/bjophthalmol-2011-300539> PMID: 22133988
- Mariotti SP, Pascolini D, Rose-Nussbaumer J. Trachoma: global magnitude of a preventable cause of blindness. *Br J Ophthalmol*. 2009 May;93(5):563–8. doi: <http://dx.doi.org/10.1136/bjo.2008.148494> PMID: 19098034
- WHO Alliance for the Global Elimination of Blinding Trachoma by the year 2020. Progress report on elimination of trachoma, 2013. *Wkly Epidemiol Rec*. 2014 Sep 26;89(39):421–8. PMID: 25275153
- Burton MJ, Mabey DC. The global burden of trachoma: a review. *PLoS Negl Trop Dis*. 2009 Oct 27;3(10):e460. doi: <http://dx.doi.org/10.1371/journal.pntd.0000460> PMID: 19859534
- Hu VH, Harding-Esch EM, Burton MJ, Bailey RL, Kadimpeul J, Mabey DC. Epidemiology and control of trachoma: systematic review. *Trop Med Int Health*. 2010 Jun;15(6):673–91. doi: <http://dx.doi.org/10.1111/j.1365-3156.2010.02521.x> PMID: 20374566
- Emerson PM, Cairncross S, Bailey RL, Mabey DC. Review of the evidence base for the 'F' and 'E' components of the SAFE strategy for trachoma control. *Trop Med Int Health*. 2000 Aug;5(8):515–27. doi: <http://dx.doi.org/10.1046/j.1365-3156.2000.00603.x> PMID: 10995092
- Ngondi J, Gebre T, Shargie EB, Adamu L, Ejigsemahu Y, Teferi T, et al. Evaluation of three years of the SAFE strategy (Surgery, Antibiotics, Facial cleanliness and Environmental improvement) for trachoma control in five districts of Ethiopia hyperendemic for trachoma. *Trans R Soc Trop Med Hyg*. 2009 Oct;103(10):1001–10. doi: <http://dx.doi.org/10.1016/j.trstmh.2008.11.023> PMID: 19178920
- Emerson PM, Ngondi J, Biru E, Graves PM, Ejigsemahu Y, Gebre T, et al. Integrating an NTD with one of "the big three": combined malaria and trachoma survey in Amhara Region of Ethiopia. *PLoS Negl Trop Dis*. 2008 Mar 19;2(3):e197. doi: <http://dx.doi.org/10.1371/journal.pntd.0000197> PMID: 18350115
- Berhane Y. Prevalence of trachoma in Ethiopia. *Ethiop J Health Dev*. 2007;21(3):211–5.
- Bailey R, Lietman T. The SAFE strategy for the elimination of trachoma by 2020: will it work? *Bull World Health Organ*. 2001;79(3):233–6. PMID: 11285668
- West SK. Blinding trachoma: prevention with the safe strategy. *Am J Trop Med Hyg*. 2003 Nov;69(5) Suppl:18–23. PMID: 14692676
- Emerson PM, Bailey RL, Mahdi OS, Walraven GE, Lindsay SW. Transmission ecology of the fly *Musca sorbens*, a putative vector of trachoma. *Trans R Soc Trop Med Hyg*. 2000 Jan-Feb;94(1):28–32. doi: [http://dx.doi.org/10.1016/S0035-9203\(00\)90427-9](http://dx.doi.org/10.1016/S0035-9203(00)90427-9) PMID: 10748893
- Emerson PM, Bailey RL, Walraven GE, Lindsay SW. Human and other faeces as breeding media of the trachoma vector *Musca sorbens*. *Med Vet Entomol*. 2001 Sep;15(3):314–20. doi: <http://dx.doi.org/10.1046/j.0269-283x.2001.00318.x> PMID: 11583450
- Emerson PM, Lindsay SW, Walraven GE, Faal H, Bøgh C, Lowe K, et al. Effect of fly control on trachoma and diarrhoea. *Lancet*. 1999 Apr 24;353(9162):1401–3. doi: [http://dx.doi.org/10.1016/S0140-6736\(98\)09158-2](http://dx.doi.org/10.1016/S0140-6736(98)09158-2) PMID: 10227221
- Montgomery MA, Desai MM, Elimelech M. Assessment of latrine use and quality and association with risk of trachoma in rural Tanzania. *Trans R Soc Trop Med Hyg*. 2010 Apr;104(4):283–9. doi: <http://dx.doi.org/10.1016/j.trstmh.2009.10.009> PMID: 19926106

16. Stocks ME, Ogden S, Haddad D, Addiss DG, McGuire C, Freeman MC. Effect of water, sanitation, and hygiene on the prevention of trachoma: a systematic review and meta-analysis. *PLoS Med*. 2014 Feb;11(2):e1001605. doi: <http://dx.doi.org/10.1371/journal.pmed.1001605> PMID: 24586120
17. Solomon AW, Zondervan M, Kuper H, Buchan JC, Mabey D, Foster A. *Trachoma control: a guide for programme managers*. Geneva: World Health Organization; 2006.
18. King JD, Endeshaw T, Escher E, Alemtaye G, Melaku S, Gelaye W, et al. Intestinal parasite prevalence in an area of Ethiopia after implementing the SAFE strategy, enhanced outreach services, and health extension program. *PLoS Negl Trop Dis*. 2013;7(6):e2223. doi: <http://dx.doi.org/10.1371/journal.pntd.0002223> PMID: 23755308
19. King JD, Teferi T, Cromwell EA, Zerihun M, Ngondi JM, Damte M, et al. Prevalence of trachoma at sub-district level in Ethiopia: determining when to stop mass azithromycin distribution. *PLoS Negl Trop Dis*. 2014 Mar;8(3):e2732. doi: <http://dx.doi.org/10.1371/journal.pntd.0002732> PMID: 24625539
20. Turner AG, Magnani RJ, Shuaib M. A not quite as quick but much cleaner alternative to the Expanded Programme on Immunization (EPI) cluster survey design. *Int J Epidemiol*. 1996 Feb;25(1):198–203. doi: <http://dx.doi.org/10.1093/ije/25.1.198> PMID: 8666490
21. *Monitoring the situation of children and women. multiple indicator cluster survey manual 2005*. New York: United Nations Children's Fund; 2006.
22. Thylefors B, Dawson CR, Jones BR, West SK, Taylor HR. A simple system for the assessment of trachoma and its complications. *Bull World Health Organ*. 1987;65(4):477–83. PMID: 3500800
23. King JD, Buolamwini J, Cromwell EA, Panfel A, Teferi T, Zerihun M, et al. A novel electronic data collection system for large-scale surveys of neglected tropical diseases. *PLoS One*. 2013;8(9):e74570. doi: <http://dx.doi.org/10.1371/journal.pone.0074570> PMID: 24066147
24. Ngondi J, Teferi T, Gebre T, Shargie EB, Zerihun M, Ayele B, et al. Effect of a community intervention with pit latrines in five districts of Amhara, Ethiopia. *Trop Med Int Health*. 2010 May;15(5):592–9. doi: <http://dx.doi.org/10.1111/j.1365-3156.2010.02500.x> PMID: 20345557
25. Diez Roux AV. A glossary for multilevel analysis. *J Epidemiol Community Health*. 2002 Aug;56(8):588–94. doi: <http://dx.doi.org/10.1136/jech.56.8.588> PMID: 12118049
26. Merlo J, Ohlsson H, Lynch KF, Chaix B, Subramanian SV. Individual and collective bodies: using measures of variance and association in contextual epidemiology. *J Epidemiol Community Health*. 2009 Dec;63(12):1043–8. doi: <http://dx.doi.org/10.1136/jech.2009.088310> PMID: 19666637
27. Carle AC. Fitting multilevel models in complex survey data with design weights: recommendations. *BMC Med Res Methodol*. 2009 Jul 14;9(1):49. doi: <http://dx.doi.org/10.1186/1471-2288-9-49> PMID: 19602263
28. Korn EL, Graubard BI. *Analysis of health surveys*. New York: John Wiley; 1999. doi: <http://dx.doi.org/10.1002/9781118032619>
29. Merlo J, Chaix B, Ohlsson H, Beckman A, Johnell K, Hjerpe P, et al. A brief conceptual tutorial of multilevel analysis in social epidemiology: using measures of clustering in multilevel logistic regression to investigate contextual phenomena. *J Epidemiol Community Health*. 2006 Apr;60(4):290–7. doi: <http://dx.doi.org/10.1136/jech.2004.029454> PMID: 16537344
30. Bright EA, Coleman PR, Rose AN, Urban ML. *LandScan 2011*. Oak Ridge: Oak Ridge National Laboratory; 2012. Available from: <http://www.ornl.gov/landscan/> [cited 2015 Apr 17].
31. Textor J, Hardt J, Knüppel S. DAGitty: a graphical tool for analyzing causal diagrams. *Epidemiology*. 2011 Sep;22(5):745. doi: <http://dx.doi.org/10.1097/EDE.0b013e318225c2be> PMID: 21811114
32. Rothman KJ, Greenland S, Lash TL. *Modern epidemiology*. 3rd ed. Philadelphia: Lippincott, Williams & Wilkins; 2008.
33. Rabe-Hesketh S, Skrondal A. *Multilevel and longitudinal modeling using Stata*. Third Edition, Volume II: Categorical responses, counts, and survival. College Station: Stata Press; 2012.
34. Kleinbaum DG, Kupper LL, Morgenstern H. *Epidemiologic research: principles and quantitative methods*. Belmont: Lifetime Learning Publications; 1982.
35. Neter J, Wasserman W, Kutner MH. *Applied linear statistical models: regression, analysis of variance and experimental designs*. Homewood: Richard D Irwin; 1985.
36. Emerson PM, Lindsay SW, Alexander N, Bah M, Dibba SM, Faal HB, et al. Role of flies and provision of latrines in trachoma control: cluster-randomised controlled trial. *Lancet*. 2004 Apr 03;363(9415):1093–8. doi: [http://dx.doi.org/10.1016/S0140-6736\(04\)15891-1](http://dx.doi.org/10.1016/S0140-6736(04)15891-1) PMID: 15064026
37. Stoller NE, Gebre T, Ayele B, Zerihun M, Assefa Y, Habte D, et al. Efficacy of latrine promotion on emergence of infection with ocular Chlamydia trachomatis after mass antibiotic treatment: a cluster-randomized trial. *Int Health*. 2011 Jun;3(2):75–84. doi: <http://dx.doi.org/10.1016/j.inhe.2011.03.004> PMID: 21785663
38. Haile M, Tadesse Z, Gebreselassie S, Ayele B, Gebre T, Yu SN, et al. The association between latrine use and trachoma: a secondary cohort analysis from a randomized clinical trial. *Am J Trop Med Hyg*. 2013 Oct;89(4):717–20. doi: <http://dx.doi.org/10.4269/ajtmh.13-0299> PMID: 24002488
39. Keenan JD, Lakew T, Alemayehu W, Melese M, House JI, Acharya NR, et al. Slow resolution of clinically active trachoma following successful mass antibiotic treatments. *Arch Ophthalmol*. 2011 Apr;129(4):512–3. doi: <http://dx.doi.org/10.1001/archophthalmol.2011.46> PMID: 21482879
40. Freeman MC, Ogden S, Jacobson J, Abbott D, Addiss DG, Amnie AG, et al. Integration of water, sanitation, and hygiene for the prevention and control of neglected tropical diseases: a rationale for inter-sectoral collaboration. *PLoS Negl Trop Dis*. 2013;7(9):e2439. doi: <http://dx.doi.org/10.1371/journal.pntd.0002439> PMID: 24086781
41. Barreto ML, Genser B, Strina A, Teixeira MG, Assis AM, Rego RF, et al. Effect of city-wide sanitation programme on reduction in rate of childhood diarrhoea in northeast Brazil: assessment by two cohort studies. *Lancet*. 2007 Nov 10;370(9599):1622–8. doi: [http://dx.doi.org/10.1016/S0140-6736\(07\)61638-9](http://dx.doi.org/10.1016/S0140-6736(07)61638-9) PMID: 17993362
42. Bateman OS, Smith S. *A comparison of the health effects of water supply and sanitation in urban and rural Guatemala*. Washington: United States Agency for International Development; 1991.
43. Root GP. Sanitation, community environments, and childhood diarrhoea in rural Zimbabwe. *J Health Popul Nutr*. 2001 Jun;19(2):73–82. PMID: 11503350
44. Fuller JA, Villamor E, Cevallos W, Trostle J, Eisenberg JN. I get height with a little help from my friends: herd protection from sanitation on child growth in rural Ecuador. *Int J Epidemiol*. 2016 Apr;45(2):460–9. doi: <http://dx.doi.org/10.1093/ije/dyv368> PMID: 26936912
45. *Estimates on the use of water sources and sanitation facilities*. New York: United Nations Children's Fund; 2014.
46. Kar K, Chambers R. *Handbook on community-led total sanitation*. London: Plan UK; 2008.