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DOI: 10.1371/journal.pone.0063840

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An Evaluation of Culture Results during Treatment for Tuberculosis as Surrogate Endpoints for Treatment Failure and Relapse

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Abstract

It is widely acknowledged that new regimens are urgently needed for the treatment of tuberculosis. The primary endpoint in the Phase III trials is a composite outcome of failure at the end of treatment or relapse after stopping treatment. Such trials are usually both long and expensive. Valid surrogate endpoints measured during or at the end of treatment could dramatically reduce both the time and cost of assessing the effectiveness of new regimens. The objective of this study was to evaluate sputum culture results on solid media during treatment as surrogate endpoints for poor outcome. Data were obtained from twelve randomised controlled trials conducted by the British Medical Research Council in the 1970s and 80s in East Africa and East Asia, consisting of 6974 participants and 49 different treatment regimens. The month two culture result was shown to be a poor surrogate in East Africa but a good surrogate in Hong Kong. In contrast, the month three culture was a good surrogate in trials conducted in East Africa but not in Hong Kong. As well as differences in location, ethnicity and probable strain of Mycobacteria tuberculosis, Hong Kong trials more often evaluated regimens with rifampicin throughout and intermittent regimens, and patients in East African trials more often presented with extensive cavitation and were slower to convert to culture negative during treatment. An endpoint that is a summary measure of the longitudinal profile of culture results over time or that is able to detect the presence of M. tuberculosis later in treatment is more likely to be a better endpoint for a phase II trial than a culture result at a single time point and may prove to be an acceptable surrogate. More data are needed before any endpoint can be used as a surrogate in a confirmatory phase III trial.

Introduction

Tuberculosis (TB) is one of the world’s oldest infectious disease and over the centuries has been responsible for more mortality, morbidity and human suffering than any other[1]. Though an effective cure is available today for no more than US $20, it is estimated that there are over 9 million new cases and almost 2 million people die every year from TB[2]. The six month standard regimen for drug-susceptible TB has been shown to be highly efficacious in clinical trials[3], but such results are rarely achieved in practice[4]. New treatment regimens are urgently needed to reduce the duration of treatment for drug-susceptible TB and effectively treat multi-drug resistant TB (MDR-TB).

TB is almost unique among bacterial infections in that failure to culture the bacilli is not necessarily indicative of cure. An effective regimen is one which renders patients culture negative by the end of treatment, but also prevents subsequent relapse. Clinical trials to evaluate new regimens for the treatment of TB therefore commonly involve follow-up of Sputum culture results throughout the end of treatment of 18-24 months[5]. A surrogate endpoint measured during or at the end of treatment could be used as a substitute for the currently used composite clinical endpoint of treatment failure and relapse in a phase III clinical trial[5] thereby substantially shortening the trial duration and speeding overall drug development.

A biomarker is any marker objectively measured and evaluated as an indicator of normal biological processes, pathogenic processes, or pharmacologic responses to a therapeutic intervention[6]. A surrogate endpoint is a biomarker that fully captures the effect of the trial intervention on the clinical endpoint and ‘for which a test of the null hypothesis of no relationship to the treatment effect is possible’[7]. Such an endpoint can only be defined in the context of a comparison of different treatment arms in a randomised controlled trial[8]. There have been several reviews on surrogate endpoints and biomarkers of TB treatment response[9–12]. Perrin et al.[10] summarised the potential markers currently available that could describe a patient’s response to treatment (biomarkers of treatment response) but no formal evaluation of surrogate endpoints was provided. The two month culture was shown to be a risk factor for relapse in two trials[13,14] but there was, however, no attempt to formally evaluate the endpoint as a surrogate endpoint.


Editor: Adithya Cattamanchi, San Francisco General Hospital, University of California San Francisco, United States of America

Received May 1, 2012; Accepted April 11, 2013; Published May 8, 2013

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Funding: This work was funded by a British Medical Research Council PhD studentship. The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Competing Interests: The authors have declared that no competing interests exist.

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Perrin et al. also referred to a review by Mitchison[15] expanding on earlier correspondence[16] which addressed the relationship between relapse rates and rates of culture positivity across several clinical trials. It is not clear which statistical methods were used; the results indicate a relationship suggestive of surrogacy, but with a note of caution that ‘the most effective time for measuring sputum conversion may vary according to the drug under test.’[15] The report from an expert consultation meeting on biomarkers in TB organised by WHO Tropical Disease Research (TDR) also identified the two month culture result as being currently accepted as a surrogate for treatment outcome, but again without any formal evaluation to support this[17]. A recent systematic review of sputum monitoring during TB treatment for predicting outcome found the two month culture had modest specificity but low sensitivity for predicting failure and relapse[18].

The objective of the present study was to evaluate sputum culture results during treatment as potential surrogate endpoints for long term outcome in the treatment of pulmonary TB using appropriate statistical methodology. Treatment comparisons are required for evaluating a surrogate endpoint and therefore it is necessary to use data from randomised controlled clinical trials where culture results during treatment are available as well as follow up for relapse for a minimum of 18 months. To achieve this objective, data were used from selected TB clinical trials conducted by the British Medical Research Council (BMRC) during the 1970s and 1980s.

Methods

Ethics Statement

This study of past clinical trial data was approved by the ethics committee of the London School of Hygiene and Tropical Medicine.

Selection of studies

Individual patient data were available from all TB clinical trials that were conducted by the BMRC in East Africa and East Asia[3]. These trials provided much of the evidence for the short course regimens which are standard treatment today[3,19,20]. The advantage of using data only from BMRC trials is that the clinical and bacteriological protocols were largely unchanged throughout the programme of trials and the level of homogeneity was therefore high. These data are therefore ideal for the evaluation of culture results during treatment as surrogate endpoints.

These large multi-centre randomised controlled trials included high quality laboratory data with frequent follow-up sampling after the end of treatment on smear positive patients with pulmonary TB. All trials included regimens comprised of various combinations of first-line drugs available and recommended for use today[2]: isoniazid, rifampicin, ethambutol and pyrazinamide, in addition to two drugs no longer used in first-line regimens: thiacetazone and streptomycin. From this pool of trials, treatment arms of duration other than six months (the duration of the WHO-recommended regimen in use today) have been excluded as were trials of regimens all with less than 2% relapses as the small numbers of relapses yield too little information for evaluating surrogates. The full list of treatment arms included is given in Table 1.

Clinical and Surrogate Endpoint Definitions

A treatment failure was defined as heavy growth on culture (at least 20 colonies) at month 3 or 6 and a relapse defined as two cultures with heavy growth within three consecutive months or three positive cultures with any growth (one colony or more) within four consecutive months, following the end of treatment. The clinical endpoint was a combined endpoint of treatment failure at the end of treatment or relapse in follow-up hereafter referred to as poor outcome. In the original publications, these were usually presented as two separate endpoints, but were combined in this study in a composite endpoint to reflect the endpoint currently used in phase III TB trials[5,21,22]. The bacteriological definitions of treatment failure and relapse were taken from with the original trial reports[3], with the exception that ‘heavy growth’ was sometimes defined as at least 5, 10 or 20 colonies. Heavy growth of at least 20 colonies was chosen for consistency across trials. Default or death from a non-TB cause during treatment were classified as a missing clinical endpoint. If a patient was lost to follow-up after a single positive culture, they were classified as a relapse if no further data were available.

Cultures on solid media during treatment were available monthly and were recorded on a semi-categorical scale: negative, 0–19 colonies, 20–100 colonies, more than 100 colonies, or confluent growth. Three endpoints were evaluated as potential surrogates: i) a positive culture of at least 20 colonies at month one, ii) a positive culture of any growth at month two and iii) a positive culture of any growth at month three. Most patients were still culture positive after one month of treatment, so the endpoint selected for this time point is a positive culture of at least 20 colonies. Insufficient patients were culture positive at month four for that to be a useful endpoint.

Statistical Methods

Patients identified as having additional extra-pulmonary TB were often withdrawn from the trials and so data on treatment outcomes were often missing. For this reason, this small number of patients were excluded from the analysis. Patients with negative cultures at baseline were also excluded.

Culture results during treatment were evaluated as surrogate endpoints for poor outcome using a two stage approach based on a frequentist application of the Bayesian methods developed to evaluate CD4 count as a potential surrogate the development of AIDS or death[23]. Both stages are repeated for each of the three candidate surrogates. The first stage involves analysis at the trial participant-level estimating the treatment effect on the surrogate endpoint $\alpha_i$ (expressed as the log odds ratio of a positive culture) and the treatment effect on the clinical endpoint $\beta_j$ (expressed as the log odds ratio of a poor outcome) for each treatment comparison $j$, of an experimental arm with the control arm in each trial $i$. Many of the trials did not have a pre-specified ‘control regimen’ and so, unless the control arm was obvious, the arm with the highest proportion of poor outcomes was identified as the nominal control. This was done so that the difference in risk of poor outcome between the experimental and control regimens is greatest and therefore the treatment ordering is such that the most amount of information is available for evaluating culture results as surrogate endpoints. Where two or more arms in a trial had the same treatment for the first one or two months, the control arm was selected to limit the number of comparisons of such arms.

The second stage involves analysis at the treatment-comparison-level fitting a linear regression model with $\beta_j$ as the response and $\alpha_i$ as the explanatory variable. Since the pairs $(\alpha_i, \beta_j)$ are estimated with varying precision, the model is fitted with weights equal to the inverse of the mean of the variances of the $\alpha_i$ and $\beta_j$ for each $i$ and $j$. Robust standard errors are used to account for the clustering of treatment comparisons within trials and the intercept term in the linear model is constrained to be zero since each treatment comparison corresponds to comparisons of different
Table 1. List of trials and treatment arms included in this study.

<table>
<thead>
<tr>
<th>Trial</th>
<th>Year of Start</th>
<th>Treatment arms included (first regimen nominated as control)</th>
</tr>
</thead>
<tbody>
<tr>
<td>East Africa</td>
<td></td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>1970</td>
<td>6SH, 6SHT, 6SHZ, 6SHR</td>
</tr>
<tr>
<td>2</td>
<td>1972</td>
<td>6HR, 25HRZ/4TH, 25HRZ/4SHZ, 6SHR</td>
</tr>
<tr>
<td>3</td>
<td>1974</td>
<td>25HRZ/4TH, 15HRZ/25TH, 25HRZ/4TH, 15HRZ/5SHZ, 6SHR</td>
</tr>
<tr>
<td>6</td>
<td>1978</td>
<td>25HRZ/4H, 25HRZ/4TH</td>
</tr>
<tr>
<td>Hong Kong</td>
<td></td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>1972</td>
<td>6SHZ, 6SH, 6SHZ</td>
</tr>
<tr>
<td>2</td>
<td>1974</td>
<td>25SHRE/4SHZE, 25SHRE/4SHZ, 45SHRE/25SHZ</td>
</tr>
<tr>
<td>3</td>
<td>1977</td>
<td>6HRSE, 6HRZE, 6HRZ, 6HRZ, 6HRZ</td>
</tr>
<tr>
<td>4</td>
<td>1979</td>
<td>2HRZ/4HRE, 6SHRE(2wZ1), 6SHRE(4wZ1), 6SHRE(8wZ1), 6SHRE(2wZ2), 6SHRE(8wZ1), 6SHRE(4wZ1)</td>
</tr>
<tr>
<td>Singapore</td>
<td></td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>1973</td>
<td>25HRZ/4HR, 25HRZ/4HRZ</td>
</tr>
<tr>
<td>3</td>
<td>1983</td>
<td>25HRZ/4HR, 20HRZ/4HR, 15HRZ/4HR, 15HRZ/SHR, 2HRZ/4HR, 25HRZ/4HR</td>
</tr>
</tbody>
</table>

Trial numbering corresponds to numbering in a comprehensive review of all MRC studies[3], where full references for trial report(s) are listed. Trials conducted in East and Central Africa are listed in Table 1.7, trials conducted in Hong Kong in Table 1.8 and trials conducted in Singapore in Table 1.9 of the review[3].

1This trial was actually of 4 month regimens, but was terminated earlier than planned and patients still on treatment at that time were continued to 6 months of treatment. The results of those on 6 months of treatment were presented in a later publication[44].

2The results of this trial were never published (personal communication, DA Mitchison) and it is therefore not included in the tables of MRC studies in Fox, Ellard and Mitchison[3]. For treatment notation: S = Streptomycin, H = Isoniazid, T = Thiacetazone, Z = Pyrazinamide, R = Rifampicin, E = Ethambutol. Where the regimen has distinct intensive and continuation phases, these are separated by a forward slash with the leading number corresponding to the duration in months. The subscript indicates the number of doses given weekly; the absence of subscript indicates daily dosing. For example, 25SHRE/4SHZE consists of a 2 month intensive phase of 4 drugs given daily followed by a 4 month continuation phase of 3 drugs each given twice-weekly. The subscript C indicates the drugs were given in a combined formulation. In the fourth Hong Kong study pyrazinamide given once or thrice weekly was added to some of the regimens for the first 2, 4 or 8 weeks. This is indicated by the text in parentheses where, for example, 4wZ1 indicates that once-weekly pyrazinamide was added only for the first 4 weeks.

doi:10.1371/journal.pone.0063840.t001

treatment regimens and therefore a non-zero intercept has no meaning. The treatment effect on poor outcome is also plotted against the treatment effect on the surrogate with the diameter of the circles corresponding to the precision of the estimates. Estimates with greater precision, and therefore larger weight in the linear model, are represented by larger circles. The proportion of variation in β̂ explained by η̂ in this situation is called the trial-level proportion of variation explained, R²trial. This is an established metric for evaluating surrogate endpoints[24], and based on a number of examples of the use of this metric, an R²trial ≥ 0.80 could be considered as evidence for a surrogate being ‘good’ and R²trial ≥ 0.95 being ‘very good’[24]. The analyses were repeated incorporating adjustment for important baseline patient risk factors (including smear and culture status, pre-treatment drug resistance, extent of cavitation, weight, age and sex) in the first stage of the two-stage analysis.

Results
Baseline Characteristics
Table 2 summarises the baseline characteristics of the trials and patients included in this study. Data were included from 12 trials, yielding 49 trial arms and 37 total possible treatment comparisons. Relapse rates and details of the treatment arms can be found elsewhere[3] along with references of the individual trial reports. Patients in the Hong Kong trials were more likely to have pre-treatment resistance to isoniazid (9% overall) or streptomycin (10%) than in the East African trials, 7% isoniazid resistance and 3% streptomycin resistance. Patients in the East African trials were much more likely to present with extensive or gross cavitation (44% overall) than in the Hong Kong trials (5% overall).

Evaluating the candidate surrogates
Treatment for TB is usually described in two phases, the intensive phase where three to four drugs are given together followed by the continuation phase where, typically, only two of these drugs are continued until the end of treatment. Some treatment comparisons involved regimens with the same drug combination in the first few months of treatment. These comparisons were therefore excluded from the evaluation of the one and two month cultures. In summary, of the 37 total possible treatment comparisons from 49 trial arms, 32 treatment comparisons were used to evaluate a positive culture of heavy growth at month one, 33 to evaluate the month two culture result and 35 the month three culture result, as surrogate endpoints. All treatment comparisons involved substantial changes to the regimen in the first few months.

Figure 1 and the first three rows of Table 3 show the results of the second stage of the analysis for all trials overall for the three potential surrogate endpoints.

There is considerable scatter about the fitted line in Figures 1A and 1B with the proportions of variation explained, R²trial only 0.36 in each case. This can be interpreted as a weak relationship between the treatment comparison on the clinical endpoint of a poor outcome and the treatment comparison on the candidate surrogate endpoint meaning that it would not be possible to use the effect of a treatment regimen on the candidate surrogate to accurately predict the effect of the treatment on the proportion of poor outcomes.

Apart from the groupings around the origin in figure 1B (showing no difference between treatment on either endpoint), there are at least seven points in the lower right quadrants indicating that the treatment direction on a poor outcome is opposite to that on the candidate surrogate. The corresponding treatments would appear to be inferior to the control when
Table 2. Summary of baseline characteristics of trial participants included in the analysis.

<table>
<thead>
<tr>
<th>Trials</th>
<th>Participants</th>
<th>Treatment arms</th>
<th>Male N (%)</th>
<th>Age Median (IQR)</th>
<th>Weight/kg Median (IQR)</th>
<th>Isoniazid N (%)</th>
<th>Streptomycin N (%)</th>
<th>Culture</th>
<th>Smear</th>
<th>Extent of cavitation</th>
<th>Extent of disease</th>
</tr>
</thead>
<tbody>
<tr>
<td>East Africa 1</td>
<td>761</td>
<td>4</td>
<td>520 (68%)</td>
<td>30 (24–41)</td>
<td>49 (43–54)</td>
<td>65 (9%)</td>
<td>16 (2%)</td>
<td>733 (98%)</td>
<td>279 (37%)</td>
<td>377 (59%)</td>
<td>326 (50%)</td>
</tr>
<tr>
<td>2</td>
<td>902</td>
<td>4</td>
<td>566 (63%)</td>
<td>32 (25–44)</td>
<td>48 (44–53)</td>
<td>55 (8%)</td>
<td>36 (4%)</td>
<td>854 (96%)</td>
<td>217 (24%)</td>
<td>468 (54%)</td>
<td>454 (52%)</td>
</tr>
<tr>
<td>3</td>
<td>421</td>
<td>4</td>
<td>268 (64%)</td>
<td>30 (25–40)</td>
<td>47 (42–52)</td>
<td>28 (7%)</td>
<td>17 (4%)</td>
<td>406 (96%)</td>
<td>157 (37%)</td>
<td>242 (59%)</td>
<td>200 (49%)</td>
</tr>
<tr>
<td>4</td>
<td>310</td>
<td>5</td>
<td>197 (64%)</td>
<td>34 (25–44)</td>
<td>48 (43–51)</td>
<td>25 (8%)</td>
<td>4 (1%)</td>
<td>299 (97%)</td>
<td>96 (31%)</td>
<td>48 (17%)</td>
<td>99 (35%)</td>
</tr>
<tr>
<td>5</td>
<td>533</td>
<td>3</td>
<td>339 (64%)</td>
<td>30 (23–40)</td>
<td>47 (42–52)</td>
<td>28 (5%)</td>
<td>9 (2%)</td>
<td>509 (97%)</td>
<td>241 (46%)</td>
<td>56 (12%)</td>
<td>200 (45%)</td>
</tr>
<tr>
<td>6</td>
<td>296</td>
<td>2</td>
<td>208 (71%)</td>
<td>34 (25–45)</td>
<td>49 (43–54)</td>
<td>24 (8%)</td>
<td>11 (4%)</td>
<td>264 (91%)</td>
<td>47 (16%)</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td>Hong Kong 1</td>
<td>246</td>
<td>3</td>
<td>175 (71%)</td>
<td>37 (23–50)</td>
<td>46 (41–51)</td>
<td>27 (11%)</td>
<td>33 (13%)</td>
<td>220 (89%)</td>
<td>129 (52%)</td>
<td>28 (11%)</td>
<td>51 (21%)</td>
</tr>
<tr>
<td>2</td>
<td>369</td>
<td>3</td>
<td>288 (78%)</td>
<td>36 (21–55)</td>
<td>47 (43–52)</td>
<td>46 (12%)</td>
<td>48 (13%)</td>
<td>343 (93%)</td>
<td>125 (34%)</td>
<td>48 (13%)</td>
<td>67 (18%)</td>
</tr>
<tr>
<td>3</td>
<td>1142</td>
<td>5</td>
<td>824 (72%)</td>
<td>31 (22–52)</td>
<td>N/A</td>
<td>89 (8%)</td>
<td>104 (9%)</td>
<td>1097 (96%)</td>
<td>298 (26%)</td>
<td>18 (2%)</td>
<td>126 (11%)</td>
</tr>
<tr>
<td>4*</td>
<td>1489</td>
<td>8</td>
<td>1061 (71%)</td>
<td>32 (23–50)</td>
<td>48 (44–53)</td>
<td>134 (9%)</td>
<td>147 (10%)</td>
<td>1428 (97%)</td>
<td>503 (34%)</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td>Singapore 1</td>
<td>198</td>
<td>2</td>
<td>129 (65%)</td>
<td>43 (29–54)</td>
<td>45 (41–50)</td>
<td>4 (2%)</td>
<td>9 (5%)</td>
<td>192 (97%)</td>
<td>123 (62%)</td>
<td>4 (2%)</td>
<td>48 (24%)</td>
</tr>
<tr>
<td>3</td>
<td>307</td>
<td>6</td>
<td>199 (65%)</td>
<td>38 (25–51)</td>
<td>49 (42–55)</td>
<td>6 (2%)</td>
<td>9 (3%)</td>
<td>302 (98%)</td>
<td>161 (52%)</td>
<td>N/A</td>
<td>N/A</td>
</tr>
<tr>
<td>Total</td>
<td>6974</td>
<td>49</td>
<td>4774 (69%)</td>
<td>33 (23–47)</td>
<td>48 (43–53)</td>
<td>531 (8%)</td>
<td>443 (6%)</td>
<td>6647 (96%)</td>
<td>2376 (34%)</td>
<td>1289 (28%)</td>
<td>1591 (34%)</td>
</tr>
</tbody>
</table>

Some data elements were not available (N/A) in certain trials. Trial numbering corresponds to numbering in a comprehensive review of all MRC studies[3], where full references for trial report(s) are listed. Trials conducted in East and Central Africa are listed in Table 1.7, trials conducted in Hong Kong in Table 1.8 and trials conducted in Singapore in Table 1.9 of the review[3].

*The results of this trial were never published (personal communication, DA Mitchison) and it is therefore not included in the tables of MRC studies in Fox, Ellard and Mitchison[3].

IQR - Inter-Quartile Range.

doi:10.1371/journal.pone.0063840.t002
Analysis by geographical region

Patient-level adjustment for baseline factors

The analyses were repeated incorporating adjustment in the first stage of the analysis for important baseline patient risk factors (including smear and culture status, pre-treatment drug resistance, extent of cavitation, weight, age and sex) with no substantial differences in the result (data not shown).

Analysis by geographical region

Most of the trials were conducted across two separate geographical regions: East Africa and Hong Kong. There is evidence of different relapses rates between clinical trials conducted in these two regions using the same treatment regimen on three separate occasions[3]. Therefore, the three candidate surrogate endpoints were evaluated in trials separately for Hong Kong and East Africa. The results are shown in Figures 2 and 3 and Table 3. Too few treatment comparisons (all from one trial) were available from trials in Singapore to draw any clear conclusions.

At months one and two, analysis restricted to data from East Africa shows great variation about the line, with the proportion of variation explained 0.29 and 0.19, respectively, with very wide confidence intervals. This contrasts with analysis of the month three culture in East Africa, where there is a clear linear trend (Figure 2B, proportion of variation explained 0.81).

Six points lie in the lower right quadrant in Figure 2A, indicating an effect of treatment on the two month culture result that is in the opposite direction to the effect on the proportion of poor outcomes. Of these, two correspond to the addition of rifampicin beyond two months, two to the addition of pyrazinamide beyond two months and two to little or no change beyond 2 months.

In contrast to the East African graphs, a linear trend is more apparent for all months in the graphs restricted to data from Hong Kong only. At months one and three the proportions of explained variation are reasonably high at 0.69 and 0.62 respectively though the widths of the confidence intervals are also very wide. The best fit is at month two with a narrow 95% confidence interval around the slope, a high proportion of explained variation at 0.86, and no points outside the lower left quadrant, except for three which are very close to the origin (Figure 3A).

Fitting one model allowing for different slopes for trials from Hong Kong and East Africa, there was no evidence for interaction with p = 0.46 and p = 0.75 for months one and two, respectively. At month three, there was evidence for a difference in slopes between trials from Hong Kong and trials from East Africa, p = 0.015.

Table 4 shows a summary of some of the differences by geographical region. 73% of the treatment comparisons in the Hong Kong trials were of two regimens that included rifampicin throughout compared to only 6% of the treatment comparisons in the East African trials. Similarly, all of the regimens evaluated in the East African trials had daily dosing in the intensive phase of treatment compared to only 7% in the Hong Kong trials and 88% in the continuation phase in East African trials compared to none in the Hong Kong trials.

Culture conversion occurred earlier on average in Hong Kong compared to East Africa. 15% of patients at two months and only 4% of patients at three months were still culture positive in Hong Kong compared with 29% and 10% respectively in East Africa.

Repeating the analyses for comparisons of two regimens that contained rifampicin throughout treatment resulted in 0.67 and 0.46 proportion of variation explained for the 2 month and 3 month cultures respectively.

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Table 3. Estimates of the slope of the fitted line and of the proportion of explained variation from the model from the second stage.

<table>
<thead>
<tr>
<th>Analysis</th>
<th>Month of Culture*</th>
<th>Trials</th>
<th>Treatment Comparisonsb</th>
<th>Slope (95% CI)</th>
<th>$R^2_{\text{total}}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Overall</td>
<td>1</td>
<td>9</td>
<td>32</td>
<td>1.35 (0.10, 2.80)</td>
<td>0.36</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>9</td>
<td>33</td>
<td>0.85 (0.13, 1.57)</td>
<td>0.36</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>11</td>
<td>35</td>
<td>1.29 (0.82, 1.76)</td>
<td>0.69</td>
</tr>
<tr>
<td>East Africa trials only</td>
<td>1</td>
<td>4</td>
<td>13</td>
<td>1.13 (-1.82, 4.11)</td>
<td>0.29</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>4</td>
<td>13</td>
<td>0.76 (-1.57, 3.09)</td>
<td>0.19</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>6</td>
<td>16</td>
<td>1.61 (1.38, 1.83)</td>
<td>0.81</td>
</tr>
<tr>
<td>Hong Kong trials only</td>
<td>1</td>
<td>4</td>
<td>15</td>
<td>1.98 (-0.92, 4.05)</td>
<td>0.68</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>4</td>
<td>15</td>
<td>0.99 (0.82, 1.16)</td>
<td>0.86</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>4</td>
<td>15</td>
<td>0.82 (0.09, 1.56)</td>
<td>0.62</td>
</tr>
<tr>
<td>Singapore trials only</td>
<td>1</td>
<td>1</td>
<td>4</td>
<td>0.19 (-4.95, 5.33)</td>
<td>&lt;0.01</td>
</tr>
<tr>
<td></td>
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<td>1</td>
<td>5</td>
<td>-0.06 (-2.31, 2.20)</td>
<td>&lt;0.01</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>1</td>
<td>4</td>
<td>-0.52 (-5.09, 4.05)</td>
<td>0.04</td>
</tr>
</tbody>
</table>

CI - Confidence Interval; $R^2_{\text{total}}$ - trial-level proportion of variation.

*The month 1 endpoint evaluated was a positive culture with heavy growth (at least 20 colonies). The month 2 and 3 endpoints evaluated were a positive culture with any growth (at least 1 colony).

bDue to similarities in regimens in the first few months of treatment, not all 37 treatment comparisons could be used for each analysis.

doi:10.1371/journal.pone.0063840.t003
Discussion

The two month culture has variously been described as ‘the main surrogate marker [for sterilizing activity]’[25], a ‘currently available surrogate marker of relapse rates’[26], ‘probably the best available surrogate marker for the relapse rate’[27] and ‘an index of efficacy of anti-TB regimens’[28]. These conclusions are based on limited published evidence and a varied understanding of what is meant by a surrogate.

Using the definitions of surrogacy outlined in the introduction, this study has shown that the two month culture appeared to be a good surrogate endpoint using data from the Hong Kong trials and the three month culture was suggestive of a good surrogate endpoint using data from the East African trials, but the reverse was not true.

There were no substantial differences in the results on adjustment for patient-level baseline factors in the first stage of the analysis. This was not unexpected as each treatment comparison was a comparison of randomised groups which would be likely to be reasonably balanced by any baseline risk factors due to the process of randomisation.

On the one hand, the results are encouraging as they suggest that culture results over the first few months of treatment can be an acceptable surrogate endpoint in certain trials in certain populations. However, it is unclear which populations this would apply to, as the heterogeneity in results could be due to a number of factors, not just the geographical location of trial sites.

More of the treatment comparisons in the Hong Kong trials were of two regimens that included rifampicin throughout and many regimens included intermittent dosing, even in the intensive phase of treatment. Culture conversion was also earlier on average in the Hong Kong trials as compared to the East African trials as has been noted elsewhere[3], this was probably due at least in part to the greater use of rifampicin in the Hong Kong trials. Delayed culture conversion has however been reported in African patients compared to non-African patients, albeit in liquid media, in a recent multi-site clinical trial[29]. The authors concluded the differences could be due to ‘modest variation in laboratory processes’ but that further investigation was needed to find other possible causes[29]. Studies have also shown cavitation to be a strong risk factor for relapse[3,14]; and differences have been seen in relapse rates[30] and culture conversion after eight weeks of treatment[31] by strain and lineage of M. tuberculosis, with a corresponding association between strain and ethnicity or geographical location.

Based on these results, a surrogate endpoint that is a summary measure of the longitudinal profile of culture results over time is likely to be more useful than a culture result at a single time point.

Two approaches for capturing the longitudinal profile of culture results that have been proposed are (i) using a parameter from repeated measures modelling of culture results over time or (ii) summarising the time to stable culture conversion in a survival analysis. These approaches have been described elsewhere[32]. Both have been used in phase II TB clinical trials that are completed[33,34] and ongoing (the former in the TB Alliance study NC-002, clinicaltrials.gov identifier NCT01498419 and the latter in PanACEA MAMS-TB, Pan African Clinical Trials Registry identifier PACTR201205000383208). There is, however, no evidence as yet that either approach will yield markers that are acceptable surrogate endpoints.

Cultures were only performed monthly in the BMRC trials and therefore neither repeated measures modelling of culture results over time or summarising the time to stable culture conversion can be conducted with these data. Data from multiple treatment comparisons across large

Figure 1. Analysis of culture results as surrogate endpoints across all trials. A. Month 1 (a positive culture with heavy growth, at least 20 colonies), $R^2_{\text{trial}}=0.36$. B. Month 2 (a positive culture with any growth), $R^2_{\text{trial}}=0.69$. C. Month 3 (a positive culture with any growth), $R^2_{\text{trial}}=0.69$. Logs odds ratio of a poor outcome plotted against log odds ratio of a positive culture. Fitted line is weighted by the precision of the estimates, and this precision is represented by the diameter of the circles around each point. The dotted line represents the 95% confidence interval on the slope.
doi:10.1371/journal.pone.0063840.g001
multi-centre trials will be necessary for a formal evaluation of a marker as a surrogate endpoint. Some data will become available in the next few years as several large phase III clinical trials will be finishing and reporting results. Adaptive trial designs and innovative clinical development pathways are critical to compensate for the current lack of suitable surrogate endpoints[35].

In the only other formal evaluation of any marker as a surrogate endpoint for treatment response in TB[36], the authors evaluated two month culture conversion as a surrogate endpoint using data from published report of trials conducted by the BMRC. They selected 30 pairs of regimens showing that the slope of the meta-regression line ‘was statistically significant (p < 0.00001)’. There is some spread around the fitted line, but the authors do not give a figure for the $R^2_{trial}$ to allow the reader to judge whether this analysis shows two month culture conversion to be an acceptable surrogate. The authors concluded that two month culture conversion ‘should be a surrogate endpoint for the registration of new drugs for the treatment of TB.’ They were, however, unable to evaluate culture status at any other time as a surrogate, as it was only common in trial reports to publish the two month culture conversion rates, and there was no evidence that they had looked at effect modification by geographical region.

Figure 2. Sub-group analysis by geographical region: East African trials. A. Month 2 restricted to East African trials, $R^2_{trial} = 0.19$. B. Month 3 restricted to East African trials, $R^2_{trial} = 0.81$. Logs odds ratio of a poor outcome plotted against log odds ratio of a positive culture. Fitted line is weighted by the precision of the estimates, and this precision is represented by the diameter of the circles around each point. The dotted line represents the 95% confidence interval on the slope. doi:10.1371/journal.pone.0063840.g002
Limitations

Apart from the limitation of the cultures only being available monthly, the cultures were done on solid media which is being used increasingly less in clinical trials. The results may well be different when liquid media are used due to the increased diagnostic sensitivity[21].

These trials were conducted before methodology was available to distinguish true endogenous relapse from exogenous reinfection caused by a new strain of *M. tuberculosis*. It is therefore possible that a number of the cases recorded as relapses were in fact as a result of reinfection. However, this number is likely to be few for two reasons. Firstly reinfections occur more frequently in HIV co-infected patients[37,38] and these trials were conducted before the HIV epidemic. Secondly, and more importantly, many of these trials had up to 5 years of follow-up finding very few recurrences in the final 2–3 years of follow-up. Results at 5 years were consistent with those after 30 months suggesting minimal impact of the inclusion of possible cases of exogenous reinfection[3].

The two-stage analysis methodology did have some drawbacks. The estimates of the $a_{ij}$ were assumed in the second stage to be without error with the variance of the estimates only entering the model through the weights. The $a_{ij}$ and $b_{ij}$ are estimated separately in the first stage and the correlation between the two

**Figure 3. Sub-group analysis by geographical region: Hong Kong trials.**

A. Month 2 restricted to Hong Kong trials, $R^2_{final} = 0.86$. B. Month 3 restricted to Hong Kong trials, $R^2_{final} = 0.62$. Logs odds ratio of a poor outcome plotted against log odds ratio of a positive culture. Fitted line is weighted by the precision of the estimates, and this precision is represented by the diameter of the circles around each point. The dotted line represents the 95% confidence interval on the slope.

doi:10.1371/journal.pone.0063840.g003
is therefore assumed to be zero, which is probably not the case since
the estimates are from the same group of trial participants.
This will result in estimates of $R^2_{\text{trial}}$ slightly above or below the
true values, but the impact is likely to be minimal.

All comparisons were of two treatments with differences in the
first few months of treatment. Many comparisons also involved
changes in regimens after the putative surrogate endpoint had
been measured. Of the six points in the lower right quadrant of
Figure 2A, evaluating the 2 month culture in East African trials,
two were of comparisons where rifampicin was added beyond two
months. Trials have shown that rifampicin throughout treatment is
critical[39], but this cannot be reflected in the two month culture
result. Restricting the analysis to all comparisons of two regimens
that contained rifampicin throughout treatment gives considerably
better results for the two month culture ($R^2_{\text{trial}} = 0.67$ compared to
$R^2_{\text{trial}} = 0.36$), but not as good as in the sub-group of Hong Kong
trials only ($R^2_{\text{trial}} = 0.86$).

Fundamentally, a marker that is measured before the end of
treatment cannot capture the full effect of the treatment regimen
and can never therefore be a perfect surrogate. This can be
illustrated in a trial comparing a six month regimen with rifampicin
throughout with an eight month regimen with rifampicin for only the first two months[39]. The two-month intensive phase was unchanged and therefore the proportion of
culture positive patients at two months was similar (17% and 14%
respectively) but the proportion with unfavourable outcomes at the
end of follow-up was significantly different (5% and 10%
respectively, $p<0.01$). The optimum time for measuring a marker
that could be a surrogate at the end of treatment. Unlike the
situation in HIV where CD4 count or viral load can always be
measured, TB patients are almost without exception negative on
culture at the end of treatment (unless they have failed treatment
or have an uninformative isolated positive) and therefore it is likely
that more sensitive methodologies that can detect the presence of
$M. \text{tuberculosis}$ later in treatment will be needed. These could
include a molecular viable count assay[40], resuscitation-promot-
ing factors[41] or cycle threshold of the Xpert MTB/RIF
assay[42,43].

Conclusions

Without a better understanding of the main cause of the
heterogeneity of results, neither the two month nor the three
month culture on solid media can be recommended for use as the
primary endpoint in a phase III clinical trial - the ultimate
objective for a putative surrogate.

The results are encouraging, however, in that culture results on
solid media during treatment capture a moderate proportion of
the treatment effect on long-term outcome and are appropriate as
endpoints for phase II trials to identify promising regimens to take
forward to phase III for more rigorous evaluation. An endpoint
that is a summary measure of the longitudinal profile of culture
results over time or that is able to detect the presence of $M. \text{tuberculosis}$
later in treatment is more likely to be a better endpoint
for a phase II trial than a culture result at a single time point and
may prove to be an acceptable surrogate. More data are needed
before any endpoint can be used as a surrogate in a confirmatory
phase III trial.

Author Contributions

Conceived and designed the experiments: PPJP KF AJN. Performed the
experiments: PPJP. Analyzed the data: PPJP. Contributed reagents/
materials/analysis tools: PPJP. Wrote the paper: PPJP KF AJN.
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PLOS ONE | www.plosone.org 10 May 2013 | Volume 8 | Issue 5 | e63840